

## On the Robustness of Anchoring Effects in WTP and WTA Experiments<sup>†</sup>

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*We reexamine the effects of the anchoring manipulation of Ariely, Loewenstein, and Prelec (2003) on the evaluation of common market goods and find very weak anchoring effects. We perform the same manipulation on the evaluation of binary lotteries, and find no anchoring effects at all. This suggests limits on the robustness of anchoring effects. (JEL C91, D12, D44)*

Decades of evidence from individual-choice experiments have documented that subjects do not always satisfy the assumptions of consumer theory and of expected utility theory. Some departures from the standard model can be accommodated while maintaining the assumption that people do have stable preferences. Experiments on anchoring effects are more troubling as they suggest that many people do not have fixed or slowly changing preferences. A defender of the standard models might argue that this is not relevant to most everyday consumer decisions, and that an average person with enough experience will have relatively fixed preferences for common market goods. This suggests that experience with the goods and decisions in question may be important in the stability of preferences. Since people are relatively familiar with risky monetary choices, this in turn suggests that decisions about uncertain prospects (lotteries) might be relatively stable.

If the monetary valuation of goods and gambles is based on constant underlying preferences, choices should be independent of irrelevant prior “anchoring” questions. In the case of lotteries, Johnson and Schkade (1989) showed that asking subjects whether they would prefer a fixed amount (the anchor) to a given gamble affected subsequently stated certainty equivalents. Chapman and Johnson (1999) had subjects generate a random anchor from their Social Security Number (SSN),

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and found very strong anchoring effects on the stated willingness to accept (WTA) for selling a purely hypothetical lottery.

In the case of goods, Ariely, Loewenstein, and Prelec (2003, henceforth ALP) used a simple anchoring manipulation to provide evidence that people do not have constant preferences. They first generated a personal random number between 0 and 99 for each subject (the anchor), asking them to turn the last two digits of their SSN into a dollar price. They then asked subjects whether they would buy particular consumer goods for this price. Subsequently, they elicited subjects' willingness to pay (WTP) for the goods using the Becker, DeGroot, and Marschak (1964, henceforth BDM) mechanism,<sup>1</sup> and found that WTP was significantly correlated with the anchors.

These results seem troubling for economic theory. First, they are large. ALP write: "The effect is even more striking when examining the valuations by quintiles of the social security number distribution [...] the valuations of the top quintile subjects were typically greater by a factor of *three*. For example, subjects with SSN in the top quintile were willing to pay \$56 on average for a cordless computer keyboard, compared with \$16 on average for subjects with bottom quintile numbers. Evidently, these subjects did not have, or were unable to retrieve personal values for ordinary products." Similarly, the average stated WTP for a bottle of average wine, for subjects whose SSN-based anchor price was in the highest quintile, was \$27.90, more than three times as high as the analogous average WTP for subjects, whose SSN-based anchor price was in the lowest quintile. Moreover, ALP present evidence that their results are robust across many dimensions: repetition, different levels of simplicity of the "hedonic experience" (consuming a good is one such experience), different subject pools, different levels of monetary incentives, and markets versus individual choice. We believe that the ALP experiments are important and influential enough that they deserve to be replicated and tested for robustness.

We conducted four sets of experiments in order to examine whether preferences for common market goods and for simple lotteries are "arbitrary" in the sense of ALP. We examined choices for both lotteries and goods. In the case of goods we examined WTP and WTA and whether or not explaining the incentive properties of the BDM procedure matters. For lotteries we focused on eliciting WTA with the BDM incentives explained. These experiments enable us to examine new questions as well as examining robustness: Are preferences for lotteries or common goods more arbitrary? Are anchoring effects more pronounced for WTA or WTP?

Our first finding is that we are unable to replicate the results of ALP: we find very weak anchoring effects both with WTP and with WTA. The Pearson correlation coefficients between the anchor and stated valuation are generally much lower than in ALP, and the magnitudes of the anchoring effects (as measured by the ratio of top to bottom quintile) are smaller. Repeating the ALP procedure for lotteries we do not find any anchoring effects at all.

Unlike ALP, we carried out laboratory rather than classroom experiments. This necessitated some minor changes—discussed below—from ALP's procedures. It is

<sup>1</sup>This mechanism is incentive compatible under risk neutrality, but not in general. However, there is strong evidence that experimental subjects are not risk neutral even for small gambles.

conceivable that these changes are responsible for the differences in our findings; if so the robustness of their results is limited.

### I. Experiment 1: Eliciting WTA for Common Market Goods

All of our experiments were paper-and-pencil as was the case in ALP. Our first experiment reexamined the effects of random anchors on subjects' valuation for common market goods. The experiment took place in the California Social Science Laboratory (CASSEL) at UCLA, in August of 2009, and subjects were UCLA students. Subjects were recruited using the standard procedure of the lab, which is sending an announcement to a large e-mail list. A total of 84 subjects participated,<sup>2</sup> and each of the three sessions had 26–28 subjects.

Our subject pool consisted almost exclusively of undergraduate students, and the goods were chosen to be of interest for them: an academic planner, a cordless keyboard, a financial calculator, a designer book, a pack of quality chocolates, and a cordless mouse.<sup>3</sup> The average retail price of the goods was \$51.70. When the experiment started, subjects were asked to press the "Enter" key and generate a random number between 0 and 99 (we used the computer program Excel). Subjects were then asked to copy this number six times in their answer sheet. After subjects had done so, we started reading the instructions loudly, so that everyone could hear.

For each good, subjects had to answer two questions. The first (anchor) question asked them whether they would be willing to sell the good for an amount of money equal to the number that they had written down. In the second question, subjects were asked to state the minimum price for which they would be willing to sell the good. They were instructed to act as if the good was theirs to keep since there was a positive probability that they would truly own it. We explained that three subjects would be randomly drawn. Each of the three chosen persons would be assigned one of the goods (a different good for each of the three subjects drawn). For each of them, one of her answers (corresponding to her assigned good) would be carried out for real.<sup>4</sup>

If a subject was the first person randomly drawn, her answer to the first question would be consequential. So if she stated that she wanted to sell the good for an amount equal to her random anchor she did so. Otherwise she kept the good and received it at the end of the experiment. If a subject was the person randomly drawn

<sup>2</sup>Although 84 subjects participated, five subjects were excluded from the data analysis. One subject failed to write a WTA for any of the goods, and another subject failed to write a WTA for one of the six goods. One subject failed to write the same anchor number for the six goods. One subject wrote a random number equal to \$100, which was not in the range of possible random numbers (0–99). Finally, one subject failed to answer any of the anchoring questions (the first question for each of the goods), leaving the space blank. It is worth emphasizing that we have not excluded a subject whose WTA for the chocolates and the cordless mouse was \$100 and \$300, respectively, because it is not a priori impossible for someone to have such WTA for these common goods. So our actual sample size consists of 79 subjects.

<sup>3</sup>Our principal objective was to replicate the ALP experiments in a laboratory. Economic laboratories, which work with large sample pools, typically use undergraduate students as subjects. This constrained the range of goods that we could use (for example, in the USA, wine cannot be procured to people below 21 years of age). In addition, changes had to be made due to the elapsed time from the ALP experiment (cordless trackballs were largely unknown to undergraduate students at the time of our experiment). We recognize that using different goods adds an additional confounding factor, but note that some of our goods were very similar to ALP.

<sup>4</sup>This process emulates the procedures of ALP.

second or third, her answer to the second question (for the assigned good) would matter. We explained that in this case, a price would be randomly drawn from a matrix containing 100 values. The values of the matrix were drawn before the experiment from a triangular distribution in  $[0, 100]$  with a mode equal to zero. Subjects were shown the matrix on the screen. If the stated WTA was higher than the random price, the subject kept the good, and if not, she received a monetary amount equal to the random price, and sold the good. We also explained why the BDM mechanism gives an incentive to truthfully reveal preferences.<sup>5</sup>

We explained that all uncertainty would be resolved using physical devices. The random price would be chosen from the 100 possible prices by throwing a ten-sided dice twice. The three consequential goods would be determined by throwing a six-sided dice three times. The three chosen subjects would be chosen using a bingo cage with balls numbered according to numbers pre-assigned to the subjects.

After we had finished explaining the instructions, we showed the six goods to subjects. We displayed each good in clear view, and a photograph of each good was also projected on a screen. After showing each good we gave subjects time to complete the answers to both questions. After they finished, they were asked to copy their responses to a second answer sheet. After we picked up the second answer sheet, we resolved all uncertainty as explained, and concluded the experiment.

Given the results of ALP, our prior expectation was that we would find a positive relationship between the personal random number and the stated valuation. Thus in all our experiments, we use one-tailed statistical tests. For example, we reject the null hypothesis of a zero Pearson correlation between the personal random number and the WTA at a certain significance level if we get a high enough positive value of the sample correlation.

The average cash payment per subject was \$13.80. Moreover, subjects received goods of total retail value equal to \$298.00, which implies that the average value of earned goods per subject was about \$3.50.<sup>6</sup>

Table 1 presents the main results. For each of the five quintiles of the distribution of the personal random number, we report the mean and the median WTA (the medians are in the parentheses). We also report the Pearson correlation coefficients, both to understand whether there is a linear relationship between the WTA and the anchor and in order to allow comparisons with ALP. As seen in the table, the Pearson correlation coefficients between the random anchor and stated WTA were relatively

<sup>5</sup>Prelec explained the BDM procedure as part of his class prior to conducting the experiment. As we used paid participants from a laboratory pool, we instead explained the procedure as part of the experiment. We based our explanation on the instructions of Plott and Zeiler (2005). In particular, we said: "Please, leave your instructions aside for a while. With the second question, we want to know how much the opportunity to have each item is worth to you. This method simply gives you an incentive to state the true minimum you would be willing to accept for selling each item. Why is your best strategy to write the minimum you would be willing to accept? Because you will not receive the amount you ask for. Instead, if you receive anything, you will receive the random price. Let's try to understand this with an example. Imagine that I own the item. Say I believe that this item is worth \$500 for me. What happens if I falsely say that the minimum I would be willing to accept is more than 500? Assume I say \$600. If the random price is, say, 530, I do not sell the item. But, had I said 500, I could receive 530 for an item that I think is only worth 500. I lose out. What happens if I say less than 500? For example, I say only 400. If the random price is, say, 420, then I have to sell my item. I lose out, because I have to sell for 420 an item that I think is worth 500." Note that these instructions implicitly tell subjects to be risk neutral.

<sup>6</sup>Each of the participants who were not chosen received a participation fee of \$13. Chosen persons also received this fee, in addition to being assigned one good.

TABLE 1—EXPERIMENT 1, AVERAGE AND MEDIAN WTA SORTED BY QUINTILE OF THE DISTRIBUTION OF THE ANCHOR

Quintile (# obs.)	Academic planner	Cordless keyboard	Financial calculator	Designer book	Milk chocolates	Cordless mouse
0–19 (13)	7.00 (5)	38.85 (35)	10.23 (10)	11.30 (10)	5.62 (4)	22.07 (25)
20–39 (16)	12.44 (10)	50.19 (47.5)	20.37 (20)	14.75 (15)	17.88 (7)	42.88 (23)
40–59 (16)	11.50 (10)	51.94 (50)	20.94 (15.5)	18.12 (17.5)	8.47 (5.25)	24.56 (20)
60–79 (20)	10.05 (8)	38.00 (30)	18.78 (11)	16.50 (13.5)	7.47 (5)	19.95 (15)
80–99 (14)	7.64 (7.5)	47.28 (30)	31.92 (23.5)	15.57 (15)	7.36 (6.5)	26.85 (20)
All data	9.90 (9)	45.07 (40)	20.46 (15)	15.46 (15)	9.46 (6)	27.10 (20)
Pearson	–0.047	–0.022	0.207	0.134	–0.096	–0.108
<i>p</i> -value	0.66	0.58	0.03	0.12	0.80	0.83
Spearman	–0.046	–0.059	0.208	0.126	0.070	–0.133
<i>p</i> -value	0.66	0.70	0.03	0.13	0.27	0.73

small: they range from  $-0.108$  to  $0.207$ , and only one of them is statistically significant. We also report the (non-parametric) Spearman correlation coefficient and find similar values. In contrast, the Pearson correlation coefficients in ALP's study range from  $0.319$  to  $0.516$  and are all highly statistically significant, despite their smaller sample size.<sup>7</sup>

To get further evidence for the importance of anchoring, we examine the ratio of stated WTA for subjects in the highest quintile of random numbers to those in the lowest quintile. Recall that ALP report ratios of about  $2.3$  and  $2.2$  for two of the goods, and between  $3$  and  $3.45$  for the remaining four goods. Table 2 reports our results. For one of our goods the ratio was  $2.9$ , for the other five the ratio ranged from  $1.1$  to  $1.4$ . Using the (nonparametric) Wilcoxon rank-sum test we are unable to reject the null hypothesis of homogeneity at the 5 percent significance level for the five goods with ratios between  $1.1$  and  $1.4$ . For the remaining good (the financial calculator) the hypothesis is rejected at the 5 percent significance level. Notice that if we test six different true hypotheses with different data it is not unlikely that one of them will be rejected at the 5 percent level. On the other hand all the ratios are greater than  $1$ , which is not likely if there is no anchoring at all.<sup>8</sup>

## II. Experiment 2: Willingness to Pay

Our second experiment examines whether the differences of our results in Experiment 1 from ALP's were driven by the fact that we elicited WTA and ALP elicited WTP. This treatment also took place at the CASSEL lab at UCLA, in April of 2010, and the participants were UCLA students. A total of 79 subjects participated, in three sessions of 26 or 27 students.<sup>9</sup> The average payment per subject was equal to  $\$21.8$ .<sup>10</sup>

<sup>7</sup>The sample of ALP consisted of 55 subjects, whereas we had 79 subjects.

<sup>8</sup>The statistical tests reported in the main text included all of our data; in Appendix 1 we verify that our conclusions about which effects are significant do not change if we exclude outliers and subjects with inconsistent responses, as has been done in some past work on this topic.

<sup>9</sup>One subject failed to answer any of the six anchoring questions and was excluded for the analysis. Hence, our final dataset consists of 78 subjects.

<sup>10</sup>The subjects who were not chosen received a participation fee of  $\$13$ .

TABLE 2—EXPERIMENT 1, WILCOXON TEST OF HOMOGENEITY OF THE DISTRIBUTIONS OF WTA FOR SUBJECTS IN THE FIRST AND FIFTH QUINTILE OF THE DISTRIBUTION OF THE ANCHOR

Item	Academic planner	Cordless keyboard	Financial calculator	Designer book	Milk chocolates	Cordless mouse
$z$	-0.587	-0.685	-3.008	-1.424	-1.342	-0.171
$p$ -value	0.279	0.247	0.001	0.077	0.090	0.423

The sessions were nearly identical to those of Experiment 1. The main difference was that subjects were asked to state the most they would be willing to pay for the goods.<sup>11</sup> Hence, the subjects that would be chosen at random would not own the good, but they would have the chance to buy it. To conform to the standards of the CASSEL lab, a large participation fee (\$93) was offered to the three chosen subjects to make sure they could buy each good, and in order to prevent subjects from losing money in the experiment.

The results are shown in Table 3. We again examine the ratio of stated WTP for subjects in the highest quintile of random numbers to that of subjects in the lowest quintile. One ratio is quite large, 1.4. Four are less than one and lie between 0.9 and 1.0, and the sixth, the financial calculator, has a very low ratio of 0.7, while in Experiment 1 it had a high ratio. However none of the ratios are statistically significant at the 5 percent level according to the Wilcoxon rank-sum test. Moreover the Pearson and the Spearman correlations are all very close to zero.<sup>12</sup>

From this data, we can conclude that anchoring effects are not stronger for WTP than WTA. Our data is consistent with the possibilities that anchoring effects are weaker for WTP than WTA and that there are negative anchoring effects, but we cannot make definitive statements about these issues.

### III. Experiment 3: Is it the BDM Explanations?

In Experiments 1 and 2 (as well as in experiment 4) we included part of the explanations of the BDM mechanism that was used by Plott and Zeiler (2005); our objective was to try to account for the explanations that were provided as a part of a classroom lecture before the experiments of ALP. As we could not replicate their classroom training it is possible that this is why our results differ from theirs. To examine the importance of the explanation of the BDM elicitation procedure our third experiment provides no explanations of the concept of WTP/WTA or of the BDM mechanism.

The experiment was also conducted at the CASSEL laboratory at UCLA, in February of 2011, and all subjects were UCLA students. A total of 81 subjects

<sup>11</sup> Four of the six goods were the same as in Experiment 1, while two goods were not the same, but very similar. A class organizer was used instead of an academic planner, and a very similar cordless mouse was used instead of the one used in Experiment 1.

<sup>12</sup> As have earlier researchers, we find a possible WTP-WTA gap, however it is a relatively small one, in the order of 20–30 percent.

TABLE 3—EXPERIMENT 2, AVERAGE AND MEDIAN WTA SORTED BY QUINTILE OF THE DISTRIBUTION OF THE ANCHOR

Quintile (# obs.)	Class organizer		Cordless keyboard		Milk chocolates		Cordless mouse		Designer book		Financial calculator	
0–19 (20)	3.25	(1)	30.70	(21)	4.82	(4)	16.75	(14.5)	7.70	(4.5)	15.05	(11.5)
20–39 (11)	2.95	(2)	33.45	(30)	4.77	(4)	15.63	(10)	5.55	(4)	12.09	(10)
40–59 (12)	4.00	(2)	32.33	(35)	4.83	(5)	17.58	(15)	8.91	(5)	7.66	(6)
60–79 (19)	4.37	(3)	33.37	(25)	5.00	(5)	16.95	(20)	8.32	(4)	13.74	(10)
80–99 (16)	4.81	(1)	28.31	(25.5)	4.19	(3)	12.87	(10)	6.25	(3)	9.81	(6)
All data	3.92	(1.5)	31.50	(25)	4.73	(4)	15.97	(15)	7.43	(4.5)	12.10	(8.5)
Pearson	–0.047		–0.022		–0.043		–0.039		–0.013		–0.080	
<i>p</i> -value	0.66		0.58		0.65		0.63		0.54		0.76	
Spearman	–0.046		–0.059		–0.007		–0.029		–0.013		–0.074	
<i>p</i> -value	0.66		0.70		0.52		0.60		0.55		0.74	

TABLE 4—EXPERIMENT 2, WILCOXON TEST OF HOMOGENEITY OF THE DISTRIBUTIONS OF WTA FOR SUBJECTS IN THE FIRST AND FIFTH QUINTILE OF THE DISTRIBUTION OF THE ANCHOR

Item	Class organizer	Cordless keyboard	Milk chocolates	Cordless mouse	Designer book	Financial calculator
<i>z</i>	0.148	–0.191	0.562	0.959	0.500	1.024
<i>p</i> -value	0.558	0.424	0.713	0.831	0.691	0.847

participated in three sessions of 27 subjects each.<sup>13</sup> Subjects' average cash payment was equal to \$15.64, and in addition subjects received goods of total value equal to \$153, for an average value of earned goods per subject of about \$1.88.<sup>14</sup>

The experiment was conducted in exactly the same way as Experiment 1, except without the additional explanations of the BDM mechanism.<sup>15</sup> The type and quality of each good was the same as in Experiment 1, except for using a telephone address book instead of an academic planner.

The results, shown in Table 5, are similar to our other two experiments. We again examine the ratio of stated WTA for subjects in the highest quintile of random numbers to the stated WTA of those in the lowest quintile. One of these is 1.5 (this time the cordless mouse rather than the financial calculator). Four lie between 1.0 and 1.1, and the sixth is 1.2. Needless to say the Wilcoxon rank-sum tests, shown in Table 6, are not significant. The Pearson correlation coefficients are very low although for one good the null hypothesis of a zero correlation is rejected at the 5 percent significance level.

<sup>13</sup>One subject failed to state a WTA for one of the goods, and another failed to copy the correct random number (wrote different numbers rather than a single one). These subjects were excluded from the observations, leaving a total of 79 subjects.

<sup>14</sup>Each of the participants who were not chosen received a participation fee of \$13. Chosen persons also received this fee, in addition to being assigned one good.

<sup>15</sup>The written instructions were short. Appendix 2 contains the instructions for the lotteries experiment, which are similar to the written instructions used in the other three experiments.

TABLE 5—EXPERIMENT 3, AVERAGE AND MEDIAN WTA SORTED BY QUINTILE OF THE DISTRIBUTION OF THE ANCHOR

Quintile (# obs.)	Address book		Cordless keyboard		Financial calculator		Designer book		Milk chocolates		Cordless mouse	
0–19 (16)	8.31	(7.5)	39.69	(35)	20.81	(18.5)	14.31	(15)	7.28	(6.5)	19.93	(20)
20–39 (15)	10.06	(8.95)	31.60	(26)	21.40	(20)	17.13	(10)	5.93	(5)	17.26	(17)
40–59 (15)	12.60	(10)	42.53	(40)	17.80	(18)	20.20	(20)	11.13	(6)	25.46	(20)
60–79 (20)	6.55	(5)	41.05	(40)	21.80	(15.5)	17.35	(15)	5.10	(5)	20.85	(19)
80–99 (13)	10.38	(10)	41.62	(40)	21.69	(20)	16.30	(15)	7.69	(6)	29.30	(25)
All data	9.35	(8)	39.35	(35)	20.74	(18)	17.06	(15)	7.27	(6)	22.25	(20)
Pearson	0.001		0.097		0.022		0.041		0.003		0.202	
<i>p</i> -value	0.50		0.20		0.42		0.36		0.49		0.04	
Spearman	−0.013		0.123		0.017		0.012		−0.042		0.152	
<i>p</i> -value	0.54		0.14		0.44		0.46		0.64		0.09	

TABLE 6—EXPERIMENT 3, WILCOXON TEST OF HOMOGENEITY OF THE DISTRIBUTIONS OF WTA FOR SUBJECTS IN THE FIRST AND FIFTH QUINTILE OF THE DISTRIBUTION OF THE ANCHOR

Item	Address book	Cordless keyboard	Financial calculator	Designer book	Milk chocolates	Cordless mouse
<i>z</i>	−0.904	−0.529	−0.507	−0.199	−0.309	−1.323
<i>p</i> -value	0.183	0.298	0.306	0.421	0.379	0.092

#### IV. Experiment 4: Eliciting WTA for Lotteries

In our fourth experiment, we elicited WTA for six money lotteries with a range of expected values similar (but not identical) to the range of the market prices of the six goods. The experiment also took place in the CASSEL lab at UCLA during November of 2008 and August of 2009. A total of 110 subjects participated,<sup>16</sup> in four sessions of 26–28 subjects each. Subjects were UCLA students, and the great majority was undergraduate students. The average total earnings per subject were equal to \$21.7.<sup>17</sup>

The six gambles are shown in Table 7. The procedures and the instructions were similar to Experiment 1, with lotteries instead of goods.<sup>18</sup> In each experimental session,

<sup>16</sup>We exclude from the data a person who did not state a WTA for one of the lotteries. We also exclude a participant who stated a WTA equal to \$300 for the first lottery, and a WTA \$400 for the sixth lottery. Since the maximum prize of any lottery is \$100, we felt that it is impossible that someone could actually have such a WTA for the lotteries. Therefore, our final sample size is 108.

<sup>17</sup>All subjects received a participation fee of \$16. The three subjects who were randomly chosen to receive a lottery could earn money on top of that. The average market value of our experimental goods in Experiments 1 and 3 was about \$52, and the large “extra fee” of the chosen subjects in Experiment 2 was equal to \$80. The average expected value for the lotteries was about \$38. We used a somewhat higher participation fee in the lottery sessions in order to roughly equalize the ex-ante average expected payoffs for the lotteries and common goods treatments.

<sup>18</sup>The only difference was that in this experiment we further explained to subjects how they should answer the questions. In particular, after presenting the two questions, we drew attention away from the written instructions, and told subjects: “Please go back to the last two pages of your instructions but please do not write anything yet. Please have a look at the lotteries and the questions you are asked to answer. The first question for each lottery is easy: do you prefer to have the opportunity to play this lottery or the specified amount of money? An example will illustrate what you should answer to the second question. Imagine that I own the right to play a particular lottery. Say this sample lottery ticket. How do I know what amount is the minimum I’d be willing to accept for the lottery? First of all, notice that, if I keep it, this lottery ticket gives me the opportunity to earn some money with a certain probability. I am not willing to give it up without getting some money. The question is, what is the minimum I would accept? Start with \$100. Would I be willing to give up my lottery in exchange for 100? If so, decrease the amount

TABLE 7—EXPERIMENT 4, AVERAGE AND MEDIAN WTA SORTED BY QUINTILE OF THE DISTRIBUTION OF THE ANCHOR

Quintile (# obs.)	0.5, \$100; 0.5, \$0	0.25, \$100; 0.75, \$0	0.1, \$100; 0.9, \$0	0.75, \$60; 0.25, \$20	0.5, \$60; 0.5, \$20	0.9, \$60; 0.1, \$20
0–19 (18)	42.11 (40)	24.67 (20)	14.67 (10)	41.11 (40)	34.39 (35)	49.39 (52)
20–39 (13)	49.69 (50)	26.62 (25)	15.77 (11)	44.69 (45)	37.00 (40)	48.69 (50)
40–59 (22)	51.54 (50)	31.82 (32.5)	23.55 (20)	46.05 (45)	37.86 (40)	55.32 (55)
60–79 (21)	41.67 (50)	22.19 (25)	10.23 (10)	39.67 (45)	34.67 (40)	49.52 (55)
80–99 (34)	50.02 (50)	31.45 (27)	21.15 (13.5)	43.15 (42.5)	34.18 (30)	51.38 (52.5)
All data	47.35 (50)	28.01 (25)	17.78 (10)	42.90 (45)	35.40 (39.5)	51.17 (55)
Pearson	0.057	0.082	0.071	−0.010	−0.076	0.032
<i>p</i> -value	0.28	0.20	0.23	0.54	0.78	0.37
Spearman	0.119	0.090	0.034	−0.026	−0.104	0.025
<i>p</i> -value	0.11	0.17	0.36	0.60	0.86	0.40

three subjects would be drawn at random; each would get to own one lottery for real, and would either keep it or sell it depending on the answers given to the experimental questions. The lottery was carried out for real, using a dice. A subject who chose to keep the lottery, by indicating a WTA higher than the random price, would receive the prize drawn from the lottery. Otherwise, he or she simply received the random price from selling the lottery. All of this was made common knowledge to subjects.

We again examine the ratio of stated WTA for subjects in the highest quintile of random numbers to those in the lowest quintile. Three of the ratios lie between 0.99 and 1.05, and the remaining three lie between 1.2. and 1.4. This is smaller even than the modest effect we found for goods. Despite the larger sample size none of the Wilcoxon rank-sum tests reported in Table 8 are significant at the 5 percent level. Nor are any of the Pearson and Spearman correlation coefficients in Table 7.

The lack of anchoring in our experiment might be due to the simplicity of the lotteries we considered. Chapman and Johnson (1999) used an SSN-based anchor, but for the complicated binary lottery (\$287, 17 percent; \$18.5, 83 percent), and found a Pearson correlation coefficient equal to 0.45, whereas the largest coefficient that we find for any lottery is equal to 0.082.

## V. Conclusion

There were several differences between our experimental design and ALP. They elicited subjects' WTP for six common goods: Belgian chocolates, a wireless keyboard, a designer book, a bottle of rare wine, a bottle of cheap wine, and a cordless

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to 95. If I am willing to accept 95 in exchange for the lottery, then decrease further. I keep decreasing until I come to the amount that makes me indifferent between keeping the lottery and getting the money. Your objective is to think like this for all six lotteries. You should think carefully before answering these questions because these questions are not purely hypothetical. There is a chance you will actually own each lottery, and each of your decisions may be carried out for real money." This part was also based on the instructions by Plott and Zeiler (2005).

We attempted to use the minimal set of instructions that made clear what the rules of the procedure were. Some elements of this are arbitrary: for example, we told subjects that they might start at 100 and consider lower amounts until they were indifferent. It is possible that different results might have obtained had they been told to start at zero and consider higher amounts. However, it is hard to see how this might impact on the framing effect of being told to record particular numbers in advance.

TABLE 8—EXPERIMENT 4, WILCOXON TEST OF HOMOGENEITY OF THE DISTRIBUTIONS OF WTA FOR SUBJECTS IN THE FIRST AND FIFTH QUINTILE OF THE DISTRIBUTION OF THE ANCHOR

	0.5, \$100; 0.5, \$0	0.25, \$100; 0.75, \$0	0.1, \$100; 0.9, \$0	0.75, \$60; 0.25, \$20	0.5, \$60; 0.5, \$20	0.9, \$60; 0.1, \$20
<i>z</i>	−1.577	−1.312	−1.125	−0.371	0.304	−0.485
<i>p</i> -value	0.058	0.095	0.130	0.355	0.619	0.313

trackball. They performed the experiment in a class with 55 MBA students, and decisions were consequential for only six subjects, drawn at random, one for each good. We conducted the experiments in a lab environment, and we changed some of the goods, because of technology changes and age restrictions. In three treatments we used part of the Plott and Zeiler (2005) explanation of the BDM mechanism, but a fourth treatment suggests that this deviation from ALP was not a key factor in the earlier results.

We also preferred not to use the SSN-based anchor, since during a pilot some subjects showed privacy concerns. This might be another possible reason for the difference in our results. There is evidence that subjects might not view such numbers as truly arbitrary. For example, one-third of participants in one of the experiments of Chapman and Johnson (1999) stated that they thought the SSN-based anchor number was informative. In another experiment, 57 percent said that they thought that the experimenter wanted the number to influence their judgment. Perhaps an SSN-based anchor is not seen as random, in the sense that the value of the SSN has been determined long ago.<sup>19</sup>

Our results do not confirm the very strong anchoring effects found in ALP. They are more in agreement with the results of Simonson and Drolet (2004) and Alevy, Landry, and List (2011). Simonson and Drolet (2004) used the same SSN-based anchor as ALP, and found no anchoring effects on WTA, and moderate anchoring effects on WTP for four common consumer goods. Alevy, Landry, and List (2011) performed a field experiment, eliciting the WTP for peanuts and collectible sports cards, and they found no anchoring effects. Bergman et al. (2010) also used the design of ALP for six common goods, and found anchoring effects, but of smaller magnitude than in ALP.

Tufano (2010) and Maniadis, Tufano, and List (2011) also failed to confirm the robustness of the magnitude of the anchoring effects of ALP, using hedonic experiences, rather than common goods. Tufano (2010) used the anchoring manipulation to increase the variance in subjects' WTA for a bad-tasting liquid, but the manipulation had no effect. Notice that this liquid offers a simple (negative) hedonic experience, like the "annoying sounds" used in Experiment 2 of ALP. Maniadis, Tufano, and List (2011) replicated Experiment 2 of ALP and found weaker (and nonsignificant) anchoring effects. Overall our results suggest that anchoring is real—it is

<sup>19</sup>The final digits of the social security number are not random and are in fact correlated with personal characteristics such as the region and time of year of birth (see, e.g., Acquisti and Gross 2009). It seems unlikely that this correlation leads to a correlation between final digits of social security numbers and bidding behavior, but at this point we cannot rule that out.

hard to reconcile otherwise the fact that in the WTA treatment with goods the ratios between highest and lowest quintile is always bigger than one—but that quantitatively the effect is small. Additionally our data supports the idea that anchoring goes away when bidding on objects with greater familiarity, such as lotteries. Our results do not suggest that anchoring is stronger with WTP than WTA—if anything the reverse is true.

#### APPENDIX 1: ROBUSTNESS CHECK OF OUR RESULTS

We tried to follow the ALP experiment as much as possible, so we did not include a long practice process, or very detailed instructions. Nevertheless, as we explained above, there is reason to believe that the notions of WTA and WTP, as well as the BDM mechanism, are not very easy to understand. Accordingly, some participants in our experiment showed signs of bidding “strategically,” or otherwise showed signs that they have not fully understood the valuation task. Moreover, several subjects had their second answers inconsistent with their response in the first question, for the same good or lottery.

In Experiment 1, one subject assigned the same WTA (\$20) to all goods, and another stated implausibly similar WTA for each of the six goods (in a range from \$22 to \$30). We believe that this behavior indicates insufficient understanding, because it is highly implausible that the actual WTA for so different goods lies in such a narrow range. Another subject stated a WTA equal to \$5 for goods 1–4 and equal to \$100 and \$300 for goods 5 and 6, respectively, which highly unlikely to express true valuation. Moreover, several subjects had “inconsistent” responses, in the sense explained above. Accordingly, we considered a total of 15 subjects to have “questionable responses.”

In Experiment 2, some subjects stated improbably low willingness to pay for all the goods (less than \$10 for any of the goods). The tendency to behave strategically in a bargaining situation is a plausible reason. Moreover, one subject stated a WTP of \$50 for the good with the lowest average WTP (the first good, the class organizer), and a WTP less than \$10 for all the other goods. This might indicate confusion. There were also a few inconsistent subjects, and, in total, 11 subjects in Experiment 3 had “questionable responses.”

In Experiment 3, two subjects stated the same WTA for all goods, one subject only used two different WTAs for the six different goods (each of this subject’s WTA were either \$20 or \$5) and four subjects’ stated WTAs had an excessively narrow range of \$10 or less. Several other subjects stated inconsistent responses, and a total of 18 observations were characterized as “questionable.”

In Experiment 4, one subject assigned the same WTA (\$5) to all lotteries. Four subjects stated only two different WTAs for the six different lotteries (for example, either \$97 or \$30, either \$10 or \$56). Three subjects stated an implausibly similar WTA for the different lotteries (for one subject, the WTA of all lotteries ranged from \$15 to \$18, for another from \$48 to \$60 and for the third from \$46 to \$60). This behavior might indicate improper understanding of the task, for the reasons explained above. Moreover, six other subjects stated implausibly large WTA values (of at least \$50 for each lottery) or wrote a WTA higher than the highest possible

TABLE 9—CORRELATIONS FOR THE SAMPLES THAT EXCLUDE OBJECTIONABLE AND SUSPICIOUS OBSERVATIONS

Experiment 1 <i>N</i> = 64	Academic planner	Cordless keyboard	Financial calculator	Designer book	Milk chocolates	Cordless mouse
Pearson	0.033	-0.063	0.264	0.120	0.054	-0.137
<i>p</i> -value	0.40	>0.5	0.02	0.17	0.34	>0.5
Experiment 2 <i>N</i> = 67	Class organizer	Cordless keyboard	Milk chocolates	Cordless mouse	Designer book	Financial calculator
Pearson	-0.026	-0.030	-0.067	-0.055	-0.020	-0.018
<i>p</i> -value	>0.5	>0.5	>0.5	>0.5	>0.5	>0.5
Experiment 3 <i>N</i> = 61	Address book	Cordless keyboard	Milk chocolates	Cordless mouse	Designer book	Financial calculator
Pearson	-0.001	-0.005	-0.022	0.164	-0.017	-0.090
<i>p</i> -value	>0.5	>0.5	>0.5	>0.10	>0.5	>0.5
Experiment 4 <i>N</i> = 85	0.5, \$100; 0.5, \$0	0.25, \$100; 0.75, \$0	0.1, \$100; 0.9, \$0	0.75, \$60; 0.25, \$20	0.5, \$60; 0.5, \$20	0.9, \$60; 0.1, \$20
Pearson	0.041	0.112	0.052	-0.047	-0.083	-0.123
<i>p</i> -value	0.35	0.13	0.32	>0.5	>0.5	>0.5

prize. We believe that this might be caused by the tendency to behave “strategically,” asking for as much as possible in a bargaining situation. A few subjects also had inconsistent responses. We assigned a total of 16 subjects in Experiment 4 as having “questionable responses.”

We should note that in Session 4 of Experiment 4, one subject complained that some other subject/subjects had not written down the actual Excel random number, but a different one (by clicking “enter” multiple times). We should note that in this type of experiment, subjects might write a different anchor number than the one that they are instructed to write. If the anchor question is incentivized, then it is in the subjects’ best interest to write a number as large as possible in WTA experiments, and a number as low as possible in WTP experiments.<sup>20</sup> Excluding from the sample the eight subjects with anchor numbers higher than 80 in Session 4, we believe that we exclude the most suspicious observations. The possibility of cheating can never be ruled out, and excluding the suspicious observations leaves a very reasonable distribution of random numbers for all our experiments.

Characterizing observations as “questionable” or “suspicious” adds a necessarily subjective, and perhaps arbitrary, element. However, we feel that it is useful to examine whether our results are driven by confusion or cheating. For doing this, we can see whether excluding these subjects significantly changes the results. Table 9 describes the correlations between the anchor number and WTA/WTP for the samples where the questionable and suspicious observations are excluded: 15 observations from the data of Experiment 1, 11 observations from the data of Experiment 2, 18 observations of Experiment 3, and 23 observations from the data of Experiment 4.

<sup>20</sup>The mean of the anchor number that subjects wrote in Experiment 1 was 51.4, and the median 55. In Experiment 2, the mean was 49.7 and the median was equal to 53. In Experiment 3, the mean was 50.1 and the median 54. In Experiment 4, the mean was equal to 56.7 with a median equal to 60.

Table 9 reveals that the anchoring effects, measured as the Pearson correlations of the anchor with the stated valuation, are similar for the sample that excludes objectionable and suspicious observations. Note that very similar results are obtained, in Experiment 4, if we exclude only the objectionable observations, but not the suspicious ones, and vice versa. Therefore, the subjective element does not seem to affect the main results of our experiments.

#### APPENDIX 2: INSTRUCTIONS FOR THE LOTTERIES EXPERIMENT (EXPERIMENT 4)

Welcome to CASSEL. Please turn off pagers and cellular phones now. It is important that you do not talk, or with any other way try to communicate during the experiment. Thank you for participating in this pricing exercise. We are interested in how much you value various lotteries. We will show you six different lotteries, and then ask you a pair of price-selling intent questions about each lottery:

- First, we will ask whether you would like to sell the lottery at a particular price. Your computer generated that price randomly.
- Next, we will ask you to state the minimum that you would be willing to accept for selling the lottery.

Three persons will be chosen at random from the class and the decisions of each about one of the lotteries will be carried out for real money. One lottery will be sold on the basis of answers to the first question, and two lotteries on the basis of answers to the second question. Each person is on the running for only one lottery.

If you are chosen and the first question counts for you, then we will look to see whether you decided to sell, or not to sell, the lottery at your random computer-generated price. If you stated that you wanted to sell the lottery at that price, then we will buy it from you at that price. If you stated that you did not want to sell the lottery at that price, you will keep it.

#### If you are chosen and the second question counts:

- The price of the lottery will be determined by drawing a number at random from the matrix in the overhead slide.
- If the value is smaller than this random price, then the lottery will be sold *at that randomly selected price*.
- If the value stated is higher than the price, then the lottery will not be sold.

Because your answer does not affect the amount you get paid, only whether you sell, it's to your advantage to state the minimum that you would be willing to be paid for selling the lottery.

Each of you will receive a \$16 participation fee. In addition, if you are one of the three persons chosen, all results from lotteries you have chosen to keep will be drawn right away. Any amount that you will earn from the sale of a lottery, or from

the results of a lottery you keep, will be added to your participation fee. You will know your total earnings when you leave the experiment.

First, please copy your computer-generated random number in each of the 6 boxes

**LOTTERY TICKET 1**

50% chance of	50% chance of
<b>\$ 100</b>	<b>\$ 0</b>

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Would you sell this lottery ticket for \$  ? Circle YES or NO  
The least I would be willing to receive to sell this lottery is \$ \_\_\_\_\_

**LOTTERY TICKET 2**

25% chance of	75% chance of
<b>\$ 100</b>	<b>\$ 0</b>

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Would you sell this lottery ticket for \$  ? Circle YES or NO  
The least I would be willing to receive to sell this lottery is \$ \_\_\_\_\_

**LOTTERY TICKET 3**

10% chance of	90% chance of
<b>\$ 100</b>	<b>\$ 0</b>

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Would you sell this lottery ticket for \$  ? Circle YES or NO  
The least I would be willing to receive to sell this lottery is \$ \_\_\_\_\_

**LOTTERY TICKET 4**

75% chance of	25% chance of
<b>\$ 60</b>	<b>\$ 20</b>

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Would you sell this lottery ticket for \$  ? Circle YES or NO  
The least I would be willing to receive to sell this lottery is \$ \_\_\_\_\_

**LOTTERY TICKET 5**

50% chance of	50% chance of
<b>\$ 60</b>	<b>\$ 20</b>

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Would you sell this lottery ticket for \$ ? Circle YES or NO  
 The least I would be willing to receive to sell this lottery is \$ \_\_\_\_\_

**LOTTERY TICKET 6**

90% chance of	10% chance of
<b>\$ 60</b>	<b>\$ 20</b>

Would you sell this lottery ticket for \$ ? Circle YES or NO  
 The least I would be willing to receive to sell this lottery is \$ \_\_\_\_\_

REFERENCES

- Acquisti, Alessandro, and Ralph Gross.** 2009. "Predicting Social Security Numbers from Public Data." *Proceedings of the National Academy of Sciences* 106 (27): 10975–80.
- Alevy, Jonathan E., Craig E. Landry, and John A. List.** 2011. "Field Experiments on Anchoring of Economic Valuations." University of Alaska Anchorage Department of Economics Working Paper 2011–02.
- Ariely, Dan, George Loewenstein, and Drazen Prelec.** 2003. "'Coherent Arbitrariness': Stable Demand Curves without Stable Preferences." *Quarterly Journal of Economics* 118 (1): 73–106.
- Becker, Gordon M., Morris H. DeGroot, and Jacob Marschak.** 1964. "Measuring Utility by a Single-Response Sequential Method." *Behavioral Science* 9 (3): 226–32.
- Bergman, Oscar, Tore Ellingsen, Magnus Johannesson, and Cicek Svensson.** 2010. "Anchoring and Cognitive Ability." *Economics Letters* 107 (1): 66–68.
- Chapman, Gretchen B., and Eric J. Johnson.** 1999. "Anchoring, Activation, and the Construction of Values." *Organizational Behavior and Human Decision Processes* 79 (2): 115–53.
- Fudenberg, Drew, David K. Levine, and Zacharias Maniadis.** 2012. "On the Robustness of Anchoring Effects in WTP and WTA Experiments: Dataset." *American Economic Journal: Microeconomics*. <http://dx.doi.org/10.1257/mic.4.2.131>.
- Johnson, Eric J., and David A. Schkade.** 1989. "Bias in Utility Assessments: Further Evidence and Explanations." *Management Science* 35 (4): 406–24.
- Maniadis, Zacharias, Fabio Tufano, and John A. List.** 2011. "One Swallow Does Not Make a Summer: How Economists (Mis)Use Experimental Methods and Their Results." Unpublished. <http://didat-ica.unibocconi.it/mypage/index.php?IdUte=93696&idr=13905&lingua=ita>.
- Plott, Charles R., and Kathryn Zeiler.** 2005. "The Willingness to Pay–Willingness to Accept Gap, the 'Endowment Effect,' Subject Misconceptions, and Experimental Procedures for Eliciting Valuations." *American Economic Review* 95 (3): 530–45.
- Simonson, Itamar, and Aimee Drolet.** 2004. "Anchoring Effects on Consumers' Willingness-to-Pay and Willingness-to-Accept." *Journal of Consumer Research* 31 (3): 681–90.
- Tufano, Fabio.** 2010. "Are 'True' Preferences Revealed in Repeated Markets? An Experimental Demonstration of Context-Dependent Valuations." *Experimental Economics* 13 (1): 1–13.