Striatal hypofunction as a neural correlate of mood alterations in chronic pain patients


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ABSTRACT

Background: Chronic pain and mood disorders share common neuroanatomical substrates involving disruption of the reward system. Although increase in negative affect (NA) and decrease in positive affect (PA) are well-known factors complicating the clinical presentation of chronic pain patients, our understanding of the mechanisms underlying the interaction between pain and PA/NA remains limited. Here, we used a validated task probing behavioral and neural responses to monetary rewards and losses in conjunction with functional magnetic resonance imaging (fMRI) to test the hypothesis that dysfunction of the striatum, a key mesolimbic structure involved in the encoding of motivational salience, relates to mood alterations comorbid with chronic pain.

Methods: Twenty-eight chronic musculoskeletal pain patients (chronic low back pain, n=15; fibromyalgia, n=13) and 18 healthy controls underwent fMRI while performing the Monetary Incentive Delay (MID) task. Behavioral and neural responses were compared across groups and correlated against measures of depression (Beck Depression Inventory) and hedonic capacity (Snaith-Hamilton Pleasure Scale).

Results: Compared to controls, patients demonstrated higher anhedonia and depression scores, and a dampening of striatal activation and incentive-related behavioral facilitation (reduction in reaction times) during reward and loss trials of the MID task (p<0.05). In all participants, lower activation of the right striatum during reward trials was correlated with lower incentive-related behavioral facilitation and higher anhedonia scores (p<0.05). Finally, among patients, lower bilateral striatal activation during loss trials was correlated with higher depression scores (p<0.05).

Conclusions: In chronic pain, PA reduction and NA increase are accompanied by striatal hypofunction as measured by the MID task.

1. Introduction

Chronic pain is highly comorbid with mood disorders and is often accompanied by increased negative affect (NA) and decreased positive affect (PA), yielding poorer health-related quality of life and greater clinical burden than either condition alone (Albrecht et al., 2019b; Arnow et al., 2006; Bair et al., 2003; McWilliams et al., 2003). Several lines of evidence support the presence of dysfunctional reward pathways in co-occurring pain and mood alteration. Anhedonia, defined as the loss of pleasure from ordinarily rewarding activities, is a cardinal symptom of major depressive disorder and is commonly associated with chronic pain (Manchikanti et al., 2002). A wealth of animal and human research indicates that pain and anhedonia share common neuroanatomical substrates involving the reward systems (de Heer et al., 2014; Garland et al., 2019; Leknes and Tracey, 2008), such as reduced activity in the striatum and anomalies in reward processing. While...
analgesia and pain relief are inherently rewarding and hedonic events, the experience of relief may differ in patients with chronic pain. For example, patients with fibromyalgia, a disorder with documented alterations in endogenous opioid analgesic activity (Harris et al., 2007), show dampened neural responses to anticipation of pain relief, which might reflect anhedonic response to rewarding stimuli (Loggia et al., 2014). Among the brain regions involved in the processing of rewards, the striatum is among those most consistently shown to be altered in chronic pain patients, including chronic low back pain (Bali et al., 2010, 2012; Berger et al., 2014; Martikainen et al., 2013, 2015), burning mouth syndrome (Hagelberg et al., 2003a, 2003b), and fibromyalgia (Wood et al., 2007b). The striatum is a key structure implicated in the learning of associations between stimuli, actions, and rewards, and motivational modulations of motor behavior (Liljeholm and O’Doherty, 2012), and as such its disruption in chronic pain may reflect different aspects of reward processing, including motivational salience and motor planning (Puglisi-Allegra and Ventura, 2012).

Notably, blunted responsiveness in regions of the reward system, including the striatum, predicts attenuated opioid-induced analgesia even in healthy participants (Vanigasekera et al., 2012). Thus, the investigation of striatal dysfunctions in chronic pain patients may enhance our understanding of the reduced efficacy of opioid treatments of chronic pain, and guide research identifying novel treatment targets that could help ameliorate the global opioid epidemic.

While a growing number of studies implicate striatal neurocircuity in the pathophysiology of mood disorders (Epstein et al., 2006; Keedwell et al., 2005; Pizzagalli et al., 2008, 2009), including when comorbid with pain (Borsook et al., 2007), our knowledge remains limited. In this study, we used the monetary incentive delay (MID) task (Knutson et al., 2000), a validated functional task that probes behavioral and neural responses to monetary rewards and losses, to test the hypothesis that striatal dysfunction co-occurs with mood alterations in chronic pain. The MID task has been used in various clinical conditions linked with anhedonia or NA and alterations in the reward circuitry, such as major depression and substance use disorders (Beck et al., 2009; Knutson et al., 2008). It has been established as a reliable tool to assess incentive-specific behavioral and neurophysiological responses (Knutson et al., 2001b; Knutson and Greer, 2008) and a useful paradigm to investigate psychiatric phenotypes (Knutson and Heinz, 2015). While the MID task contains separate anticipation and outcome/consumption phases, in this study we focused only on the former, due to our specific interest in the striatum. In fact the anticipation phase is a more sensitive probe of striatal function (in healthy subjects, anticipation of reward or loss is ordinarily accompanied by strong activation of striatal regions, which is dampened in several psychiatric conditions such as major depression), whereas the feedback phase of the task tends to recruit more the orbitofrontal and ventromedial prefrontal regions, with lesser striatal engagement (Oldham et al., 2018; Wilson et al., 2018).

2. Methods and materials

2.1. Participants

Twenty-eight patients diagnosed with chronic (>6 months) musculoskeletal pain and 18 healthy, pain-free controls (HC) completed all study procedures. Patients had either chronic low back pain (CLBP; n=15), with and without radicular pain complaints, or fibromyalgia (FM; n=13), a disorder characterized by widespread pain, muscle tenderness, and other symptoms (Wolfe et al., 1990). CLBP was defined as ongoing low back pain for more than 6 months, with a self-reported average pain intensity of at least 3 (on a 0-10 scale) during a typical week for at least half the week. All FM patients met the criteria of the American College of Rheumatology (Wolfe et al., 2011).

Exclusion criteria included any magnetic resonance imaging (MRI) contraindications (e.g., pregnancy, claustrophobia), history of notable medical disorders, illicit drug use confirmed by subjective report and urine drug screening, and routine moderate-to-high use of opioids (>60 mg morphine equivalents). Because participants were simultaneously scanned with the PET radioligand [(11C)PBR28], which binds to the 18 kDa translocator protein (TSPO) (Albrecht et al., 2019a, 2019b; Loggia et al., 2015), we also excluded for the use of benzodiazepines other than alprazolam, lorazepam, and diazepam, due to their documented low affinity for TSPO (Kalk et al., 2013). However, the PET results are beyond the scope of this investigation, which focuses solely on functional MRI (fMRI) responses to the MID task, and will not discussed further here.

The study was conducted at the Athinoula A. Martinos Center for Biomedical Imaging at Massachusetts General Hospital. The participants in this dataset were included in research evaluating the role of neuroinflammation using PET in FM (Albrecht et al., 2019a) and CLBP (Albrecht et al., 2019b; Loggia et al., 2015), which did not report on any task-based fMRI results. All participants were enrolled between 10/30/2015 and 11/29/2017 and provided written informed consents to a protocol approved by the Partners Human Research Committee.

2.2. Procedure

After an initial phone screening, eligible participants completed a 2-hour visit for clinical assessment and training, including a medical history intake and a physical examination.

On a separate day, participants underwent a simultaneous PET/MRI scan in conjunction with the MID task. Periodically throughout the scan, participants rated their pain on a visual analog scale, anchored by 0 ("No pain at all") and 100 ("Most intense pain tolerable"). One during-scan pain rating for one CLBP patient was missing. Participants were compensated for their participation in the study and received instructions that they could earn an additional $17-$22 depending on their task performance (see section 2.3 for details). During the study, all participants (minus one CLBP patient) completed the Beck Depression Inventory (BDI-1A; Beck et al., 1961), which has shown good psychometric properties in individuals with chronic pain (Geisser et al., 1997). In addition, with the exception of 2 FM patients, all participants completed the Snaith-Hamilton Pleasure scale (Snaith et al., 1995), which specifically assesses anhedonia. Using a dimensional approach (Franken et al., 2007), each SHAPS item was scored on a 1-4 scale (1 = "Strongly agree"; 2 = "Agree"; 3 = "Disagree"; 4 = "Strongly disagree"). The scoring proposed in the original publication of the scale (Snaith et al., 1995) was also used, recoding the four response categories in dichotomous categories (i.e., 0 for either "Agree" or "Strongly agree", 1 for either "Disagree" or "Strongly disagree"), but only to classify participants as anhedonic per the original cutoff (score > 2, original scoring).

2.3. MID task

The specific trial structure of the MID task used in this study followed that of previous works (Admon et al., 2017) (see Fig. 1). The task consisted of three runs, each lasting approximately 5 min and containing 24 trials (8 reward trials, 8 loss trials, and 8 no-incentive trials) in a pseudorandomized order, following an initial practice run. At the onset of each trial, the anticipatory visual cue appeared for 0.5s and indicated the potential outcome ("+$", "$-", or "$0" for reward, loss or no-incentive trials, respectively). Following a variable jittered anticipatory period (2.25–3.75s), the participants saw a red target square appear for 0.15s and pressed a key as soon as possible upon seeing the target. A successful trial was defined as button press within the 70th percentile of the participant’s response time (RT) from the immediately preceding run. After a second variable interval (2.4–3.9s), visual feedback (1.25s) indicated the trial outcome. If successful, participants gained money ($1.98–2.32; pseudorandomized) on reward trials and avoided losing money ("no change") on loss trials. Conversely, participants did not gain any money ("no change") on reward trials and lost money ($1.82–2.19; pseudorandomized) on loss trials if their RTs fell outside of the 70th percentile. No-incentive
trials always yielded “no change” feedback. “Wrong moves” (penalty: $2) occurred when participants either pressed the button before the target square appeared or gave no response. There was no feedback on cumulative earning, and a variable interval (1.5–4.5s) separated the trials. An initial calibration run, identical to the task runs but without any feedback, was completed immediately before the first experimental MID run to generate baseline RT calculations. In healthy subjects, anticipation of reward and loss is ordinarily accompanied by strong activation of striatal regions (including the ventral striatum, head of the caudate, and putamen), and faster incentive-related RT speed compared to that in the non-incentivized trials (Oldham et al., 2018; Wilson et al., 2018). In patients with major depression, however, the reward and loss anticipation has been associated with dampened striatal activations, and the cues lose their facilitatory effect on RTs (Pizzagalli et al., 2009).

2.4. Neuroimaging data acquisition and processing

Scans were performed on an integrated PET/MRI scanner consisting of a dedicated brain avalanche photodiode-based PET scanner in the bore of a Siemens 3T Tim Trio MRI (Siemens Corp, Erlangen, Germany) (Kolb et al., 2012). A multi-echo T1-weighted magnetization-prepared rapid acquisition with gradient echo (MEMPRAGE) volume was acquired prior to tracer injection (TR/TE1/TE2/TE3/TE4 = 2530/1.64/3.5/5.3/7.22 ms, flip angle = 7°, voxel size = 1mm isotropic) for anatomical localization and spatial normalization of the imaging data (and generation of attenuation correction maps (Tzquierdo-Garcia et al., 2014)).

The participants underwent four ~5-min BOLD fMRI scans, each corresponding to one initial calibration and three experimental MID task runs (TR/TE = 2s/30ms, flip angle = 90°, voxel size = 3.1 × 3.1 × 3mm, 37 slices, 142 vol). The calibration imaging data were collected to generate the same acoustic and physical environment the participants would experience during the experimental MID runs, thus providing an accurate calibration of the RTs to be used in the first run. This run also allowed us to probe the participants’ ability to comfortably complete the task. Although patients reported some pain during scanning, no participant reported discomfort significant enough to interfere with task completion. fMRI data were pre-processed and analyzed using FSL (FMRIB’s Software Library, http://www.fmrib.ox.ac.uk/fsl/), AFNI (Automated Functional NeuroImaging, http://afni.nimh.nih.gov/afni), and FreeSurfer (http://surfer.nmr.mgh.harvard.edu/) software packages. Data were corrected for slice-timing, motion, and B0 field inhomogeneities. Subsequently performed was brain extraction, coregistration to the MEMPRAGE, spatial smoothing with a 6mm Gaussian kernel, and nonlinear registration to Montreal Neurological Institute (MNI) standard space.

2.5. Statistical analysis

All group effects in demographic, behavioral, and imaging outcomes were first tested by combining all patients into a singular “Pain” group, to maximize our statistical power (primary analysis). While CLBP and FM may have different etiology and pathophysiology, both are characterized by musculoskeletal pain. Combining data is justified in this context because 1) this project focuses on the impact of chronic pain on reward processing, and not on the somatic aspects of pain, and 2) disruptions in reward processing have been similarly hypothesized in multiple chronic pain disorders, including FM and CLBP (Berger et al., 2014; Loggia et al., 2014). To evaluate whether any effects observed in the primary analyses might be driven by one subgroup, follow-up secondary group analyses were performed, entering CLBP and FM separately in the statistical models.

2.5.1. Demographic and behavioral outcomes

Group differences in age and questionnaire scores were tested using an analysis of variance (ANOVA) or unpaired t-tests as applicable, while group differences in sex distribution were tested using a chi-square test.

RTs were averaged separately per trial type (reward, loss, no-incentive) after removal of “wrong moves,” as previously described (Pizzagalli et al., 2009). One FM participant was excluded from the behavioral analysis due to a calibration error where the 50th percentile was used instead of the 70th. While this calibration may have slightly affected the difficulty of the task, the same participant was however included in the imaging analyses, as it was presumed that the subjective experience of reward/loss anticipation (as opposed to the performance of the button press itself) would not have been meaningfully affected by a minor change in the difficulty of the task. Moreover, repeating the fMRI analyses after excluding the same patient still yielded significant group effects (see Results).

We first evaluated the RTs using analyses of covariance (ANCOVA) with Group and Trial as factors and age and sex as covariates of no interest. The aim of these analyses was to test for a significant Trial type effect on RTs, as a marker for a successful experimental manipulation (shorter RT in reward/loss trials indicating behavioral facilitation and thus a correct implementation of the MID task (Pizzagalli et al., 2009)). Although these analyses enabled observation of Group effects or Group × Trial interactions, the primary evaluation of these effects utilized a different ANCOVA model, in which the no-incentive RT was added as a covariate rather than a level in the factor Trial. This model produced maximal sensitivity in assessing the expected incentive-related RT reduction, by correcting for the general RT differences across groups. While age and sex were not statistically different in the primary analyses comparing all pain patients with controls, there were marginal or significant group differences when the pain groups were evaluated separately (see Supplementary Materials). Therefore, these variables were used as covariates in all group analyses.
Significant effects and interactions were analyzed using Tukey’s HSD. Pearson’s correlations were employed to investigate the relationship between behavioral measures (RT difference scores, BDI scores, SHAPS scores) in the patients. Demographic and behavioral data were analyzed using Statistica 13 (StatSoft).

2.5.2. Neuroimaging data

In the first-level fMRI analysis, the anticipatory time period (including visual cue and anticipatory period prior to target presentation), button press, feedback/outcome (gain, loss, no change, wrong move) and six head motion parameters (3 translations, 3 rotations) were modeled as regressors per trial type. Parameter estimates for each contrast of interest (“reward > no-incentive” and “loss > no-incentive”) were computed for each run, then averaged for each subject using a fixed-effect analysis (3 runs for all subjects, except for one CLBP and two FM participants, for whom technical and calibration issues prevented the execution of the third run). Mixed-effect analyses (FSL’s FLAME1) were used to create group maps for each group, an omnibus group map generated by averaging all three groups (in order to functionally define region-of-interests (ROI); see below) and to compare HC against pain patients, both separately and combined, with sex and age as regressors of no interest. Results were corrected for multiple comparisons using a cluster-forming threshold of $z \geq 3.1$ and a corrected cluster significance threshold of $p < 0.05$. Images were visualized using FreeSurfer's Freeview tool (https://surfer.nmr.mgh.harvard.edu/fswiki/FreeviewGuide).

In addition to the whole-brain voxelwise analyses, we pursued ROI analyses using functionally-defined striatal ROIs. These ROIs were obtained from a conjunction analysis of the entire sample using the omnibus activation maps for “reward > no-incentive” and “loss > no-incentive” anticipation contrasts applying easythresh_conj (https://warwick.ac.uk/fac/sci/statistics/staff/academic-research/nichols/scripts/fsl/easythresh_conj.sh). The resulting regions were intersected with the binarized Oxford-Imanova striatal structural atlas, then lateralized to define functionally-localized striatal ROIs. Note that this strategy was pursued instead of running masked voxelwise analyses using anatomically defined striatal masks because the striatal activations typically observed during the MID anticipatory phase do not completely overlap with the available anatomical striatal masks (e.g., they tend to load more on ventral striatum and ventromedial aspects of caudate and putamen, and typically extend slightly outside the anatomical boundary of striatal labels (Oldham et al., 2018)).

From these functionally defined striatal ROIs, we extracted the mean contrast of parameter estimates (COPEs), or beta weights. Applying statistical models analogous to those in the RT analyses, we first evaluated the striatal COPEs with ANCOVAs, setting Group and Trial as factors and statistical models analogous to those in the RT analyses, we compare HC against pain patients, both separately and combined, with sex and age as regressors of no interest. Results were corrected for multiple comparisons using a cluster-forming threshold of $z \geq 3.1$ and a corrected cluster significance threshold of $p < 0.05$. Images were visualized using FreeSurfer’s Freeview tool (https://surfer.nmr.mgh.harvard.edu/fswiki/FreeviewGuide).

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3. Results

3.1. Demographics

Demographic characteristics of the groups are presented in Table S1. No differences in age ($F(1,44) = 1.72$, $p = 0.2$) or sex ($p = 0.14$) were observed between the pain group and healthy controls. When controlling for sex and age, patients had significantly higher BDI scores than controls ($F(1,41) = 10.76$, $p < 0.01$) (Fig. 2A). The average BDI score for patients was 9.15, although 12 patients (4 CLBP and 8 FM patients) had a score of at least 10 (the threshold for mild-to-moderate depression for the version used in the current study (Beck et al., 1988)). Pain patients also had significantly higher SHAPS scores (i.e., higher anhedonia) ($F(1,40) = 7.18$, $p = 0.01$) (Fig. 2B). Fig. 2C and D display data separating the pain group into CLBP and FM subgroups (for more information, see Supplementary Materials). Using the original SHAPS scoring, 9 patients (2 CLBP and 6 FM) scored above a 2, indicating anhedonia (Snaith et al., 1995). BDI and SHAPS scores were significantly correlated in the pain group ($r = 0.53$, $p < 0.01$) (Fig. 2E). Positive associations between pain ratings and BDI ($r = 0.31$) and SHAPS ($r = 0.28$) were observed but did not reach statistical significance ($p s \geq 0.12$).

3.2. Behavioral responses to the MID task

A significant effect of Trial type emerged using absolute RTs ($F(2,82) = 4.27$, $p = 0.02$). Post-hoc pairwise comparisons clarified that all participants were faster during reward ($p = 0.0001$) and loss trials ($p = 0.0002$) than during no-incentive trials, indicating that the task elicited the intended effects. RTs for reward and loss trials did not differ ($p = 0.19$). The pain group was significantly slower than controls ($F(1,41) = 6.0$, $p = 0.02$), irrespective of trial type (Group × Type interaction: $F(2,82) = 1.13$, $p = 0.33$) (Fig. 3A).

When controlling for differences in RTs during no-incentive trials, the RTs from reward or loss trials were significantly longer in patients ($F(1,40) = 5.48$, $p = 0.02$), indicating that the relative facilitatory effect of reward or loss cues on RTs was attenuated in patients. No significant Group × Trial interaction emerged from the relative RTs ($F(1,40) = 0.34$, $p = 0.56$), suggesting similar effects from loss and reward trials (Fig. 3B). Neither BDI nor SHAPS correlated with RT difference scores ($p s > 0.33$). Stratifying the pain patients into CLBP and FM subgroups (Fig. 3C and D) revealed a gradient, where CLBP patients performed intermediately to the HC and FM groups, paralleling the pattern observed in the depression and anhedonia scores (Fig. 2C and D) (for more information, see Supplementary Materials).

Overall, our results are consistent with the hypothesis that chronic pain patients display dampened behavioral response to reward and loss cues, even after correcting for the general slowness of pain patients.

3.3. Neural responses to the MID task

In a whole-brain voxelwise analysis of healthy controls, both reward and loss trials were associated, as expected, with significant activation of the striatum (in the caudate nucleus: right-sided for reward, bilateral for loss; right nucleus accumbens, and right putamen) and the supplementary and pre-supplementary motor areas (SMA/pre-SMA), compared to the no-incentive trials. Chronic pain patients, in contrast, did not demonstrate any significant activations (see Fig. 4A; Table 1). Fig. 4B displays the brain responses extracted from the striatum of both pain and HC groups. While the direct group comparison of these contrast maps was not significant in the whole-brain voxelwise analyses, ROI analyses (Fig. 4C) showed that when controlling for anticipatory brain responses in the no-incentive trials, the right striatal responses in the reward and loss trials
were significantly smaller in patients than in controls ($F(1,41) = 7.21, p = 0.01$). Again, no significant Group × Trial interaction was observed for these contrasts, ($F(1,41) = 0.29, p > 0.05$), signifying that anticipatory striatal hypofunction was similarly observed in reward and loss trials. When the analyses were repeated after the exclusion of the FM patient with the RT calibration error, the Group effect remained significant ($F(1,40) = 8.34; p = 0.006$) and the Group × Trial interaction remained not significant ($F(1,40) = 0.23, p > 0.05$). Similar to the behavioral results, a gradient also appeared in the striatal responses, with the CLBP patients displaying intermediate activation between the HC and FM groups (Fig. 4D and E; for more information, see Supplementary Materials).

Overall, the imaging results demonstrate a dampening in the brain responses to anticipation of reward and loss, paralleling the patterns observed in the behavioral results.

3.4. Association between imaging and behavioral/clinical measures

Across all participants, lower activation of the right striatum during anticipation of reward was significantly correlated with slower incentive-related behavioral facilitation ($r = 0.37, p = 0.013$; Fig. 5A), and higher SHAPS scores ($r = -0.31, p = 0.037$; Fig. 5B). The same correlations did not reach statistical significance for the right striatum during loss trials ($p \geq 0.11$), or for the left striatum in either reward or loss trials ($p \geq 0.09$).
In patients, diminished striatal activation was correlated with higher BDI scores during loss trials (right striatum: $r = -0.42$, $p = 0.026$; left striatum: $r = -0.42$, $p = 0.027$) (Fig. 6). The correlation between striatal activation and BDI scores was not significant in the reward trials, for both right ($r = -0.37$, $p = 0.056$) and left striatum ($r = -0.29$, $p = 0.14$). Pain ratings also were not significantly correlated with either reaction times ($-0.101 < r < -0.018$, $p > 0.64$) or striatal responses ($-0.185 < r < -0.085$, $p > 0.40$) in partial correlation analyses correcting for sex and age.

For results from the analyses separating patients into CLBP and FM subgroups, see Supplementary Materials.

4. Discussion

In the current study, we report that chronic pain patients, when compared with healthy controls, show 1) increased negative affect and anhedonia, 2) smaller incentive-related behavioral facilitation from reward and loss cues and 3) lower anticipatory activation in the striatum. Striatal hypofunction across groups was associated with slower behavioral responses and greater levels of anhedonia, while in patients, higher depression scores were correlated with low bilateral striatal activation during loss trials.

The absence of robust striatal activations in chronic pain patients supports the general view that dysfunctions within the mesolimbic dopamine pathways play a role in the pathogenesis of chronic pain (Baliki et al., 2012), and its comorbidity with mood alterations (Schwartz et al., 2014). Notably, the most pronounced activation difference between pain patients and healthy controls was observed in the caudate nucleus, a region of the striatum known to control motivation (Delgado et al., 2004) and is highly innervated by dopaminergic neurons whose projections are sent from the substantia nigra pars compacta. We also observed activity in the nucleus accumbens, another striatal region with high density of dopaminergic neurons and whose activity is consistently associated with anticipation of reward as well as motivational salience and reward-oriented motor planning as probed by the MID task (Knutson et al., 2001a). While our understanding of the neurophysiological mechanisms underlying negative affect and anhedonia in chronic pain remains incomplete, preclinical pain models provide some clues. For instance, the excitatory synaptic transmission of D2 dopamine receptor expressing medium spiny neurons in the nucleus accumbens show a galanin receptor 1-mediated depression, implicating alterations in the indirect pathway of the basal ganglia (Schwartz et al., 2014). Human imaging studies in FM patients have shown hypoactivation in the mesolimbic dopamine systems (Loggia et al., 2014), low D2 dopamine receptor binding potential in the striatum (Wood et al., 2007b), and low presynaptic dopamine activity (Wood et al., 2007a), while CLBP patients have also demonstrated low striatal dopamine receptor binding potential (Martikainen et al., 2015). Even though the research is mixed in relating particular striatal substrates and subnuclei with specific dopaminergic functions and behavioral manifestations, converging lines of evidence support the role of dopaminergic alterations as mediators of anhedonia and NA comorbid with chronic pain (Finan and Smith, 2013; Jarcho et al., 2012; Scott et al., 2006; Taylor et al., 2016; Tiemann et al., 2014).

Moreover, in healthy participants, ventral striatum activation positively covary with the activation of the periaqueductal gray and the pain reduction induced by positive mood change (Villemure et al., 2012), further providing a link between striatum, pain, and affect.

Although a previous study probed reward processing in FM patients using the MID task (Martucci et al., 2018), the current study differs from the existing study in important ways. In particular, the former study did
not observe any significant group differences in RTs or striatal activation. A possible explanation for this discrepancy is the fact that while our study employed an anticipatory period of variable duration, the previous study used a fixed duration. The latter experimental choice could have rendered the task performance more dependent on the ability to correctly estimate the duration of the anticipatory period, than on the actual identification of the target itself. In addition, our use of shorter runs (three 5-min 24-trial blocks, vs. 90 trials in 2 blocks) may have helped minimize attention fatigue, and the use of a functionally-defined striatal mask (as opposed to anatomically defined ROIs) may have afforded greater sensitivity to detect group differences. Our study also included both FM and CLBP patients, which enabled us to observe a HC > CLBP > FM gradient in all principal outcomes evaluated. The dampening of the behavioral and striatal responses to reward and loss cues, and the levels of negative affect and anhedonia were all greatest in FM patients, followed by CLBP patients. This gradient provides further support for the
Among the limitations of this study, we note a relatively small sample size, particularly within each pain subgroup (CLBP and FM). Because in this study it was assumed that striatal hypofunction was a general feature of chronic pain, irrespective of specific etiology and clinical characteristics of the patients evaluated, we elected to combine data of patients with different pain disorders. While we feel that the focus on reward/motivational (and not somatosensory) processing justified this approach, it is certainly possible that some nuanced differences may exist in how striatal physiology is affected in different pain disorders. Future studies with larger sample sizes will be necessary to fully evaluate how certain features of chronic pain (e.g., etiology, degree of “centralization”) may relate to striatal dysfunction. Similarly, future studies will need to compare patients with or without pain, but comparable levels of anhedonia, in order to evaluate the specific contributions of pain state to the striatal alterations reported here. In addition, the increase in reaction times during the MID task was not selective for the reward and loss trials but was also observed during the neutral trials. This overall trend in the patients is suggestive of a generalized reduction in motor processing speed, which could be due to a number of factors including an effect of pain on cognition (Seminowicz et al., 2004), or the use of pain medications (Hienz et al., 2001). However, it should be noted that the group differences in the incentive-related trials remained significant even after correcting for differences in reaction times in the neutral trials, indicating

Table 1
MNI coordinates and cluster size from the voxelwise analyses.

<table>
<thead>
<tr>
<th>Cluster Size (# voxels)</th>
<th>Cluster size (p value)</th>
<th>Z</th>
<th>MNI x (mm)</th>
<th>MNI y (mm)</th>
<th>MNI z (mm)</th>
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<td>5.02</td>
<td>6</td>
<td>12</td>
<td>4</td>
<td>R head of the caudate nucleus</td>
</tr>
<tr>
<td></td>
<td></td>
<td>4.42</td>
<td>8</td>
<td>16</td>
<td>–2</td>
<td>R nucleus accumbens</td>
</tr>
<tr>
<td></td>
<td></td>
<td>3.79</td>
<td>14</td>
<td>12</td>
<td>–6</td>
<td>R putamen</td>
</tr>
<tr>
<td></td>
<td>254 0.029</td>
<td>4.03</td>
<td>0</td>
<td>–2</td>
<td>66</td>
<td>SMA/pre-SMA</td>
</tr>
</tbody>
</table>

Loss minus no-incentive Controls

| Reward minus no-incentive Controls | 768 1.26E-05 | 5.02 | 8          | 8          | 4          | R head of the caudate nucleus |
|                          |                      | 4.15 | –8         | 8          | 2          | L head of the caudate nucleus |
|                          |                      | 3.69 | 10         | 12         | –4         | R nucleus accumbens       |
|                          |                      | 3.38 | 14         | 10         | –6         | R putamen                |
|                          | 270 0.0118 | 3.92 | 0          | –4         | 66         | SMA/pre-SMA              |

Fig. 5. Correlation between imaging and behavioral measures in reward trials. A. Change in reward COPEs (vs. no-incentive COPE) of the functionally-defined right striatum correlated with change in reward-trial RTs. B. Change in reward COPEs (vs. no-incentive COPE) of the functionally-defined right striatum correlated with BDI scores.

Fig. 6. Correlation between imaging and behavioral measure in loss trials. Change in loss COPEs (vs. no-incentive COPE) of the functionally-defined right striatum correlated with BDI scores.
With the ever-growing demand for effective treatments of chronic pain, the need to advance our understanding of chronic pain is more critical than ever. Unraveling the interactions between chronic pain and the nervous systems could lend a new insight into identifying better predictors and novel treatment targets for pain and its comorbidities.

Author contribution

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Vitaly Napadow: Writing - Review & Editing.
Diego A. Pizzagalli: Conceptualization, Methodology, Resources, Writing - Review & Editing.
Marco L. Loggia: Funding acquisition, Conceptualization, Methodology, Formal analysis, Writing - Reviewing and Editing.

Declaration of competing interest

MK, IM, DSA, RA, AT-C, CB, EP, PK, RRE, AS, and VN have no biomedical financial interests or potential conflict of interest to report. Over the past three years, M.L.L. received consulting fees from Shionogi Inc for activities unrelated to the current study. D.A.P received consulting fees from Akili Interactive Labs, BlackThorn Therapeutics, Boehringer Ingelheim, CompassPathway, Posit Science and Takeda Pharmaceuticals USA and an honorarium from Alkermes for activities unrelated to the current study.

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Appendix A. Supplementary data

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References


