



Racial disparities in health among nonpoor African Americans and Hispanics: The role of acute and chronic discrimination



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ABSTRACT

Racial disparities in health tend to be more pronounced at the upper ends of the socioeconomic (SES) spectrum. Despite having access to above average social and economic resources, nonpoor African Americans and Latinos report significantly worse health compared to nonpoor Whites. We combine data from the parents and children of the 1979 National Longitudinal Survey of Youth (NLSY79) to address two specific research aims. First, we generate longitudinal SES trajectories over a 33-year period to estimate the extent to which socioeconomic mobility is associated with exposure to discrimination (acute and chronic) across different racial/ethnic groups (nonHispanic Whites, nonHispanic Blacks, and Hispanics). Then we determine if the disparate relationship between SES and self-rated health across these groups can be accounted for by more frequent exposure to unfair treatment. For Whites, moderate income gains over time result in significantly less exposure to both acute and chronic discrimination. Upwardly mobile African Americans and Hispanics, however, were significantly *more* likely to experience acute and chronic discrimination, respectively, than their socioeconomically stable counterparts. We also find that differential exposure to unfair treatment explains a substantial proportion of the Black/White, but not the Hispanic/White, gap in self-rated health among this nationally representative sample of upwardly mobile young adults. The current study adds to the debate that the shape of the SES/health gradient differs, in important ways, across race and provides empirical support for the diminishing health returns hypothesis for racial/ethnic minorities.

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1. Introduction

In the United States, racial disparities in health are well established across a diverse set of outcomes (Williams, 2012). These inequalities are primarily the result of excessive rates of chronic disease among racial minorities (Geronimus et al., 2011; Harper et al., 2012; Murray et al., 2006). NonHispanic Blacks face rates of hypertension, diabetes, and obesity that are 25%, 49%, and 59% higher, respectively, than those found among nonHispanic Whites, while Hispanics experience rates of diabetes and obesity that are 25% and 20% higher, respectively than their nonHispanic White

counterparts (NCHS, 2016). Despite the amount of attention that has been paid to documenting these stark racial disparities in morbidity and mortality, many questions remain regarding the social and economic conditions that give rise to such unequal chances of an outcome that, quite literally, determines who lives and who dies.

One of the strongest predictors of health is socioeconomic status (SES) – so much so, that it is considered to be a fundamental cause of health disparities (Phelan et al., 2010). The association between SES and health has been shown to be positive, linear, and monotonic, such that with each additional increase in socioeconomic standing, we expect an individual's health to improve in a step-wise fashion (Adler and Rehkopf, 2008). This SES/health gradient has been noted across a wide range of settings as well as historical

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time periods (Adler and Stewart, 2010; Braveman et al., 2010; Elo, 2009). However, among some racial/ethnic groups the association between SES and health is not as straightforward as was once thought. For example, Black/White disparities in health are often more pronounced among high as opposed to low SES subpopulations and upward mobility does not seem to bring similar improvements in health for African Americans compared to Whites (Colen et al., 2006; Colen, 2011; Farmer and Ferraro, 2005; Williams and Sternthal, 2010). For Hispanics, particularly those born outside the U.S. (Williams et al., 2010), the association between SES and health is so tenuous that investigators have coined the term Hispanic health paradox to describe this particularly unexpected finding and spent much of the previous three decades trying to determine if the Hispanic health paradox is “real” or primarily the result of multiple data artifacts (Markides and Eschbach, 2005; Ruiz et al., 2013; Smith and Bradshaw, 2006; Vega et al., 2009).

Clearly more research is needed to understand, on a deeper level, why some racial/ethnic groups stand poised to benefit, with respect to their health, from gains in their socioeconomic standing while others do not experience similar positive effects. One potential explanation for this finding is that for racial minorities, ascending the socioeconomic hierarchy does not guarantee fewer encounters with discrimination, or unfair treatment (For the purposes of this paper, the terms discrimination and unfair treatment will be used interchangeably. Unless otherwise specified, they refer to forms of discrimination, or unfair treatment, that are unattributed to any specific social characteristic such as race, ethnicity, gender, age, etc.). In fact, some studies have revealed the opposite – that individuals with higher levels of SES tend to report more instances of interpersonal discrimination (Everson-Rose et al., 2015) or that the association between unfair treatment and health is significantly more pronounced among higher vs. lower SES groups (Fuller-Rowell et al., 2012; Hudson et al., 2012, 2013).

The proposed study seeks to examine if and how exposure to discrimination can help to explain why nonpoor African Americans and Latinos tend to report worse health compared to nonpoor Whites, despite typically having access to above average socioeconomic resources via educational attainment, occupational prestige, or family income. We use data from two cohorts of the 1979 National Longitudinal Survey of Youth (NLSY79) to address two specific research aims. First, we estimate the extent to which longitudinal SES trajectories are associated with exposure to discrimination across different racial/ethnic groups (nonHispanic Whites, nonHispanic Blacks, and Hispanics). Then, we determine if the disparate relationship between SES and self-rated health across these racial/ethnic groups can be accounted for by more frequent exposure to racial discrimination.

1.1. SES and health

The relationship between SES and health is one of the most consistent and persistent findings within the health disparities literature (Adler and Rehkopf, 2008). It has been shown across a wide range of outcomes, SES indicators, subpopulations, and time periods that as SES increases, physical wellbeing also improves (Braveman et al., 2010; Elo, 2009; Phelan et al., 2010). Although there remains some debate regarding whether the mechanisms through which SES influences health are largely a result of differential access to material resources or primarily psychosocial in nature (Demakakos et al., 2008; Link et al. 2008, 2013; Lynch et al., 2000; Marmot, 2004; Matthews et al., 2010; Seeman et al., 2014), there is overwhelming evidence to suggest that SES is a fundamental cause of population health disparities (Clouston et al., 2016; Phelan et al., 2010).

And yet, despite the seemingly ubiquitous nature of this

association, there is a growing body of evidence to suggest that the association between SES and health depends on race. Some studies note the existence of diminishing returns to SES among African Americans (Colen et al., 2006; Geronimus et al., 2006; Farmer and Ferraro, 2005; Hudson et al., 2012, 2013; Kahn and Fazio, 2005). Not only does it appear that the impact of SES on general indicators of physical wellbeing, such as self-rated health and functional impairment, might be muted for African Americans compared to Whites (Kahn and Fazio, 2005; Farmer and Ferraro, 2005), newer research suggests that these differences might occur on a deeper physiological or cellular level. For example, Geronimus et al. (2006, 2015) demonstrate that allostatic load scores and telomere length are similar for working-age poor and nonpoor Blacks but significantly different for their White counterparts. Diminishing health returns to SES for African Americans is apparent when fluctuations in socioeconomic resources are taken into account across generations as well as within a single generation. Colen et al. (2006) report significant reduction in the odds of low birthweight following increases in family income over time only for White but not Black NLSY79 respondents who spent their childhoods in or near poverty.

Among Hispanics, the relationship between SES and health is less apparent – so much so, that researchers frequently employ the term, Hispanic health paradox, to refer to the unexpected population health patterning displayed by this racial/ethnic group, which reports relatively low levels of SES but tends to experience better than average health outcomes (Markides and Eschbach, 2005; Ruiz et al., 2013; Smith and Bradshaw, 2006; Vega et al., 2009). This pattern is more consistently noted when examining rates of death (Hummer et al., 2007) as opposed to specific disease outcomes, health behaviors, or biological risk profiles (Abraido-Lanza et al., 2005; Bates et al., 2008; Crimmins et al., 2007) as well as among certain Hispanic subgroups, particularly Mexican-Americans (Markides and Eschbach, 2005) and recent immigrants to the U.S. (Williams et al., 2010). It is thought that better adherence to healthier diets, lower rates of smoking, and higher levels of social support could be protecting the health of Latinos in the U.S. despite their low SES, at least during the first decade of time spent in the U.S. before acculturation becomes more complete (Abraido-Lanza et al., 2005). Although much debate remains regarding the underlying factors that are driving this quizzical finding, with some investigators arguing that the HHP is not even real but simply the result of data artifacts (Palloni and Arias, 2004; Smith and Bradshaw, 2006), there remains ample evidence to suggest that the association between SES and health is less pronounced among Hispanics than Whites.

1.2. Discrimination and health

Exposure to discrimination has been posited as an explanation, at least in part, for the stark and consistent racial disparities in health. Despite the fact that for most individuals experiences of unfair treatment remain a relatively rare event with the majority of respondents reporting that they “never,” “rarely,” or “sometimes” encounter interpersonal discrimination, a rapidly growing literature links experiences of unfair treatment to subsequent morbidity and even mortality (Lewis et al., 2015; Pascoe and Smart Richman, 2009). There is some evidence that the association between unfair treatment and wellbeing might be more pronounced for mental as opposed to physical health outcomes (Lewis et al., 2015; Paradies et al., 2015; Schmitt et al., 2014; Williams and Mohammed, 2009) although a recent meta-analysis concluded that this difference is not statistically significant (Pascoe and Smart Richman, 2009).

Discrimination has been assessed using a number of different indicators, some of which capture unattributed unfair treatment while others identify unfair treatment that the respondent

specifically ascribes to a sociodemographic characteristic such as race, ethnicity, sex, etc. (Bastos et al., 2010). There is not consensus in the literature concerning which approach provides a more valid measure of interpersonal discrimination. Some investigators argue that in order to more accurately estimate the effects, health and otherwise, of racial discrimination, respondents should identify the underlying reason for reported unfair treatment (Krieger, 2012), while other researchers note that it is often difficult and possibly disadvantageous for a respondent to be required to attribute discrimination to a particular sociodemographic characteristic (Lewis et al., 2015; Williams, 2012). This might be especially problematic for individuals who occupy more than one disadvantaged social status, such as women of color, for whom separating out the underlying source of unfair treatment might be all but impossible. Of particular importance for the current study, there appears to be little difference in the impacts of racial, as opposed to other forms, of discrimination on health (Williams and Mohammed, 2009) suggesting that similar physiological reactions underlie the human response to unfair treatment, regardless of the underlying cause. Moreover, general (unattributed) as opposed to specific (attributed) measures of discrimination tend to be just as or more strongly associated with a variety of negative health outcomes including coronary artery calcification (Lewis et al., 2006), hypertension (Roberts et al., 2008), and cigarette smoking (Chae et al., 2008).

The two most widely used discrimination measures, particularly among studies that explore its impact on health, are Williams' major events discrimination scale (MEDS) and everyday discrimination scale (EDS) as well as Krieger's experiences of discrimination scale. The MEDS and EDS, which are used in the current study, invite respondents to report instances of unfair treatment in a two-step process, the first of which inquires about the frequency of discriminatory occurrences while the second asks about the sociodemographic attribute likely to be driving this interpersonal interaction. On the other hand, Krieger's experiences of discrimination scale requires that respondents report instances of discrimination that have happened specifically as a result of their race; thus, study participants must simultaneously recall an event while attributing the cause of this event to an underlying characteristic (Krieger, 2012; Lewis et al., 2015). Although the psychometric properties of both measures are convincing (Bastos et al., 2010), there are concerns that a one-step approach leads to artificially inflated reports of discrimination primarily due to interviewer effects and response bias (Lewis et al., 2015; Williams and Mohammed, 2009) while a two-step approach captures unfair treatment that is not specifically attributed to one's race (Krieger, 2012).

The biological pathway through which discrimination is thought to negatively influence physical wellbeing is the human stress response. The stress process model has long served as a powerful theoretical framework through which to gain a deeper understanding of population health patterns and inequalities (Pearlin, 1989; Turner, 2009). At its core, this theory emphasizes the interaction between exposures to psychosocial stressors and access to coping resources in the production of health and illness (Thoits, 2010). Thus, holding the number or efficacy of coping resources constant, we expect individuals who encounter more frequent or more challenging stressors to have worse health outcomes than individuals facing fewer or less challenging stressors (Schwartz and Meyer, 2010). In more recent years, the stress model has been expanded to place a greater emphasis on macro-level social conditions, such as the role of social statuses (race, sex, etc.) and neighborhood conditions, that are likely to unequally pattern exposure to psychosocial stressors across population subgroups (Aneshensel, 2009).

It is thought that by interpreting unfair treatment as a psychosocial threat, an individual's stress response system will be activated - most notably, via his/her HPA axis - thus releasing a potent cocktail of hormones and biochemical (e.g. cortisol and epinephrine) that, in turn, trigger a cascade of physiological changes (e.g. increased heart rate, blood pressure, and respiration) specifically produced to deal with the perceived threat (Brondolo et al., 2008; Cohen et al., 2006; Cunningham et al., 2012; Friedman et al., 2009; Fuller-Rowell et al., 2012). As the concept of allostatic load predicts, the worst health outcomes will be seen among those who have this stress response repeatedly activated over time or who cannot turn off this physiological cascade once it is initiated (Geronimus et al., 2006; McEwen, 1998; Seeman et al., 2001). Thus, with respect to the production of racial disparities in health, the chronicity rather than the severity of exposure to discrimination is often considered to be stronger predictor of worse outcomes, either across subgroups or over time. This is supported by a growing body of empirical evidence which shows that everyday discrimination is more consistently associated with higher rates of morbidity and mortality than acute instances of unfair treatment (Paradies, 2006; Williams and Mohammed, 2009). This finding echoes similar results from the broader stress literature that reveal chronic, as opposed to acute, stressors tend to more predictive of suboptimal health outcomes (Cohen et al., 1997).

Although the link between discrimination and health is most apparent for cardiovascular outcomes, particularly those that have been studied within a laboratory setting or using continuous, real-time monitoring (Brondolo et al., 2008), they remain somewhat mixed and require further investigation and replication before definitive conclusions can be drawn (Chae et al., 2010; Lewis et al., 2006, 2009). Greater exposure to unfair treatment has been demonstrated to result in significantly higher daytime blood pressure readings as well as worse nocturnal blood pressure recovery patterns, more rapid heart rates, and elevated cortisol secretions (Smart Richman et al., 2010; Tomfohr et al., 2010; Zeiders et al., 2014). These findings are particularly notable since this type of physiological reactivity is a key pathway through which stressful environmental stimuli is thought to impact subsequent health outcomes and, in particular, has been implicated in the emergence of hypertension and other negative cardiovascular outcomes (D'Agostino et al., 2008; Din-Dzietham et al., 2004). A key question for researchers in this area is the extent to which results from experimental studies can be generalized to more diverse populations.

The empirical evidence linking discrimination to health generated from community-based or observational research is less straightforward but, nonetheless, provides preliminary support for the notion that exposure to unfair treatment is likely to negatively impact physical wellbeing. Several studies reveal that the frequency of discrimination is predictive of worse health across a wide range of physical health outcomes including all-cause mortality (Barnes et al., 2008), chronic health conditions (Gee et al., 2007; Ryan et al., 2006), self-reported health (Borrell et al., 2006; Harris et al., 2006; Schulz et al., 2006), asthma incidence (Coogan et al., 2014), coronary artery calcification (Lewis et al., 2006), obesity (Cozier et al., 2014; Hunte, 2011; Hunte and Williams, 2009), low birthweight and preterm birth (Lauderdale, 2006; Mustillo et al., 2004), breast cancer incidence (Taylor et al., 2007), and uterine fibroids (Wise et al., 2007). Exposure to interpersonal discrimination has also been implicated in important subclinical physiological processes that are thought to be a marker of accelerated aging such as higher allostatic load scores (Brody et al., 2014; Ong et al., 2017) and shorter telomeres (Chae et al., 2014). However, these findings are not consistent across all outcomes or subpopulations. For example, the association between discrimination and health

appears more pronounced for mental as opposed to physical outcomes (Williams and Mohammed, 2009) and somewhat quizzically is sometimes stronger among Whites than nonWhites (Hunte and Williams, 2009).

The extent to which exposure to unfair treatment is a key mechanism through which racial disparities in health are produced remains unclear (Williams and Mohammed, 2009). A few notable studies provide preliminary evidence that perceived discrimination is, at the very least, one reason why the Black/White gap in health outcomes remains relatively unchanged despite an ever-increasing average life expectancy, widespread medical advancements, and more inclusive health care coverage. Mustillo et al. (2004) demonstrate that perceived racial discrimination explains as much or more of the Black/White disparity in poor birth outcomes (low birthweight and preterm birth) as maternal education, income, cigarette smoking, alcohol consumption, and depressive symptomatology *taken together*. These findings are further supported by research that found women of Arab descent living in California were significantly more likely to give birth to a low birthweight or preterm baby in the six-month period following September 11th, 2001 than in the six-month period before this date (Lauderdale, 2006). This result is particularly noteworthy since (1) Arab and Arab American women typically have very healthy pregnancies and consistently low rates of LBW and PTB and (2) similar increases were not noted among any other racial/ethnic group during the same time period. Furthermore, Tomfohr et al. (2010) report that encounters of everyday discrimination fully accounted for differences in nocturnal blood pressure dipping between African Americans and Whites. There is also mounting evidence that differential exposure to unfair treatment across race is driving important racial/ethnic inequalities in both sleep quantity and quality (Tomfohr et al., 2012).

The extant literature concerning discrimination and health suffers from some notable shortcomings that the current study was specifically designed to address. First, many existing studies rely on small or convenience samples that are not nationally representative. While this might help to understand how the health of a specific group is impacted by unfair treatment, these findings should not be used to generalize to the U.S. population as a whole. Second, few studies are specifically designed to determine the extent to which exposure to discrimination can help to explain the stark racial disparities in health that currently characterize the American social landscape. Third, the way in which unfair treatment is captured differs substantially across studies, with some investigators relying on indicators with unproven psychometric properties. Fourth, much of what we know about the association between discrimination and health has been derived from cross-sectional research, thus our ability to discern the true causal ordering that undergirds this statistical relationship remains methodologically hampered.

Most notably, few extant studies examine how encounters with discrimination differ across both race and SES, and an even smaller number then link these unequal exposures to subsequent health outcomes. Among a sample of African Americans, Hudson et al. (2012) demonstrate that more frequent encounters with unfair treatment predict an increased risk of depression for high SES men but not their low SES counterparts. In a similar study that examined differences across race among a nationally representative cohort of young adults, Hudson et al. (2013) report that for Whites but not Blacks, the association between discrimination and self-rated health is more pronounced at higher levels of SES. There is also preliminary evidence to suggest that more frequent encounters with unfair treatment predict less healthy diurnal cortisol rhythms but only among low as opposed to high SES African Americans (Fuller-Rowell et al., 2012). Clearly, much more work is needed to

begin to understand, on a deeper level, how race and SES interact to produce health disparities and the extent to which these complex trends can be explained by differential exposure to discrimination.

To this end, the current study combines prospective cohort data from two NLSY surveys, the 1979 cohort (NLSY79) and the young adult cohort (NLSY79-YA) to answer the following research questions:

1. As SES increases over time, are individuals exposed to significantly less acute and chronic discrimination?
2. Does this association vary across race, such that upwardly mobile Whites are significantly less likely than upwardly mobile nonWhites to encounter instances of unfair treatment?
3. Can this differential exposure to discrimination account, at least in part, for racial disparities in health among individuals experiencing upward mobility?

2. Data and methods

2.1. Description of the data

The 1979 National Longitudinal Study of Youth is a prospective cohort study that began in 1979 by enrolling a nationally representative sample of young men and women between the ages of 14 and 21. This data collection effort was originally undertaken to investigate the labor market experiences and outcomes of young adults as they completed high school and began college, entered the work force, and established their own households. Racial minorities, low-income individuals, and members of the armed forces were oversampled. In addition to gathering data on educational attainment and employment, the original 1979 cohort also provided a wealth of information on childbearing and childrearing. To this end, all children born to female members of the 1979 cohort were followed up over time and form the basis of the NLSY79-Young Adult survey. Prospective data was collected on these children biennially from 1986 until 2012. As of 2012, the ages of the young adult cohort ranged from 14 to 41 years and African American and Hispanic respondents accounted for 28% and 19%, respectively, of the total sample.

For the purposes of this study, overall health was assessed using a measure of self-rated health. NLSY79 young adults were asked to describe their present health as excellent, very good, good, fair, or poor. Our decision to use this particular health indicator to capture overall health status was due to three important considerations. First, this type of simple, straightforward survey question has been repeatedly shown to be a powerful predictor of subsequent morbidity and mortality, even more so than physician examination (DeSalvo et al., 2006; Idler and Benyamini, 1997; Jylhä, 2009). Second, our study sample is relatively young and most have yet to develop chronic illnesses or serious health conditions, which are easier to capture with more specific indicators. Finally, self-rated health is a commonly used measure of overall wellbeing particular within the health disparities literature; thus, our findings will be easily and directly comparable to a wide swathe of the existing research.

Questions concerning discrimination were incorporated into the 2012 round of the NLSY79-YA survey and form the analytical basis of the current study. Experiences of acute discrimination were captured using the Major Experiences of Discrimination Scale (Williams et al., 1997; Kessler et al., 1999), while exposure to chronic discrimination was assessed by means of the Everyday Discrimination Scale (Williams et al., 1997). These multi-item measures were originally developed by Williams and colleagues (Williams et al., 1997) and possess several methodological advantages (Bastos et al., 2010; Krieger et al., 2005), including their strong

psychometric properties, comprehensive and respondent driven approach to capturing experiences with unfair treatment, their clarity and the ease with which it is understood by participants, and their comparability across studies.

Due to existing time constraints and considerable demands already placed on survey respondents, shortened versions of both discrimination scales were incorporated into existing data collection efforts. The Major Experiences of Discrimination Scale (MEDS) was limited to 5 questions that asked respondents if they ever encountered unfair treatment in the following settings: educational attainment, occupational advancement (hiring, promotion, and dismissal), and police harassment. Thus, this indicator is a lifetime rating of discrimination exposure. Only NLSY79-YA who were 21 years of age or older in 2012 were asked to complete the MEDS, since individual questions inquired about discrimination in settings that affected young people once they transitioned from adolescence to adulthood. The version of the Everyday Discrimination Scale (EDS) that was used in the NLSY79-YA survey contained 9 questions about chronic, as opposed to acute, instances of unfair treatment. For example, participants were asked, "In your day to day life, how often have any of the following happened to you? You have been treated with less courtesy than other people." Response categories ranged from (1) very often; (2) fairly often; (3) not too often; (4) hardly ever; to (5) never and were reverse coded, so that higher scores indicated more exposure to chronic discrimination. Similar to the wording used in the MEDS, this indicator does not require that the respondent recall a specific time period during which the unfair treatment occurred. For both discrimination measures, respondents had the opportunity to attribute reported unfair treatment to specific sociodemographic categories including race or ethnicity; however, small cell sizes across income trajectories prevented us from incorporating these data into the current study. A full description of both the MEDS and the EDS is presented in [Appendix A](#).

In the NLSY79-YA, race/ethnicity is based upon how the original NLSY79 female participant (ie. mother) was identified. We divided young adults into three distinct racial/ethnic groups: nonHispanic Whites, nonHispanic Blacks, and Hispanics. We excluded any respondent who did not fall within one of these three racial categories due to small cell sizes. Since the NLSY79-YA sample is comprised of children born to original NLSY79 female respondents, nativity status did not vary across Hispanic young adults.

For the current study, SES is captured using a time-varying indicator of household income that combines detailed information from both the NLSY79 and NLSY79-YA surveys for the years between 1979 and 2012. If, in a given year, the respondent was less than 21 years of age, did not work full-time, or did not live independently, we used parents' income to generate this measure. If all three inclusion criteria were met, we relied on the young adult's household income of the NLSY79 to calculate these SES trajectories. Latent growth curve analysis (LGCA) was employed to identify which NLSY79-YA respondents experienced downward mobility, upward mobility, or socioeconomic stability based on these income trajectories. We used household income to capture SES fluctuations over time because it (1) comprehensively captures the amount of economic resources available to the NLSY79 respondent, regardless of if he/she (or her mother) is currently in the labor force; (2) is consistently measured and can easily be adjusted for temporal changes by accounting for inflation whereas educational attainment or occupational prestige across generations cannot; (3) provides an indicator with enough variation both across respondents and over time to allow us to create stable SES trajectories; and (4) is correlated with parental educational attainment and occupational prestige – two indicators that do not as consistently vary over time and for which we have substantial amounts of missing data,

particularly among fathers and stay at home mothers.

LGCA identifies clusters of individuals who follow similar pathways along some variable over a specific unit of time (e.g. age or year) (Jones and Nagin, 2007, 2013). The method is a specialized application of finite mixture modeling using maximum likelihood to estimate the trajectory of each group, the proportion of the entire sample assigned to each group, and the probability of group membership for all observations (Jones and Nagin, 2007; Nagin, 2005). Unlike growth curve models, which assume that the entire population shares similar risks and follows similar trajectories over time, group-based models allow clusters of individuals to adhere to distinct time-varying trajectories (Jones and Nagin, 2007; Nagin, 2005; Nagin and Odgers, 2010). Although the groups are approximations and do not reflect actual entities, they provide a useful classification scheme for describing developmental pathways over time (Nagin, 2005; Nagin and Odgers, 2010; Nagin and Tremblay, 2005; Petts, 2009).

In the current study, we assign individuals into groups based on fluctuations in household income between 1979 and 2012. We used the following model to predict income as a polynomial function of age:

$$y_j = \beta_0 + \beta_1 \text{Age} + \beta_2 \text{Age}^2$$

where y is total family income for individuals in group j and β_1 and β_2 are population parameters that determine the shape of the trajectory. Because each trajectory can theoretically have its own unique shape, we estimate a unique set of parameters for each group (Nagin, 2005). After examining BIC/AIC values and posterior probabilities, completing a visual inspection of graphs, and in keeping with prior research on income trajectories in the United States (e.g. Cerdá et al., 2011), we finished with a 5-group LGCA model. The largest group (40%) experienced moderate upwardly mobility, while two other groups – the socioeconomically stable (30%) and the slow upwardly mobile (22%) – also accounted for a sizeable proportion of the sample. Only 3% of NLSY79 young adults could be classified as downwardly mobile. Finally, a small group (5%) had household incomes that were relatively high and grew steadily over time.

We control for a number of potential confounders in multivariate regression models, all of which were assessed in 2012. These include respondent's age, sex (male vs. female), marital status (not married, married, divorced, separated, widowed; cohabiting), region of residence (Northeast; Midwest; South; West), and employment status (neither in school or working; in school; working; in school and working).

2.2. Analytic strategy

A series of multilevel regression models were estimated to determine (1) whether the association between SES and acute or chronic discrimination significantly differed across the three racial/ethnic groups and (2) the extent to which differential exposure to acute or chronic discrimination could explain, at least in part, lower levels of self-rated health among upwardly mobile Blacks and Hispanics compared to upwardly mobile Whites. Ordered logit models were used to predict categorical outcomes including acute discrimination and self-rated health, while linear regression models were employed to generate findings for continuous outcomes such as chronic discrimination scores. Since we are estimating regression models for both continuous and categorical outcomes, we report slope coefficients (b) as opposed to odds ratios (OR) in [Tables 2](#) through [4](#) to maintain consistency.

NLSY79-YA data have a hierarchical structure wherein multiple young adults are nested within families. Multilevel models allow us

to deconstruct the variation in the dependent variable in two ways – across individuals and across families. Because it would be unrealistic to assume that error terms are uncorrelated if NLSY79 respondents were born to the same mother and grew up in the same family, our regression models estimate random intercepts—that is, the model provides an average intercept, as well as a variance parameter that summarizes individual-specific deviations from the average intercept. As such, our models parse between-individual and between-family variation in self-rated health, conditional on the variables included in the model. Given the non-independence of error terms and the presence of heteroskedasticity, we calculated robust standard errors using the Huber/White correction method and clustered them at the highest level of aggregation – the original NLSY79 family. All analyses were based on maximum likelihood estimation and conducted using Stata/MP 14.2.

We restrict our study sample to all NLSY79 young adults who were 18 years or older in 2012 and had nonmissing values for at least one discrimination measure as well as self-rated health. Because the MEDS was only administered to NLSY79 young adults who were 21 or older, while the EDS was administered to all NLSY79-YA participants, our sample size differs across these two discrimination measures, ranging from 4412 respondents (acute discrimination) to 5248 (chronic discrimination). Proportions of missing data for both the MEDS and EDS were exceedingly low, at 1.6% and 0.13%, respectively.

To handle issues of missing data on additional variables, we relied on multiple imputation techniques to generate values for all covariates included in descriptive and multivariate analyses. Evidence from the relationships of missing data with individual time varying and time invariant characteristics suggests the data are not missing completely at random (unconditional on the observed covariates) which makes typical approaches such as listwise deletion inappropriate (Allison, 2001). In the imputation phase, the procedures use a diverse set of predictors to estimate five sets of plausible values for each missing value. The imputed values in the five data sets include a random component based on draws from the posterior predictive distribution of the missing data under a posited Bayesian model and, under the missing-at-random assumption – a more plausible assumption than is made by listwise deletion – provide unbiased estimates of variance (Allison, 2001). Following Von Hippel (2007), we impute values for all variables in a given model and then delete observations with missing data on self-rated health.

3. Results

Descriptive findings are presented in Table 1. Among this sample of NLSY79 young adults, 45% were nonHispanic White, 34% were nonHispanic Black, and 21% were Hispanic. This overrepresentation of nonwhite respondents allowed us to generate stable comparisons across race and SES. Regarding the classification of income trajectories we generated using LGCA, the majority of NLSY79 young adults fall into one of three categories – socioeconomically stable (30%), slow upward mobility (22%), and moderate upward mobility (40%). Only a small proportion were found to adhere to either downwardly mobile (3%) or rapid upwardly mobile (5%) trajectories. Furthermore, we found substantial racial disparities in income trajectories, with larger proportions of nonHispanic Blacks and Hispanics experiencing downward mobility or slow upward mobility and smaller proportions experiencing moderate or rapid upward mobility compared to nonHispanic Whites.

A sizable fraction of the NLSY79-YA sample report exposure to both acute and chronic discrimination. Overall, 38% of respondents experienced at least one instance of acute discrimination, with almost 15% reporting frequent occurrences (3 or more) of acute discrimination. The mean chronic discrimination score for the entire sample was 17.50 with a standard deviation of 6.2. Table 1 reveals sizeable racial differences in exposure to unfair treatment among this young adult population. Most notably, nonHispanic Blacks experienced acute discrimination more frequently than nonHispanic Whites, with 29% and 21% of Blacks reporting 1–2 and 3 + instances of acute discrimination, respectively, compared to only 20% and 10% of Whites. Similarly, mean chronic discrimination scores were significantly higher among nonHispanic Blacks (18.61) than nonHispanic Whites (16.88). For Hispanics, racial disparities in exposure to unfair treatment follow a slightly different pattern. Hispanics are more likely than nonHispanic Whites to experience moderate (24% vs. 20%) and frequent (14% vs. 10%) instances of acute discrimination but just as likely as their White counterparts to experience chronic discrimination (17.01 vs. 16.88). Table 1 also reveals significant disparities across race with respect to age and sex of the respondent, marital status, region of the country, and employment status, with more pronounced differences noted between Blacks and Whites than Hispanics and Whites.

We present the distribution of unfair treatment by race and mobility status in Figs. 1 and 2 and find greater racial disparities in the patterning of exposure for acute as opposed to chronic discrimination. As Fig. 1 makes evident, nonHispanic Whites and Hispanics who are upwardly mobile (moderate and rapid) tend to experience less acute discrimination than those who are socioeconomically stable or downwardly mobile. For example, among Whites, mean acute discrimination scores for moderate and rapid upwardly mobile respondents are 0.38 and 0.32, respectively, compared to 0.48 for socioeconomically stable respondents. Blacks follow a different pattern with respect to exposure to acute unfair treatment. Upward socioeconomic mobility does not appear to shield nonHispanic Black NLSY79 young adults from encountering instances of acute discrimination. In fact, Black moderately upwardly mobile respondents report *more* experiences (0.82) of acute discrimination than their Black socioeconomically stable counterparts (0.78).

Fig. 2 reveals similar trends regarding the distribution of chronic discrimination across race and income trajectory, albeit less pronounced than those presented in Fig. 1 for acute discrimination. With the exception of downwardly mobile young adults, exposure to chronic discrimination tends to decrease as nonHispanic White and Hispanic respondents become more upwardly mobile. For nonHispanic Whites, mean chronic discrimination scores decreased from 17.6 in the socioeconomically stable category to 16.5 and 16.3 in the moderate and rapid upwardly mobile categories. However, the trend of falling levels of chronic unfair treatment with rising levels of mobility is not evident among Black NLSY79-YA respondents. On average, socioeconomically stable Black young adults report chronic discrimination scores of 18.4, while their counterparts in the slow, moderate, and rapid upwardly mobile groups report chronic discrimination scores of 19.1, 18.1, and 18.5.

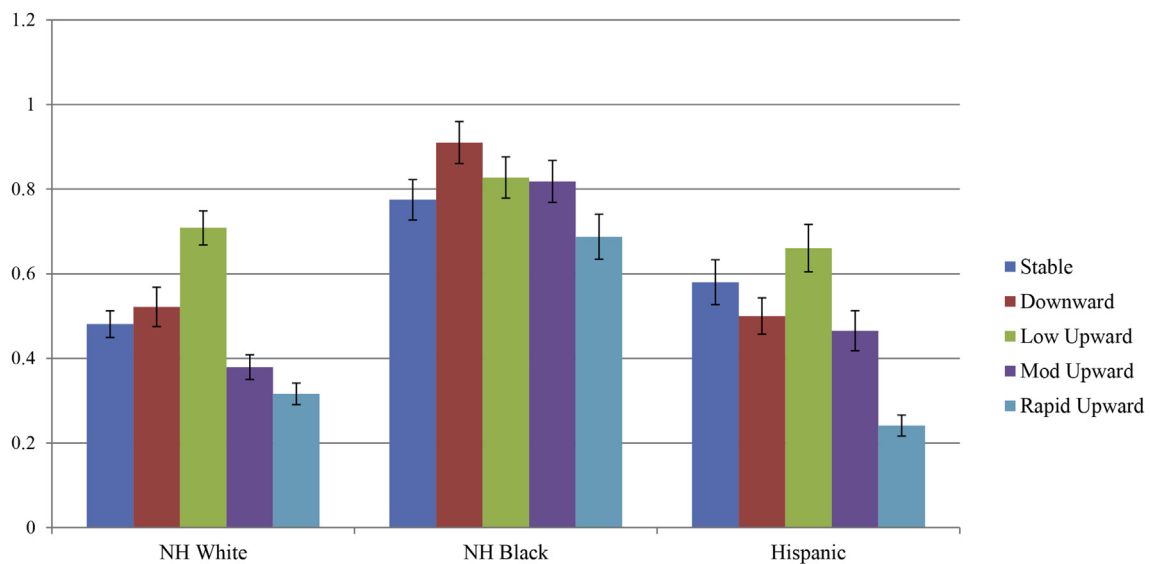
In Table 2, we present findings from multilevel regression analyses designed to test whether the racial differences we notice in the association between income trajectories and unfair treatment vary significantly across race. All multivariate models controlled for age, sex, marital status, region of residence, and employment status. Regression results reveal that compared to nonHispanic

Table 1
Descriptive statistics for NLSY79 young adults by race of respondent.

| | | NH Whites | NH Blacks | Hispanics |
|-----------------------------|--------|-----------|-----------|-----------|
| <i>Race</i> | | | | |
| NH White | 45.24 | | | |
| NH Black | 33.54 | | | |
| Hispanic | 21.22 | | | |
| <i>Income Trajectory</i> | | | | |
| Stable | 29.70 | 25.73 | 33.16 | 32.68 |
| Downward | 2.99 | 1.14 | 5.79*** | 2.51+ |
| Slow Upward | 21.98 | 8.25 | 36.68*** | 28.01*** |
| Moderate Upward | 40.32 | 56.42 | 23.06*** | 33.3*** |
| Rapid Upward | 5.01 | 8.46 | 1.31*** | 3.5*** |
| <i>Acute Discrimination</i> | | | | |
| None | 61.52 | 69.97 | 50.39 | 62.39 |
| 1–2 Times | 23.56 | 19.58 | 28.58*** | 23.52** |
| 3 + Times | 14.92 | 10.45 | 21.02*** | 14.09** |
| Chronic Discrimination | 17.49 | 16.88 | 18.61*** | 17.01 |
| | (6.22) | (5.72) | (6.82) | (5.98) |
| <i>Age</i> | | | | |
| | 25.93 | 25.14 | 26.82*** | 26.21*** |
| | (4.74) | (4.44) | (5.01) | (4.65) |
| <i>Sex</i> | | | | |
| Male | 48.48 | 48.93 | 46.34 | 50.90 |
| Female | 51.52 | 51.07 | 53.66+ | 49.10 |
| <i>Marital Status</i> | | | | |
| Not Married | 60.76 | 57.31 | 67.86 | 56.91 |
| Married | 18.69 | 23.49 | 11.93*** | 19.12* |
| Divorced/Sep/Widow | 4.40 | 3.37 | 5.79* | 4.40 |
| Cohabiting | 16.15 | 15.83 | 14.42** | 19.57* |
| <i>Region</i> | | | | |
| Northeast | 13.85 | 15.71 | 11.75 | 13.20 |
| Midwest | 24.38 | 34.48 | 20.14* | 9.61*** |
| South | 41.70 | 32.17 | 61.22*** | 31.15 |
| West | 20.08 | 17.64 | 6.93*** | 46.05*** |
| <i>Employment Status</i> | | | | |
| Neither | 7.98 | 5.31 | 11.02 | 8.89 |
| In School | 4.51 | 3.58 | 5.39 | 5.12 |
| Working | 62.48 | 62.27 | 61.16*** | 64.99*** |
| In School & Working | 25.03 | 28.84 | 22.43*** | 21.01*** |
| N | 5,250 | 2,375 | 1,761 | 1,114 |

***p < 0.001; **p < 0.01; *p < 0.05; + p < 0.10.

Source: NLSY79-Young Adult Survey.



Note: Acute discrimination scores can range from a minimum of 0 to a maximum of 5.

Fig. 1. Mean acute discrimination scores by race & income trajectory.

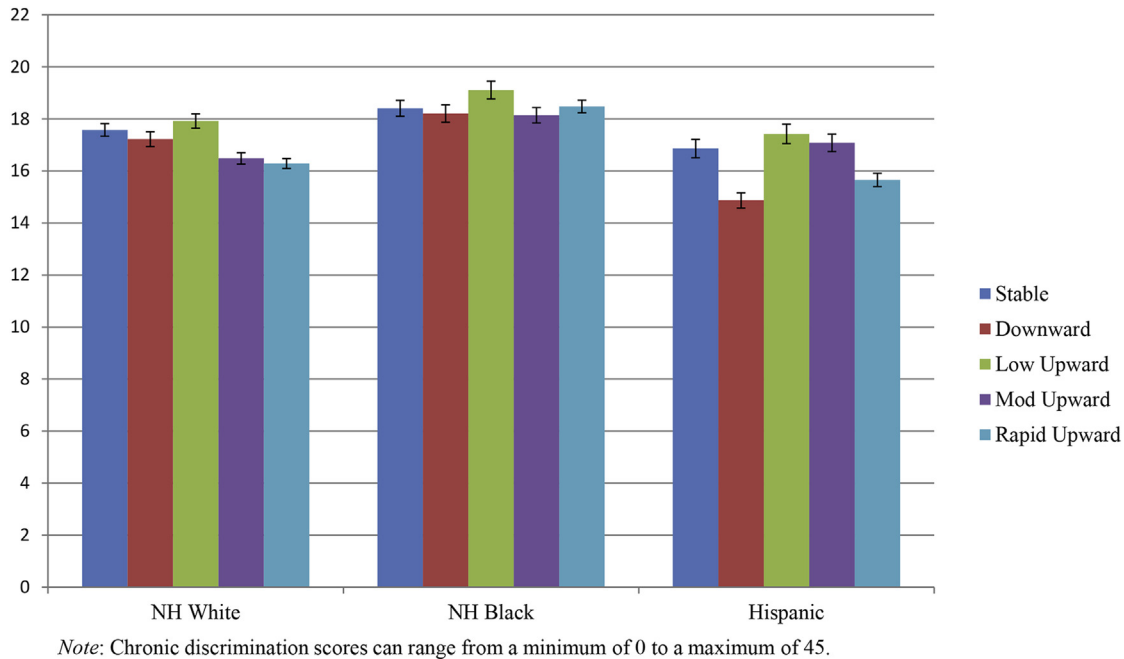


Fig. 2. Mean chronic discrimination scores by race & income trajectory.

Table 2
Results from multilevel regression models predicting acute and chronic discrimination scores by race and mobility trajectory among NLSY79 young adults.

| | Acute Discrimination | | Chronic Discrimination | |
|--------------------------------|----------------------|-----------------|------------------------|-----------------|
| | <i>b</i> | SE ^a | <i>b</i> | SE ^a |
| <i>Race</i> | | | | |
| NH White | Ref | | Ref | |
| NH Black | 0.6700*** | 0.142 | 0.8444* | 0.370 |
| Hispanic | 0.1306 | 0.169 | -0.8594* | 0.399 |
| <i>Income Trajectory</i> | | | | |
| Stable | Ref | | Ref | |
| Downward | -0.1684 | 0.586 | -0.3052 | 1.435 |
| Slow Upward | 0.5744** | 0.207 | 0.2923 | 0.570 |
| Moderate Upward | -0.2669* | 0.133 | -1.0700*** | 0.291 |
| Rapid Upward | -0.4648+ | 0.253 | -1.1761** | 0.446 |
| <i>Race*Income Interaction</i> | | | | |
| Black*Downward | 0.5045 | 0.634 | 0.0652 | 1.616 |
| Black*Slow Upward | -0.4561+ | 0.243 | 0.3410 | 0.688 |
| Black*Mod Upward | 0.4244* | 0.199 | 0.9179+ | 0.507 |
| Black*Rapid Upward | -0.0490 | 0.598 | 1.1842 | 0.835 |
| Hispanic*Downward | 0.3384 | 0.755 | -2.3653 | 1.839 |
| Hispanic*Slow Upward | -0.3423 | 0.281 | 0.1950 | 0.743 |
| Hispanic*Mod Upward | -0.0386 | 0.231 | 1.3033* | 0.515 |
| Hispanic*Rapid Upward | -0.3664 | 0.520 | -0.3508 | 0.892 |
| N | 4,371 | | 5,250 | |

Note: All regression models controlled for age, sex, marital status, region of residence, and employment status.

***p < 0.001; **p < 0.01; *p < 0.05; + p < 0.10.

^a Robust standard errors were calculated using the Huber/White correction method and clustered at the level of the original NLSY79 mother.

Source: NLSY79-Young Adult Survey.

Whites, nonHispanic Blacks face 95% higher odds [$1 - e^{(0.6700)} * 100$] of experiencing acute discrimination. Regarding the association between socioeconomic mobility and exposure to acute discrimination, respondents in the moderate and rapid upwardly mobile categories face 24% [$1 - e^{(-0.2669)} * 100$] and 37% [$1 - e^{(-0.4648)} * 100$]

lower odds, respectively, of experiencing acute discrimination compared to those in the socioeconomically stable category although the regression coefficient for the rapid upwardly mobile fails to reach statistical significance ($p < 0.05$). Results for chronic discrimination echo those found for acute discrimination with two notable exceptions. First, although Black NLSY79-YA respondents report significantly higher chronic discrimination scores ($b = 0.8444$) than Whites, Hispanic NLSY79-YA respondents report significantly lower chronic discrimination scores ($b = -0.8594$) than members of the referent group. Second, young adults in both the moderate and rapid upward mobility categories revealed significantly lower chronic discrimination scores ($b = -1.07$ and -1.18 , respectively) compared to young adults in the socioeconomically stable group.

Of particular interest for this study, we find evidence of significant interactions between race and income on exposure to unfair treatment among respondents who experienced both slow and moderate upward mobility, with more pronounced racial differences noted for African Americans than Hispanics. For nonHispanic Whites, those whose income trajectories moderately rose over time faced 24% [$1 - e^{(-0.2669)} * 100$] lower odds of encountering acute discrimination. However, for nonHispanic Blacks, those whose income trajectories moderately rose over time faced 17% [$1 - e^{(-0.2669+0.4244)} * 100$] higher odds of encountering acute discrimination. Similarly, within the slow upward mobility category, nonHispanic Whites had a 77% [$1 - e^{(0.5744)} * 100$] increase in the odds of acute discrimination while nonHispanic Blacks had a 13% [$1 - e^{(0.5744+-0.4561)} * 100$] increase in the odds of acute discrimination.

With respect to the association between socioeconomic mobility and chronic discrimination exposure, we only found empirical evidence of racial disparities for those experiencing moderate income gains over time. These tended to be more pronounced among Hispanics than nonHispanic Blacks. White NLSY79 young adults in the moderate upwardly mobile group reported mean chronic discrimination scores that were 1.07 lower than White NLSY young adults in the socioeconomically stable group.

Table 3
Results from multilevel ordered logit regression models predicting self rated health by race and acute discrimination among upwardly mobile NLSY79 young adults.

| | Model 1 | | Model 2 | |
|-----------------------------|------------|-----------------|------------|-----------------|
| | b | SE ^a | b | SE ^a |
| <i>Race</i> | | | | |
| NH White | Ref | | Ref | |
| NH Black | -0.1850* | 0.093 | -0.0782 | 0.095 |
| Hispanic | -0.3392** | 0.106 | -0.3193** | 0.105 |
| <i>Acute Discrimination</i> | | | | |
| None | | | Ref | |
| 1–2 Times | | | -0.4322*** | 0.087 |
| 3 + Times | | | -0.5913*** | 0.125 |
| Age | -0.0426*** | 0.011 | -0.0405*** | 0.011 |
| Female | -0.2785*** | 0.071 | -0.3425*** | 0.071 |
| <i>Marital Status</i> | | | | |
| Not Married | Ref | | Ref | |
| Married | 0.3018** | 0.100 | 0.2711** | 0.100 |
| Divorced/Sep/Widow | -0.0871 | 0.197 | -0.0343 | 0.196 |
| Cohabiting | -0.2113* | 0.098 | -0.1870+ | 0.098 |
| <i>Region</i> | | | | |
| Northeast | Ref | | Ref | |
| Midwest | -0.0780 | 0.126 | -0.1058 | 0.125 |
| South | -0.0076 | 0.121 | -0.0301 | 0.121 |
| West | -0.1399 | 0.135 | -0.1359 | 0.134 |
| <i>Employment Status</i> | | | | |
| Neither | Ref | | Ref | |
| In School | 0.5626+ | 0.310 | 0.5578+ | 0.304 |
| Working | 0.6925*** | 0.150 | 0.7025*** | 0.151 |
| In School & Working | 0.9704*** | 0.167 | 0.9532*** | 0.167 |
| N | 2,890 | | 2,890 | |

Note: All regression models controlled for age, sex, marital status, region of residence, and employment status.

***p < 0.001; **p < 0.01; *p < 0.05; + p < 0.10.

^a Robust standard errors were calculated using the Huber/White correction method and clustered at the level of the original NLSY79 mother.

Source: NLSY79-Young Adult Survey.

However, nonHispanic Black respondents whose incomes moderately increased experienced only 0.15 [-1.0700 + 0.9179] fewer instances of chronic discrimination than their nonHispanic White counterparts, while Hispanic respondents in this mobility group experienced 0.23 [-1.0700 + 1.3033] more instances of chronic discrimination.

Tables 3 and 4 reveal results from statistical analyses regressing self-rated health on acute discrimination and chronic discrimination, respectively, among upwardly mobile NLSY79 young adults. We highlight how the magnitude of the regression coefficients for race (NH Black and Hispanic) change from Model 1 to Model 2 once we control for discrimination in order to determine the extent to which differential exposure to unfair treatment is likely to restrict the health returns to upward mobility for racial minorities. Model 1 of Table 3 illustrates that upwardly mobile nonHispanic Blacks and Hispanics face 17% [$1 - e^{(-0.1850)} * 100$] and 29% [$1 - e^{(-0.3392)} * 100$], respectively, lower odds of being in good health than nonHispanic Whites, net of the effects of age, sex, marital status, region of residence, and employment status. Once we adjust for differential exposure to acute discrimination, the magnitude of the regression coefficient for Blacks decreases by 58% from -0.185 to -0.078 and fails to retain statistical significance ($p < 0.05$), while the coefficient for Hispanics only decreases by 6% from -0.3392 to -0.3193 and remains statistically significant ($p < 0.05$). These findings suggest that higher than expected exposure to acute discrimination helps, at least in part, to explain why African Americans do not tend to experience the same beneficial health returns to upward mobility that Whites typically do.

Similar results concerning chronic discrimination are presented

in Table 4. Among upwardly mobile young adults, nonHispanic Blacks and Hispanics experience 23% [$1 - e^{(-0.2543)} * 100$] and 28% [$1 - e^{(-0.3270)} * 100$] lower odds of being in better health than non-Hispanic Whites. These estimates are slightly different than those found in Table 3 due to our reliance on a larger sample size that resulted from fewer missing values on the measure of chronic as opposed to acute discrimination. Once we account for more frequent exposure to chronic discrimination among racial minorities, the regression coefficient for nonHispanic blacks decrease by 47% from -0.2543 to -0.1354 and no longer retained statistical significance ($p < 0.05$), while the regression coefficient for Hispanics only decreased by 3% from -0.3270 to -0.3180. In both Tables 3 and 4, we also note that older ages, being female, and cohabiting were significantly associated with worse self-rated health, while being married as well as being employed, in school, or being employed and in school were significantly associated with better self-rated health.

4. Discussion

As the analyses presented here reveal, we find evidence of striking racial/ethnic differences in exposure to unfair treatment. In contrast to other studies which reveal Whites reporting more frequent encounters of discrimination than nonWhites (Williams and Mohammed, 2009), the Black and Hispanic young adults in our study demonstrate higher levels of acute and chronic discrimination. Although we cannot say with certainty what is driving these findings, it could be a result of our reliance on a younger cohort of survey participants who might be more comfortable reporting instances of unfair treatment than older cohorts, especially in an era in which subtle forms of discrimination are more common than explicit instances of unfair treatment

Table 4
Results from multilevel ordered logit regression models predicting self rated health by race and chronic discrimination among upwardly mobile NLSY79 young adults.

| | Model 1 | | Model 2 | |
|-------------------------------|------------|-----------------|------------|-----------------|
| | b | SE ^a | b | SE ^a |
| <i>Race</i> | | | | |
| NH White | Ref | | Ref | |
| NH Black | -0.2543** | 0.084 | -0.1354 | 0.085 |
| Hispanic | -0.3270** | 0.100 | -0.3180** | 0.100 |
| <i>Chronic Discrimination</i> | | | | |
| Age | -0.0406*** | 0.009 | -0.0390*** | 0.009 |
| Female | -0.2131*** | 0.065 | -0.2848*** | 0.065 |
| <i>Marital Status</i> | | | | |
| Not Married | Ref | | Ref | |
| Married | 0.2778** | 0.099 | 0.1979* | 0.099 |
| Divorced/Sep/Widow | -0.1046 | 0.197 | -0.0893 | 0.197 |
| Cohabiting | -0.2421* | 0.096 | -0.2301* | 0.096 |
| <i>Region</i> | | | | |
| Northeast | Ref | | Ref | |
| Midwest | -0.0327 | 0.112 | -0.0512 | 0.111 |
| South | 0.0638 | 0.107 | 0.0248 | 0.106 |
| West | -0.1119 | 0.123 | -0.1215 | 0.121 |
| <i>Employment Status</i> | | | | |
| Neither | Ref | | Ref | |
| In School | 0.5649** | 0.210 | 0.5067* | 0.211 |
| Working | 0.5791*** | 0.142 | 0.5502*** | 0.145 |
| In School & Working | 0.8608*** | 0.153 | 0.8034*** | 0.156 |
| N | 3,534 | | 3,534 | |

Note: All regression models controlled for age, sex, marital status, region of residence, and employment status.

***p < 0.001; **p < 0.01; *p < 0.05; + p < 0.10.

^a Robust standard errors were calculated using the Huber/White correction method and clustered at the level of the original NLSY79 mother.

Source: NLSY79-Young Adult Survey.

(Bonilla-Silva, 2006).

We also note key differences in the association between SES and encounters with unfair treatment across race. For Whites, moderate income gains over time resulted in significantly less exposure to both acute and chronic discrimination. African Americans, on the other hand, could not count on reaping these potentially health protective benefits from upward mobility. In fact, African Americans whose incomes moderately increased over time were significantly *more* likely to experience acute discrimination and just as likely to experience chronic discrimination than their socioeconomically stable counterparts. Although we did not find evidence of racial disparities in the association between socioeconomic mobility and acute discrimination for Hispanics, a similar pattern emerges when examining outcomes for chronic discrimination. For Hispanics, income gains over time actually *increase* the frequency of encounters of chronic discrimination. Thus, it appears that regarding exposure to unfair treatment, upward mobility serves a protective function for Whites but places African Americans and Hispanics at a greater risk.

Although this result is somewhat surprising, it is not unsubstantiated in the literature given that other investigators have found exposure to discrimination to be more frequent or the relationship between unfair treatment and health to be more pronounced among high as opposed to low SES groups (Fuller-Rowell et al., 2012; Hudson et al., 2012, 2013). Hudson et al. (2012) demonstrate that the positive association between both acute and chronic discrimination and depression is more pronounced among African Americans with the highest levels of educational attainment (16 + years) and household incomes (\geq \$55,000). Similarly, Fuller-Rowell et al. (2012) find that for African Americans reports of perceived discrimination are associated with flatter (less healthy) diurnal cortisol trajectories, but this relationship appears to be significantly less pronounced among higher rather than lower SES participants. These findings lend credence to the argument that the “payoffs” typically associated with gains in SES are likely to be muted for members of racial minorities due to their continued exposure to race-based unfair treatment despite having access to more than adequate levels of socioeconomic resources.

Although the current study provides additional evidence that higher SES African Americans are more likely to report instances of discrimination, it does not shed light on why this social trend might be occurring. Researchers have posited it might be due to increased exposure to Whites in more integrated educational, occupational, or neighborhood settings (Gee, 2002; Fuller-Rowell et al., 2012), while others underscore the fact that nonpoor African Americans might be more likely to notice unfair treatment than their poor counterparts (Hudson et al., 2012), thus differences in discrimination within race but across SES is being driven more by disparate reporting than actual unequal exposure. This finding might also largely result from the frustration and discouragement higher status African Americans experience by not being able to socially and economically benefit from upward mobility to the same extent as similarly situated Whites (Colen, 2011; Lewis et al., 2015). Given that we are unable to further explore the extent to which these potential explanations account for differences in the association between discrimination and health across SES using the NLSY79-YA, we look toward future research efforts to conduct more pointed analyses and provide more indepth insights into this specific issue.

Our findings also lend credence to the idea that racial disparities in health result, at least in part, from exposure to discrimination among a group of upwardly mobile young adults – a group that is

typically thought of as having access to higher than average social and economic resources. We show that, even after controlling for a whole host of potential confounders (age, sex, marital status, region of residence, and employment status), the health of African Americans still lags well behind that of Whites. Once we account for the increased likelihood of encountering unfair treatment among African Americans, the Black/White gap in self-rated health is reduced to such an extent as to render it statistically insignificant.

Although the number of studies designed to explore the effects of discrimination on health have been steadily increasing in recent years, there are few research efforts that directly test whether exposure to discrimination is a driving force behind the remarkably persistent racial disparities in health that characterize American society today (Lewis et al., 2015; Williams and Mohammed, 2009). Our study provides preliminary evidence to support this notion and echoes previous results linking exposure to unfair treatment to Black/White inequalities in poor birth outcomes (Lauderdale, 2006; Mustillo et al., 2004), nocturnal blood pressure dipping (Tomfohr et al., 2010), and sleep patterns (Tomfohr et al., 2012).

We do not, however, find that exposure to either acute or chronic discrimination explained much, if any, of why Hispanic young adults in our sample report worse self-rated health than their White counterparts. This result is even more striking given that all of the Hispanic NLSY79 young adults were born in the U.S., with a sizeable majority to parents who, themselves, were not immigrants. Sternthal et al. (2011) note that native-born Hispanics report frequency of discrimination similar to nonHispanic Blacks, while their foreign-born counterparts report instances of unfair treatment similar to nonHispanic Whites. This is not to suggest, however, that unfair treatment does not matter for members of this racial/ethnic group or that it does not negatively impact their health. On the contrary, there is a rapidly growing body of evidence to support the idea that discrimination erodes the mental and physical health of individuals from a wide range of racial/ethnic backgrounds, including those who identify as Hispanic or Latino (Everson-Rose et al., 2015; Gee and Walsemann, 2009; Gee et al., 2007; Ryan et al., 2006).

The paucity of a mediating effect of discrimination on health for Hispanics that we report could stem from a lack of cultural resonance among the questions that comprised the acute and chronic discrimination measures or could be highly dependent upon level of acculturation. There is some evidence to suggest that Hispanics understand and answer questions about self-rated health differently than other racial/ethnic groups and changes the longer they live in the U.S. (Acevedo-Garcia et al., 2010; Finch et al., 2002; Viruell-Fuentes et al., 2011). Primarily due to data constraints, we were unable to further investigate this key difference in the association between discrimination and self-rated health across race in the current paper, but it remains an interesting avenue for future research.

4.1. Limitations

The findings from the current study should be viewed in the context of some important methodological shortcomings. First, the classification system we used as a result of LGCA to classify respondents as downwardly mobile, upwardly mobile, and socioeconomically stable does not reflect actual divisions but instead identifies groups of individuals who adhere to similar income trajectories. Thus, it is unclear how generalizable our findings will be to other datasets in which respondents experience very different fluctuations in income over time. We chose this more data driven

analytic approach to identifying mobility trajectories because of the many problems inherent in quantifying comparing changes in SES over time and across disparate subgroups.

Second, we relied on family income, as opposed to educational attainment or occupational prestige, to determine the socioeconomic status of NLSY79 young adults. This was done purposefully since the majority of our respondents had not moved beyond the lifecourse stage known as emerging adulthood -- a period of time characterized by entrance into or exit from college, frequent changes in employment, and transitional living situations - thus making comparisons based on educational attainment and/or occupational prestige difficult and imprecise. By 2012, more than half of the study sample was less than 25 years old, the age at which educational attainment is typically assessed. Moreover, we rely on data from 1979 to 2012 to estimate mobility trajectories. During this period of time the NLSY79 young adults were children and adolescents who were nowhere near completing their schooling. Given these data constraints, we would not be able to determine if educational attainment reflected achieved social status or was simply a reflection of the respondent's young age. This is not to suggest, however, that the effects of discrimination on health do not operate through or otherwise depend on educational attainment. Other investigators have found evidence to suggest that African Americans with higher levels of educational attainment report more instances of racial discrimination (Hunt et al., 2007; Borrell et al., 2006; Krieger et al., 2011).

Household income as an indicator of SES offered several advantages to a study, such as ours, that was designed to examine the association between socioeconomic mobility, discrimination, and health over time among a sample of young adults, many of whom have not yet completed their education or gained a solid foothold on their career paths. It is more likely to comprehensively capture the financial resources available to a respondent if he/she is not currently working full-time by focusing on an expansive rather than restrictive measure of SES that is captured at the household rather than the individual level. Moreover, there is substantial variation in family income, both over time and across respondents, that allows us to more accurately discern relatively small fluctuations in SES. We do not know the extent to which our findings are consistent across disparate indicators of socioeconomic status.

Third, our measures of acute and chronic discrimination reflect exposure to unattributed unfair treatment rather than unfair treatment that has been attributed to the respondent's race and/or ethnicity. Although we would have preferred to rely on more specific indicators that directly tied a respondent's report of discrimination to his/her racial group identity, small cell sizes across income trajectories prevented us from doing so. It should be noted, however, that using more general measures of unfair treatment allows us to capture instances of discrimination that stem from multiple group identities or for which the respondent is not sure how to accurately assess attribution. This could be particularly useful in more accurately gauging frequency of unfair treatment among women of color for whom it might be difficult to compartmentalize discriminatory experiences as primarily stemming from one's race or gender. Moreover, the existing literature regarding the health effects of unfair treatment includes examples of both types of discrimination (unattributed and attributed) indicators (Lewis et al., 2015); thus, our approach is not without precedent.

Finally, due to the cross-sectional nature of our discrimination measure and our outcome variable, self-rated health, we were unable to fully examine changes in key indicators of interest over time. Thus, the conclusions that can be drawn regarding the effect

of discrimination on health and the extent to which this differs across race and SES are limited in scope and require further replication using methodological approaches that can more stringently account for unobservables. By integrating income data from both the NLSY79 young adults and their parents, however, we were able to construct longitudinal income trajectories using data that spanned up to 33 years to predict both discrimination and self-rated health in 2012. Thus, our results are not completely based on cross-sectional data. Since the NLSY79-YA survey is a prospective cohort study for which interviews are conducted biennially, it is possible that in the future we will be able to examine how exposure to discrimination, SES, and health all fluctuate over time, thus gaining a clearer understanding of the complex ways in which these social conditions shape the life chances of Americans as they move through successive lifecourse stages.

4.2. Conclusion

The association between SES and health is typically thought to be remarkably consistent, whether it is across time, place, or health outcome. However, more recent research has begun to call this seemingly evident "truth" into question. As more evidence comes to light, it appears that the shape of the SES/health gradient might differ across race. The current study adds to this debate by providing additional evidence that for young African American and Hispanic Americans, gains in SES might result in *more* as opposed to less exposure to discrimination and these encounters with unfair treatment explain a sizeable proportion of the Black/White gap in self-rated health among upwardly mobile, young adults in the U.S.

Our findings can and should be taken to suggest that efforts to alleviate poverty, while incredibly important in their own right, will not be enough to eradicate or at least substantially lessen racial disparities in health. When it comes to designing public health interventions or public policies to reduce these inequalities, nonpoor subpopulation are often overlooked. This is not only unfortunate but results in a lost opportunity to affect change. Although absolute levels of morbidity and mortality tend to be higher among lower SES groups, the racial gaps in outcomes tend to be substantially more pronounced among higher SES groups. The results of this study lend credence to the idea that in order to reduce racial disparities in health, we need to seriously address the unfair treatment racial minorities experience – not only acute instances of discrimination but those slights that occur on a daily basis, and do so among the poorest segments of our society as well as among those who are somewhat better off.

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Appendix A. Description of major experiences and everyday discrimination scale used in NLSY-YA survey, 2012

| Description of Discrimination Measures | Response Categories |
|---|---------------------|
| <i>Major Experiences of Discrimination Scale (MEDS)</i> | |
| In the following questions, we are interested in the way other people have treated you or your beliefs about how other people have treated you. Can you tell me if any of the following has ever happened to you: | |
| 1 Have you ever been unfairly discouraged by a teacher or advisor from continuing your education? | Yes/No |
| 2 Have you ever not been hired for a job? | |
| 3 Have you ever been unfairly denied a promotion? | |
| 4 Have you ever been unfairly fired? | |
| 5 Have you ever been unfairly stopped, searched, questioned, physically threatened, or abused by the police? | |
| <i>Everyday Discrimination Scale (EDS)</i> | |
| In your day-to-day life, how often have any of the following happened to you? | |
| 1 How often have you been treated with less courtesy than other people? | Very often |
| 2 How often have you been treated with less respect than other people? | Fairly often |
| 3 How often have you received poorer service than other people at restaurants or stores? | Not too often |
| | Hardly Ever |
| | Never |
| 4 How often have people acted as if they think you are not smart? | |
| 5 How often have people acted as if they are afraid of you? | |
| 6 How often have people acted as if they think you are dishonest? | |
| 7 How often have people acted as if they are better than you? | |
| 8 How often have you been called names or insulted? | |
| 9 How often have you been threatened or harassed? | |

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