

Quantitative Methods in Economics

Functional form

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Roadmap, Part I

1. Linear predictors and least squares regression
2. Conditional expectations
3. **Some functional forms for linear regression**
4. Regression with controls and residual regression
5. Panel data and generalized least squares

Takeaways for these slides

Functional forms:

- ▶ Quadratic: decreasing or increasing returns
- ▶ Interactions: returns vary with covariates
- ▶ Discrete regressors, dummy variables, and saturated regressions
- ▶ Polynomial
- ▶ Linear in logarithms: elasticities
- ▶ Justification via Mincer model

► **Quadratic polynomial:**

- Y = earnings, Z = experience; $X_1 = Z$, $X_2 = Z^2$

$$E^*(Y | 1, X_1, X_2) = \beta_0 + \beta_1 X_1 + \beta_2 X_2$$

- Evaluating this at $Z = c$ gives $\beta_0 + \beta_1 c + \beta_2 c^2$.

► **Interactions:**

- Z_1 = experience, Z_2 = education; $X_1 = Z_1$, $X_2 = Z_2$, $X_3 = Z_1 \cdot Z_2$

$$E^*(Y | 1, X_1, X_2, X_3) = \beta_0 + \beta_1 X_1 + \beta_2 X_2 + \beta_3 X_3$$

- Evaluating this at $Z_1 = c$, $Z_2 = d$ gives $\beta_0 + \beta_1 c + \beta_2 d + \beta_3 c \cdot d$.

Questions for you

- ▶ Interpret these two functional forms.
- ▶ What happens as education is increased?
- ▶ How does that depend on the education we start with?
- ▶ How does that depend on experience?

- Recall conditional expectation: Solution to

$$\min_g E[Y - g(Z)]^2$$

is the *regression function*:

$$r(z) = E(Y | Z = z).$$

- Orthogonality Conditions: Consider any function $h(\cdot)$. Define

$$U = Y - r(Z).$$

- Then $U \perp h(Z)$, i.e. $E[Uh(Z)] = 0$, and in particular

$$E^*(Y|r(Z), h(Z)) = \beta_1 r(Z) + \beta_2 h(Z) = r(Z).$$

- Put differently: If $E(Y|X = x)$ is linear in x , then

$$E(Y|X = x) = E^*(Y|X = x).$$

Discrete regressors

- ▶ Assume

$$Z_1 \in \{\lambda_1, \dots, \lambda_J\}, \quad Z_2 \in \{\delta_1, \dots, \delta_K\}.$$

- ▶ Dummy Variables:

$$\begin{aligned} X_{jk} &= \begin{cases} 1, & \text{if } Z_1 = \lambda_j, Z_2 = \delta_k \\ 0, & \text{otherwise} \end{cases} \\ &= 1(Z_1 = \lambda_j, Z_2 = \delta_k). \end{aligned}$$

- ▶ Claim: $E(Y | Z_1, Z_2) = E^*(Y | X_{11}, \dots, X_{J1}, \dots, X_{1K}, \dots, X_{JK})$

Questions for you

Prove this.

Solution:

- ▶ Any function $g(Z_1, Z_2)$ can be written as

$$g(Z_1, Z_2) = \sum_{j=1}^J \sum_{k=1}^K \gamma_{jk} X_{jk}$$

with $\gamma_{jk} = g(\lambda_j, \delta_k)$.

- ▶ Thus

$$E(Y | Z_1, Z_2) = \sum_{j=1}^J \sum_{k=1}^K \beta_{jk} X_{jk},$$

where

$$\beta_{jk} = E(Y | Z_1 = \lambda_j, Z_2 = \delta_k).$$

- ▶ Since $E(Y | X = x)$ is linear in x , we get

$$E(Y | Z_1, Z_2) = E(Y | X) = E^*(Y | X).$$

Sample Analog

- ▶ Data: $(y_i, z_{i1}, z_{i2}), i = 1, \dots, n$.
- ▶ Dummy Variables: $x_{i,jk} = 1(z_{i1} = \lambda_j, z_{i2} = \delta_k)$,

$$y = \begin{pmatrix} y_1 \\ \vdots \\ y_n \end{pmatrix}, \quad x_{jk} = \begin{pmatrix} x_{1,jk} \\ \vdots \\ x_{n,jk} \end{pmatrix}.$$

- ▶ Least Squares:

$$\min_b \sum_{i=1}^n \left(y_i - \sum_{j=1}^J \sum_{k=1}^K b_{jk} x_{i,jk} \right)^2$$

- ▶ This gives:

$$b_{jk} = \frac{\sum_{i=1}^n y_i x_{i,jk}}{\sum_{i=1}^n x_{i,jk}} = \bar{y} | \lambda_j, \delta_k.$$

Polynomial regressors

- ▶ Assume

$$E(Y|Z_1 = s, Z_2 = t) \cong \beta_0 + \beta_1 s + \beta_2 s^2 + \beta_3 t \cdot s + \beta_4 t + \beta_5 t^2.$$

- ▶ Example: Jacob Mincer, *Schooling, Experience and Earnings*, 1974, Table 5.1; 1 in 1000 sample, 1960 census; 1959 annual earnings; $n = 31093$;
- ▶ $y = \log(\text{earnings})$, $s = \text{years of schooling}$, $t = \text{years of work experience}$;

$$\hat{y} = 4.87 + .255s - .0029s^2 - .0043t \cdot s + .148t - .0018t^2.$$

Predictive Effect

- Returns to college:

$$\begin{aligned} E(Y|Z_1 = 16, Z_2 = t) - E(Y|Z_1 = 12, Z_2 = t) \\ \cong \beta_1 \cdot 4 + \beta_2(16^2 - 12^2) + \beta_3 \cdot 4 \cdot t. \end{aligned}$$

- Returns to high school:

$$\begin{aligned} E(Y|Z_1 = 12, Z_2 = t) - E(Y|Z_1 = 8, Z_2 = t) \\ \cong \beta_1 \cdot 4 + \beta_2(12^2 - 8^2) + \beta_3 \cdot 4 \cdot t. \end{aligned}$$

Plugging in the estimates

Experience	Returns to college	Returns to high school
0	.70	.79
10	.52	.62
20	.35	.44

Questions for you

Verify this.

From predicting $\log W$ to predicting W

- ▶ Suppose $E(\log W | Z)$ is a linear function:

$$E(\log W | Z) = \beta_0 + \beta_1 Z.$$

- ▶ Define $U = \log W - \beta_0 - \beta_1 Z$, so $E(U | Z) = 0$.

▶

$$W = \beta_0 + \beta_1 Z + U$$

$$\Rightarrow W = \exp(\beta_0 + \beta_1 Z) \cdot \exp(U).$$

- ▶ If U and Z are independent, $E[\exp(U) | Z] = E[\exp(U)]$ and

$$\frac{E(W | Z = d)}{E(W | Z = c)} = \exp[\beta_1(d - c)] \cong \beta_1(d - c) + 1,$$

$$100 \left[\frac{E(W | Z = d)}{E(W | Z = c)} - 1 \right] \cong 100\beta_1(d - c).$$

Mincer model

- ▶ Compound Interest (Δ = fraction of one year):

$$\$1 \rightarrow \$(1 + r\Delta) \rightarrow \$(1 + r\Delta)^2 \rightarrow \$(1 + r\Delta)^3 \rightarrow \dots$$

- ▶ Annual Return $(1 + r\Delta)^{1/\Delta}$.

$$\lim_{x \rightarrow 0} \frac{f(1+x) - f(1)}{x} = f'(1); \quad \lim_{x \rightarrow 0} \frac{\log(1+x)}{x} = 1$$

$$\lim_{\Delta \rightarrow 0} \log[(1 + r\Delta)^{1/\Delta}] = r \cdot \lim_{\Delta \rightarrow 0} \frac{\log(1 + r\Delta)}{r\Delta} = r$$

- ▶ Thus

$$\lim_{\Delta \rightarrow 0} (1 + r\Delta)^{1/\Delta} = \exp(r) \approx 1 + r$$

- ▶ Works for small r :

$$\exp(.06) = 1.062; \quad \exp(.35) = 1.42; \quad \exp(.70) = 2.01.$$

- ▶ $PV(S)$ = present value at $t = 0$ of earning 0 while in school for an additional S years and then earning $W(S)$ for a very long time:

$$PV(S) = W(S) \int_S^{\infty} \exp(-rt) dt = W(S) \cdot \exp(-rS)/r.$$

- ▶ Returns such that students are indifferent about dropping out:

$$PV(S) = PV(0) \Rightarrow W(S) \cdot \exp(-rS)/r = W(0)/r,$$

- ▶ Thus:

$$\log(W(S)) = \log(W(0)) + rS.$$

- ▶ Linear Predictor:

$$E^*(Y|1, S) = \gamma_0 + (r + \gamma_1)S,$$

where $Y = \log(W(S))$, $A = \log(W(0))$, and

$$E^*(A|1, S) = \gamma_0 + \gamma_1 S.$$