Conducting Sparse Feature Selection on Arbitrarily Long Phrases in Text Corpora with a Focus on Interpretability

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Abstract: We propose a general framework for topic-specific summarization of large text corpora, and illustrate how it can be used for analysis in two quite different contexts: an Occupational Safety and Health Administration (OSHA) database of fatality and catastrophe reports (to facilitate surveillance for patterns in circumstances leading to injury or death), and legal decisions on workers' compensation claims (to explore relevant case law). Our summarization framework, built on sparse classification methods, is a compromise between simple word frequency-based methods currently in wide use, and more heavyweight, model-intensive methods such as latent Dirichlet allocation (LDA). For a particular topic of interest (e.g., mental health disability, or carbon monoxide exposure), we regress a labeling of documents onto the high-dimensional counts of all the other words and phrases in the documents. The resulting small set of phrases found as predictive are then harvested as the summary. Using a branch-and-bound approach, this method can incorporate phrases of arbitrary length, which allows for potentially rich summarization. We discuss how focus on the purpose of the summaries can inform choices of tuning parameters and model constraints. We evaluate this tool by comparing the computational time and summary statistics of the resulting word lists to three other methods in the literature. We also present a new R package, **textreg**. Overall, we argue that sparse methods have much to offer in text analysis and is a branch of research that should be considered further in this context. © 2016 Wiley Periodicals, Inc. Statistical Analysis and Data Mining: The ASA Data Science Journal, 2016

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1. INTRODUCTION

Regularized high-dimensional regression can extract meaningful information from large text corpora by producing key phrase summaries that capture how a specific set of documents of interest differs from some baseline collection. This text summarization approach has been called Concise Comparative Summarization (CCS) [1], underscoring two fundamental features of this tool: (i) the comparison of a collection of documents to a baseline set in order to remove generic terminology and characteristics of the overall corpus; and (ii) the resulting production of a short, easy-to-read summary comprised of key phrases. Such summaries can be useful for understanding what makes a document collection distinct and can be used to inform media analysis, understand incident reports, or investigate trends in legal decisions.

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Many classic methods of text summarization tend to focus on single words or short phrases only. Approaches such as latent Dirichlet allocation (LDA) [2] also do not extend naturally to phrases. On the other hand, one regression-based method [3,4] that does allow for longer phrases does not allow for rescaling of the counts of phrases in the text based on the overall frequency of appearance of such phrases, which can negatively impact summary quality. In this paper, we merge two CCS approaches to allow for rescaled, arbitrary-length key phrases that can include gaps. We briefly discuss how this is done below. Our new CCS tool can be easily used via our new R package, **textreg**, which allows for rapid exploration of text corpora of up to a few gigabytes in size.

Even given these tools, when a researcher desires to conduct a specific analysis, he or she is faced with many choices. In particular, the implementation and regularization of the regression itself can be done in several ways, and the impact of choosing from among these ways is one of the foci of this paper. In particular, we argue

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that if the researcher has specific goals for interpretation in mind, these goals can inform the choice of the tuning parameters. For example, when faced with a corpus where only a few documents are of interest and the rest are to be used as a baseline, a researcher may choose to allow only positive weights on phrases, in order to simplify interpretation. Similarly, choice of the tuning parameter can be governed by a researcher's level of interest in pruning rare phrases. We also offer a method for testing for a significant relationship between the text and document type, which also provides a threshold regularization value.

We compare this tool with other related state-of-the-art methods. First, we compare it with multinomial inverse regression (MNIR) [5], a text regression method that is primarily designed to be distributed across many cores in order to handle massive data. We also compare it with a classic Lasso approach (see, e.g., [6]), which is similar to this method run on precomputed document-term matrices without some of the flexibility. We finally compare to the original Ifrim et al's. [3,4] method, which is one of the building blocks of this work. In these comparisons, we investigate computation time, prediction accuracy, and different features of the resulting word lists. The different approaches give very different types of lists, and we hope this work gives some guidance to the practitioner as to how to sort through the options.

As a case study, we use this tool to examine a large collection of occupational fatality and catastrophe reports generated by the Occupational Safety and Health Administration (OSHA) in the United States. As a motivating example, we examine hazardous exposure to methylene chloride, a neurotoxin, during bathtub refinishing operations. In 2013, OSHA and the National Institute for Occupational Safety and Health (NIOSH) jointly issued a Hazard Alert calling attention to a recurring pattern of this nature following the deaths of least 14 workers since 2000 in related circumstances. However, the sheer volume of information describing occupational fatalities and catastrophes (Fat/Cats) may have initially obscured this pattern in the years preceding its detection. Although OSHA maintains a database of narrative reports describing Fat/Cats, similar patterns of preventable exposure to occupational hazards may be difficult to identify efficiently through manual review alone, given the large number of narratives in this database. Thus, using methylene chloride as a case study, we consider whether text mining techniques can help identify important patterns in circumstances of hazardous exposures.

In our framework, a summary list of key words and phrases ideally represents and reveals the overall content of a collection of narrative reports. For example, one summary for all narratives related to 'Methylene Chloride' contained the words 'bathtub' and 'stripper'. To qualitatively evaluate our tool, we manually examine these words and phrases in the context of the original reports and consider whether our text summarization tool effectively characterizes the circumstances of the bathroom refinishing fatalities.

In general, we explore whether we can construct text mining algorithms that, when applied to an entire corpus, can uncover 'needles in the haystack' patterns such as the connection between bathroom refinishing and overexposure to methylene chloride. At this stage, we are not focused on rates or relative risks of particular patterns; we are, instead, focused on the crude detection of textual patterns that may represent meaningful information about how certain types of injuries and fatalities occur. Such findings, even if they involve only a few recorded deaths or injuries, may facilitate the prevention of many future fatalities, particularly in the context of emerging hazards.

We also examine our tool's ability to extract information from a collection of legal decisions from the Employees' Compensation Appeals Board (ECAB), which handles appeals of determinations of the Office of Workers' Compensation Programs (OWCP) in the U.S. Department of Labor (DOL). Here we investigate what information we can extract about different categories of cases. In particular, we examine cases involving a question as to whether the work environment caused a mental health condition (an 'emotional condition', in the parlance of ECAB). We find that while the CCS tool does extract meaningful information relating to the cases of interest, further work needs to be done to obtain more nuanced summaries.

Overall, the CCS approach does allow for exploration of text and does extract meaningful information. Extending the earlier fixed-length phrase tools to allow for longer phrases and phrases with gaps does increase the expressive capability of the summaries. The methods for picking a tuning parameter, while possibly a bit aggressive and conservative, do provide an alternative paradigm for data analysis with an eye to extracting human meaning from text.

2. OVERVIEW OF SUMMARIZATION

This paper extends the concept of CCS discussed in [1], incorporating a prior approach proposed by [3] to result in an overall improved methodology. CCS involves comparing a prespecified set of documents to a baseline set. One can think of it as a regularized regression of some labeling of the *m* documents (normally +1 and -1) onto the collection of all possible summary key phrases. For example, in one analysis we label documents relating to incidents involving carbon monoxide (CO) as +1 and the remaining documents as -1. Each potential key phrase, or

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'feature', is considered to be a covariate in the regression, and is in principle represented as an m-vector of measures of the feature's presence (e.g., appearance count) in each of the m documents. By using sparse regularization, only a small subset of these possible summary key phrases is selected. These phrases are taken as the final summary. So, for example, the resulting phrases of our CO-related regression would ideally indicate what is different about CO-related events when compared to other workplace injuries and fatalities. At root, we are taking those phrases most useful for classifying a set of documents by the given labeling as the summary of the positive set of documents as compared to the negative, baseline, set.

It is worth emphasizing that the focus is not predictive quality; the selected features themselves are the object of interest, and the quality can only be measured by their usefulness in the original human interpretationbased question that motivated the exploration. Thus, these methods in principle require human analog experiments to validate their findings. This can be done; see [7,8], or [1] for examples. We are using text classification tools, but the classification is a byproduct. There are many possible choices for how to implement this regression including whether to use logistic or linear regression, and whether to rescale the frequency appearance of the phrases before regression. Prior work has shown that rescaling phrase frequency is quite important; failing to appropriately do so can result in summaries that have little informative content even while predictive accuracy is maintained. This is not surprising; term frequency in text is a known and serious concern when data mining large text corpora (e.g., [1,9]), as was first illustrated in the information retrieval literature (e.g., [10,11]).

In text classification and, by extension, key-phrase extraction via text classifiers, there is some desire to allow for phrases as well as unigram (single word) features. One approach is to calculate and use all phrases up to n words long as the overall feature set. For long phrases, this can quickly become intractable, as there is a blowup in the number of possible phrases a corpus may contain. To solve this problem, Ifrim et al. [3,4] allow for arbitrary length phrases by generating the features 'on-the-fly' as part of the optimization. As an added benefit, this approach easily allows for 'gaps', i.e. phrases with wildcard words, which greatly enhances the potential expressiveness of the resultant summaries.

Ifrim et al.'s algorithm, based on the work of [12] and, even earlier, [13], fits an elastic-net-penalized logistic regression with the features consisting of the entire space of all phrases (also see [14] or [15] for other examples of regression on text and [16] or [17] for an overview of elastic nets and other regularized regression methods in general). Ifrim et al. initially propose an algorithm to solve

a penalized logistic regression of

$$\hat{\beta} = \arg\min_{\beta} \sum_{i=1}^{m} y_i c'_i \beta + \log\left(1 + \exp(y_i c'_i \beta)\right) + CR(\beta),$$

with $\beta \in R^p$, c_i being the feature vector for document *i* with the c_{ii} as binary indicators of the presence of feature j in document i, y_i the -1/1 class label, p the number of features including all phrases, C a regularization tuning parameter, and $R(\beta)$ some regularization function. They later extend this to allow for alternate loss functions such as a hinge loss. However, they do not allow for rescaling features. By modifying their methods, we show how rescaling can be incorporated into their overall approach. They also do not allow for an intercept term, which can introduce difficulty with the summarization process if the number of positive features is not close to 50%. We extend their algorithm to allow for a (nonpenalized) intercept term as discussed in [1]. We implemented these modifications by extending their code, and then wraped the resulting algorithm in a new R package, textreg, to make it easier to use in a rapid and exploratory manner. We also provide some useful tools and visualizations to help researchers understand the resulting summaries.

The core idea behind the algorithm is a greedy coordinate descent coupled with a branch-and-bound algorithm. With each step, the algorithm searches the entire feature space for the feature that has the highest gradient at the current β . This is obviously a very large search space, but it can be pruned using a relationship that bounds the size of a gradient of a subphrase by a known calculation on any parent phrase. In the search, we track the current optimal feature, and then for each new feature considered, if the bound on all the children of that feature is smaller than the current optimum, prune all those children from the search.

2.1. Related Work

CCS is distinct from classification. Classification is focused on sorting documents, such as for attributing authorship [18,19] or counting types of news events [20,21]. Text classification has been attempted using a wide array of machine learning methods such as naive Bayes, linear discriminant analysis, or support vector machines [22], which easily allow for large numbers of features (the words and phrases) to be incorporated. For comparisons of these different methods on text, see [19], [23], or [24]. For Support Vector Machines (SVMs) as an approach in particular for text, see the book of [25]. For such methods and these evaluations, however, the features themselves are not of primary interest, but classification is. We, instead, attempt to extract meaning from documents by contrasting

sets to each other. This is most similar to key phrase extraction, a literature in its own right. See, for example, [26–28].

Interpreting text is a difficult task, and can be done in a variety of ways. For example, [29] use text to predict roll call votes with an LDA algorithm [2] in order to understand how the language of law is correlated with political support. Ref. [7] model political text to explore who dominates policy debates. Truly validating a finding is generally quite difficult and requires a great deal of effort to do properly, as these papers illustrate quite well. Our tools are primarily intended for exploration; validation is not within CCS's scope without additional human validation efforts or alternative techniques.

Of the many approaches to text analysis, variations of LDA [2] in particular have recently been widely investigated and used. These consist of a Bayesian hierarchical model that describes documents in a corpus as a mixture of some number of different distributions on the words. They can reveal structure in a corpus and have been shown to capture human meaning. However, generating novel models for specific circumstances is difficult. Even mild changes to the model can be technically quite challenging and consist of an entire research topic in its own right. They are either computationally expensive or only solved approximately (via, e.g., variational methods). In the spirit of diversity in research approaches, we take a different path.

This is not to say that using sparse regression methods on text is new; see, for example, [23,30,31]. [32] use sparsity to model topics by representing them as small collections of phrases that stand out against a background. [33] showcase several methods, such as sparse principal components analysis (PCA), to investigate large corpora. There are many others.

One aspect of our approach, which we believe is novel, is allowing for complex features in the form of overlapping phrases, especially phrases with wildcard words, while maintaining the ability to rescale features. This allows great flexibility in the expressiveness of the possible summaries generated, and it is not obvious how to naturally extend methods such as LDA, which rely on a generative model where words are picked i.i.d. from some distribution, to do this.

3. RESCALED N-GRAM REGRESSION

Initial methods regress the y_i on the c_i , where c_i is either the vector of counts with c_{ij} being how often phrase jappears in document i or of binary indicators of appearance, with the elements $c_{ij} \in \{0, 1\}$ indicating the presence of phrase j in document i. This can be problematic in that common phrases (e.g., 'the', or, less obviously, 'usually') end up having much higher variance than less common ones, and thus it is easier to pick them up due to random fluctuations. See Section 4 for further discussion.

Rescaling the features can correct this as pointed out in, e.g. [34]. In particular, L^q -rescaling (for $q \ge 1$) transforms the vectors c_i into new covariate vectors x_i as

$$x_{ij} = \frac{c_{ij}}{z_j}$$
, where $z_j \equiv \left(\sum_{i=1}^n c_{ij}^q\right)^{1/q}$.

This is similar to standardizing columns in a Lasso regression; if you do not, then phrases with high variability need smaller coefficients to have similar impacts on prediction. This makes them 'cheaper' under the regularization and therefore appears more frequently due to random chance.

Once our feature space has been standardized with each phrase having an L^q -length of 1, we regress our y onto these rescaled x and an intercept. This is a high-dimensional problem with $p \gg m$. As we want a small number of phrases, we use a sparse regularization L^1 penalty. We also use a squared hinge loss to obtain

$$\mathcal{L}(\beta) = \sum_{i=1}^{n} \left[\left(1 - y_i (\beta_0 + x'_i \beta) \right) \vee 0 \right]^2 + C \sum_{j=1}^{p} |\beta_j|,$$

with $a \lor b$ denoting the maximum of a and b. For this loss function, an *over*-prediction is not penalized. From a prediction standpoint, we wish $\beta_0 + x'_i\beta = y_i$; if we fall short, we have quadratic loss; if it does, the loss is zero; and if we overshoot, we still have zero loss. There is no penalty for 'over-predicting' a document's label. We use the squared hinge loss, as this is similar to the Lasso, shown to be effective in [1], but also monotonic, which is needed for the optimization algorithm. Also note that the penalty term does not include the intercept, β_0 .

To generalize this framework, taking notation from Ifrim et al., let our loss for an individual document be $\xi(m_i)$ with $m_i = y_i(\mu + x'_i\beta)$. We can use any monotonic loss functions with $\xi' \leq 0$ everywhere. The squared hinge loss from above is $\xi(m) = ((1 - m) \vee 0)^2$; this is very similar to an ordinary least-squares (OLS)-type penalty of $(1 - m)^2$. Logistic would be $\xi(m) = \log(1 + e^{-m})$.

Regardless of the choice of ξ , the loss term can be expressed in the original counts as

$$\mathcal{L}(\beta) = \sum_{i=1}^{n} \xi \left(1 - y_i \left(\beta_0 + \sum_{j=1}^{p} \frac{c_{ij}}{z_j} \beta_j \right) \right) + C \sum_{j=1}^{p} |\beta_j|$$
$$= \sum_{i=1}^{n} \xi(m_i) + C \sum_{j=1}^{p} |\beta_j|.$$

We then obtain $\hat{\beta}$ as

$$\hat{\beta} = \arg\min_{\beta} \mathcal{L}(\beta). \tag{1}$$

Alternatively, by letting $\tilde{\beta}_j \equiv \beta_j/z_j$, we can move the z_j to the penalty term and regress on the counts c_i ; this gives the identical loss as seen by considering $\tilde{m}_i = y_i(\mu + c'_i \tilde{\beta})$ on the rescaled columns: scaling a column by z_j is the same as penalizing the associated β_j by z_j . This has ties to weighted Lasso approaches such as the adaptive Lasso [35], in that we now have feature-specific penalties. The gradients change, however, which can affect the optimization (see Appendix B).

Algorithm 1 Greedy Coordinate Descent
$\beta = \emptyset$
$features = \emptyset$
while Not Converged do
$\beta[intercept] = updateFeature(intercept)$
f = findHighestGradient
features.add(f)
$\beta[f] = updateFeature(f)$
end while

Solving Eq. (1) is done with greedy coordinate descent. See Algorithm 1. For greedy coordinate descent, we repeatedly find the feature with the highest gradient, and then optimize its corresponding β_i with a line search over the loss function. Because this is a convex problem, this will converge on the global maximum as each iteration will decrease $\mathcal{L}(\beta)$ and since the gradient along all the coordinates can only be 0 if we are at a maximum. For a proof see, e.g. [36] or [37]. We keep a cache of all the nonzero features in our model; we do not need to ever calculate or store all possible features. The main computational cost of the algorithm is in finding the feature with the largest gradient. To do this, we dynamically generate the features by exploiting the nested structure of any multiword phrase having a smaller phrase as a prefix. This inner algorithm is shown in Algorithm 2. Here we first examine all unigrams, then bigrams, and so forth, until there are no more eligible phrases. We first calculate the gradient for all unigrams and enter them into the queue. Phrases in the queue are placeholders for their family of superphrases. When we pull a phrase out of the queue, we check to see if we can prune all of its children, and if we cannot, we determine the phrases' children, calculate gradients for these children, and finally enter them into the queue. This algorithm would work without pruning, but if we were able to prune all the at-zero children of a feature before examining them, we could achieve large speed-ups. And, indeed, some pruning is possible due to a trick of bounding a child's gradient based on the structure of its parent, although the rescaling makes keeping this bound tight more difficult than in the original Ifrim et al. presentation. The main idea is, if a bound on the gradients of a family of features is less in magnitude than our current best gradient, we can prune them all. We discuss finding such bounds next.

Algorithm 2 findHighestGradient
features = all non-zero features so far.
$best_f = \arg\max_{f \in features} gradient(f)$
$u_1, \ldots, u_{p_1} = $ all unigrams in dictionary
Q = queue(), a queue of all features to check
for $u \in u_1, \ldots, u_{p_1}$ do
if $gradient(u) > gradient(best_f)$ then
$best_f = u$
end if
Q.add(u)
end for
while Q is not empty do
f = Q.next()
if not can $Prune(f, best_f)$ then
for $c \in \operatorname{children}(f)$ do
if $gradient(c) > gradient(best_f)$ then
$best_f = c$
end if
Q.add(c)
end for
end if
end while

3.1. Bounding Gradients

Take any feature j with corresponding appearance pattern across the documents c_j . For any feature k with feature j as a prefix, we know that $c_{ki} \leq c_{ji}$ for i =1, ..., n, which we write as $c_k \leq c_j$. We also know that $c_{ki} \geq 0$ for i = 1, ..., n, because they are counts, so $0 \leq c_k$. That is, given a phrase j, any phrase with phrase j as a prefix can only have a count vector bounded between 0 and phrase j's count vector.

During our search, we consider phrases from shorter to longer. For each phrase j, we, based on that phrase's appearance pattern in the text, calculate a bound b_j on the magnitude of the highest gradient a 'best case' hypothetical superphrase with that prefix could have. If this b_j is smaller than the current best achieved gradient $|\Lambda|$, then we can prune from consideration all phrases with phrase j as a prefix because, if $b_j \leq |\Lambda|$, we have for any

superphrase k of j

$$\left|\frac{d}{d\beta_k}L(\beta)\right| \le b_j \le |\Lambda|.$$

Therefore we want b_j to be as small as possible, i.e. tight, to make phrases easier to prune.

As derived in Appendix B, one such overall bound is, for any $q \ge 1$,

$$\left| \max_{k:0 \le x_k \le x_j} \frac{d}{d\beta_k} L(\beta) \right| \le 0 \bigvee \max\left\{ \left(\sum_{y_i = -1, c_{ij} > 0} \left| \xi'(m_i) \right|^r \right)^{1/r} \right\}$$
$$\times \left(\sum_{y_i = 1, c_{ij} > 0} \left| \xi'(m_i) \right|^r \right)^{1/r} \right\} - C,$$

where 1/r = 1 - 1/q. For any phrase *j*, these bounds can be computed by summing over only those documents that have phrase *j*, rendering them computationally tractable. Because the rescaling allows for theoretically very predictive yet relatively rare phrases, these bounds are unfortunately quite loose, making it hard to substantially trim the search tree. One possible avenue for improvement would be to integrate the preprocessing step suggested by [38].

As special cases, q = 1 gives $r = \infty$ and a bound of the maximum $|\xi'(m_i)|$ for any document *i* that has phrase *j*, and $q = \infty$ gives r = 1 and a bound of the maximum of the two sums of the $|\xi'(m_i)|$ across the negative and positive documents containing phrase *j*—which is Ifrim's bound, corresponding to no rescaling, up to the scaling of the maximum occurrence of the phrase in any single document.

All of the above is easily extended to an elastic net [17]. See note in Appendix B.

4. CHOICES OF RESCALING AND ADDITIONAL CONSTRAINTS

Choices of rescaling (e.g., the q in the L^q -rescaling) and further restrictions on the optimization problem can focus the CCS tool on different aspects of the summary. We can seek to generate summaries with more general or more specific words for example, or enforce a contrast of a target set to a larger background set which eases interpretability. We discuss how to do this in the following subsections.

4.1. Rescaling

Phrases vary greatly in their overall appearance in text, with a very long tail of words and phrases that appear in nearly every document, and the bulk of phrases appearing only one or two times. A phrase's rate of appearance is connected to its underlying variance if we represent the phrase with its count vector. This can cause problems when selecting the most meaningful phrases. In particular, common phrases can easily dominate because they have greater variance. Typically, this is handled with stop-word lists, which are lists of words that are a priori deemed lowinformation and dropped before analysis. For a thorough discussion of this, see [34]. And, as [34] discusses, stopword removal is finicky, not general, and imperfect. They cannot easily be adapted to differing contexts or languages. Furthermore, how to implement stop-word removal when phrases are the object of interest is unclear. Rescaling, however, can not only serve the function of a stop-word list but also do a superior job [1].

Rescaling is critical, as is widely known in information retrieval. Without rescaling, stop words are easily selected by virtually all text mining procedures. Even with stop words being dropped, typically the 'runner up' most common phrases are then selected, primarily due to random variation in their appearance pattern. To see this, test nearly any off-the-shelf text mining tool without removing stop words first; more often than not, these methods will fail, and their results will be dominated by these lowinformation words. Stop-word lists are a hard-threshold solution when a soft-threshold tapering is more appropriate. Rescaling offers such a tapering approach. The question then becomes which rescaling to use, or, alternatively, how much tapering do we want. We use L^q -rescaling because it offers a class of choices (and integrates well into the gradient descent algorithm). With L^q -rescaling, different choices of q weight phrases relatively differently, allowing for focus on more common or uncommon phrases at the desire of the researcher.

Overall, lower q means generally higher normalization factors Z, which will change the appropriate C for regularization. The main point of concern, however, is the relative sizes of the weights for rare phrases compared to common phrases. In general, an $L^{1.25}$ -rescaling heavily penalizes common phrases while an L^3 -rescaling does not. On the other hand, L^3 -rescaling penalizes rare phrases slightly more than lower choices of q. To illustrate, see Fig. 1; here we consider a sequence of phrases that appear once in each of m out of 1000 documents. The different series of weights have been rescaled, so a phrase that appears 5% of the time (50 times) has the same weight for all choices.

4.1.1. tf-idf rescaling.

A related strategy for rescaling is tf-idf rescaling, from information retrieval [10]. It is typically something like the following (variations exist):

$$x_{ij} = \frac{c_{ij}}{n_i} \cdot r_j$$
 with $r_j \equiv \log \frac{n}{d_j}$,

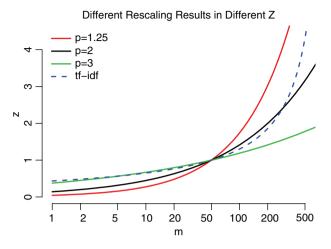


Fig. 1 Impact of different choices of rescaling. Here we see different rescaling factors z for phrases with a single appearance for each of m documents out of 1000 total documents. More common phrases are penalized more greatly relative to rare phrases. [Color figure can be viewed in the online issue, which is available at wileyonlinelibrary.com.]

with d_j being the total number of documents containing the phrase *j*. It differs in that it corrects for each document length with the n_i ; we do not do so. If documents are roughly the same length, this becomes less relevant. Tf-idf also puts more of an extreme difference between weights for rare and common phrases, scaled by the total number of documents. For rare and mid-range phrases, the tf-idf rescaling is similar to a L^q -rescaling with large q.

4.2. Interpretation and Negative Coefficients

CCS returns a list of phrases with nonzero coefficients β . Interpreting these coefficients can be quite difficult. Just as in OLS, a model-based interpretation for β_k would be that changing feature k by 1 would change the prediction by β_k *holding other features fixed* (which, given normalization, means changing feature k by a count of 1 would change the outcome by β_k/z_k). However, given the lack of a wellmotivated model in our context, interpreting the magnitude of these coefficients is somewhat dubious.

Nonetheless, we still wish to interpret the sign of the coefficients: positive indicates that a feature is associated with our positive class of documents, and negative indicates the negative class. When the negative group is a baseline, however, it is not an object of direct interest. This is especially the case if it is much larger and more diverse than the positive class. In this case, the regression is ideally subtracting out baseline characteristics, leaving the researcher with what makes the positive class distinct and noteworthy. Here, interpreting negative coefficients can be difficult. One interpretation would be that such features are 'conspicuous in their absence'.

Unfortunately, even when there is a mix of positive and negative features, we can still end up with unclear interpretations of the sign due to the 'holding other features constant' aspect of the above. For example, a negative weight for a feature might be offsetting the positive weight for a highly correlated alternate feature, and in fact both features may have a positive correlation with the labeling. In this case, interpreting the negative sign as, e.g. conspicuous in its absence, is erroneous. It is more accurate to say the feature is conspicuous in its absence given its normal association with the second feature. This can be hard to communicate.

One solution is to extend the optimization to consider only the set of positive β , forcing negative coefficients to not exist. This is an easy extension of the above algorithm: simply drop the lower bound on the gradient search and truncate any line-search update of a β_k at 0.

This is not to say that negative features are useless. For example, if we allow negative features and find that all the coefficients are positive, it would suggest that the positive group has a clearer signal than the negative group. Only phrases found in the positive group are differentiating the groups. This might suggest distinct language use, larger vocabulary, or specific turns of phrase on the part of the positive group, which could be of interest in its own right.

5. PICKING THE REGULARIZATION PARAMETER

For most regularized regression settings, picking the regularization parameter C is a notoriously difficult problem. In general, higher Cs lead to fewer features, i.e. more concise summaries. Low C summaries will be more verbose. However, an overly low C allows for overfitting, which in our context means obtaining features that are detected solely due to random fluctuations in the appearance patterns of phrases. We need to ensure that C is sufficiently large to mostly prune out such noise.

Classically, selection of C is done using methods such as cross-validation to optimize prediction accuracy on out-ofsample predictions. As prediction is not our primary focus, we look for other methods to select C that enhance the quality or interpretability of the summaries generated. The lack of appropriateness of prediction accuracy is somewhat motivated by the literature; prediction accuracy is, for example, not the same as model selection, as is illustrated by the choice between AIC and BIC (Akaike and Bayesian information criterion) selection methods in regression. We present two methods, rooted in the goals of CCS, to select C. The first is to conduct a permutation test to select a C that gives a statistically significant summary in that the summary being nonempty indicates the presence of

systematic differences in the text between the positively and negatively marked documents. The second is to select a minimum C to guarantee the pruning of very rare phrases. We discuss how one might select which approach to use in the discussion after the case studies, below.

5.1. A Permutation Test on the Summary

One might wonder if the phrases returned by CCS are simply due to random chance. There are so many different phrases, so it is reasonable to believe some will randomly be associated with any document labeling. We can control this with a permutation test. This is an exact test, and the resulting *p*-value is easy to interpret.

To test whether it is meaningful to generate a summary at all, repeatedly randomize the labeling across the documents, regress, and find the corresponding C^* that zeros out all the features, given our random permutation of the labels. This gives a 'null distribution' of what *C* is appropriate if there were no signal. Finally, compare our originally observed C^{obs} to this distribution of fake C^* s. We calculate a *p*-value of

$$p = \Pr\left\{C^{obs} \ge C^*\right\}.$$

If p is small, we conclude that the needed regularization to zero out our originally observed signal is greater than that for a random signal; i.e. there is a relationship between the labeling and the text. Similarly, if we pick a C that is at the 95th percentile of our permutation distribution, we are 95% confident that the resulting summary being nonempty is due to the relationship of the labeling and the text, and not due to random chance.

The individual phrases, however, are *not* specifically tested as being significant; it is possible that they would change, for example, given mild perturbations to the data. Nevertheless, this test provides a useful minimum bound for the final C. Any C lower than this bound could result in a nonempty summary purely due to random chance.

In a similar manner, one can check the coherence of the final summaries by generating summaries under different permutations of the labeling (potentially adjusting C with each iteration to get similar length summaries for all permutations) and creating a list of lists with the actual summary randomly inserted. If humans can then reliably pick out the actual summary from the fake ones, this is indication that the structure of the summary is not due to random chance. This idea is based on assessing the quality of EDA (exploratory data analysis) visualizations; see [39].

5.2. Pruning Rare Phrases

Most potential phrases are rare, showing up only a handful of times in even very large corpora. Selecting from such phrases introduces a severe multiple testing problem, and we seek to appropriately regularize the regression with C to solve it. In particular, rare phrases that show up only a few times can be selected if they happen to fall only in the positive set. More generally, with improper rescaling of features, a term that shows up once in the positive examples with a high count and several times in the negative examples with a low count can also be selected. This often is contrary to the interpretive goal behind selecting predictors. We want phrases that are general summaries, informing the researcher of aspects across multiple documents. These problems can be partially remedied by proper selection of the tuning parameter C. Here we investigate minimal C to guarantee that quite rare phrases are dropped.

We find such *C* by investigating the so-called perfect predictors. Consider a feature *j* such that $c_{ij} = 0$ if $y_i = -1$ and $c_{ij} = 1$ for some of the documents, where $y_i = +1$. For the moment, assume we do not have multiple counts in any document. This is a perfect predictor, predicting $r \equiv \sum_i c_{ij} \le s$ of the *s* positive documents. These perfect predictors could be used to identify a subset of the positive examples while incurring no loss for the negative examples. The only cost of including such a predictor is due to the regularization term *C*. If we set *C* high enough, the cost will be prohibitive and we will not select. In fact, the cut-off of

$$C^* = 2(1-\mu)r^{1-1/q}$$

with $\mu = \frac{1}{m}\hat{Y}_i$, with \hat{Y}_i being the prediction for document *i* without any such hypothetical phrase, suffices. See Appendix B for a derivation.

For this C^* , any perfect predictor of r documents will be pruned. For comparison, see Table 1, which shows (for both $\mu \approx -1$, the case of few positive and many negative examples, and $\mu \approx 0$, the case of roughly equal numbers of positive and negative documents) the needed cutoffs for r = 1, 2, 4. This cutoff will generally be overly aggressive; if other predictors also predict these documents, then the gain of including the perfect predictor is potentially less.

5.2.1. No rescaling

As discussed at the end of Section 3.1, no rescaling is bad for appropriately handling common phrases. No rescaling is also bad for appropriately handling rare ones, as we can see by its connection to the infinity-norm and the top row of the table. No rescaling of features makes it very difficult to prune perfect predictors.

5.2.2. Singleton predictors

As a special case, 'singleton predictors' are those that appear only once in the entire corpus, and appear for a

 Table 1.
 Needed C to prune rare perfect predictors.

			$\mu \approx -1$			$\mupprox 0$			
q	C^*	r = 1	r = 2	r = 4	r = 1	r = 2	r = 4		
$\overline{\infty}$	$2(1-\mu)r$	4	8	16	2	4	8		
4	$2(1-\mu)r^{3/4}$	4	pprox 6.7	≈ 11.3	2	≈ 3.4	≈ 5.6		
2	$2(1-\mu)\sqrt{r}$	4	$4\sqrt{2}$	8	2	$\sqrt{2}$	4		
4/3	$2(1-\mu)r^{1/4}$	4	pprox 4.75	≈ 5.7	2	≈ 2.4	pprox 2.8		
1	$2(1-\mu)$	4	4	4	2	2	2		

positive example. Normally, if such a rare phrase appears once in a positive example, it can be pruned as described above. Regardless of q, in order to remove singletons that predict for a document with no other predictors, we must have $C \ge 2(1 - \mu)$.

This can be generalized somewhat. Consider, for q = 2, whether the count of a phrase for a single positive document is *s* and the count for *t* negative documents for that same phrase are 1 each. The L^2 normalizing constant is then

$$Z = \sum_{i=1}^{n} c_{ki} = s^2 + t$$

and $x_{ki} = s/\sqrt{s^2 + t} \approx 1$ for the single positive document and $x_{ki} = 1/\sqrt{s^2 + t} \approx 1/s$ for the few negative documents. This is approximately the same as the singleton phrase circumstance, and will therefore be pruned as above.

Proper selection of the tuning parameter is a better approach for pruning than cutting by dropping phrases with low counts (ignoring computational issues), as it can also prune near-singleton phrases with high counts. This circumstance indeed arises. In a study of the Fat/Cats corpus, below, the word 'lion', used eight times for a report involving a plague-ridden mountain lion corpse (positive example) and nine times in various negative examples, was kept as a predictor in the final summary of 'Disease' when *C* was too low (C = 0.5).

5.3. Regularization with Cross-Validation

The traditional form of selecting a tuning parameter is via cross-validation, where some metric of predictive performance is optimized. For example, in our context we could calculate the predicted labeling of set-aside documents and select C such that the average squared distance between predictions and actual is minimized. As we will see, this tends to give longer lists that can be less interpretable, as a more important signal can be buried among less relevant terms. Generally, predictive performance is not directly a measure of our primary focus: the interpretability and significance of the selected phrases.

5.4. Regularization with Early Stopping and the Elastic Net

The original Ifrim et al. method includes a penalty term in the loss function but also regularizes by stopping before full convergence. Different choices of C do affect the resulting model, but early stopping is an easy way to obtain a list of specified length quite quickly (although if the list is nonsensical, then this is obviously not a good move). Ifrim's initial paper in fact has no penalty term in the loss function at all—their entire regularization is due to early stopping.

The relationship of these forms of regularization is unclear; early stopping clearly has great computational advantages and there is no need for convergence checks; we simply stop when we have found enough features of interest. However, it is unclear, for example, whether this approach alone will successfully prune out rare phrases.

6. COMPUTATIONAL COMPARISONS

We compare our CCS tool to three other methods in two studies. The first compares running time and general characteristics of the final word lists, generally using default values and recommended approaches for setting the tuning parameters. The second compares prediction accuracy for the four methods. In a third study, we also examine the CCS tool under different choices of q.

For our data, we use our Fat/Cats corpus and our ECAB corpus, both of which we describe more fully in the case studies section below.

6.1. The Four Methods

The main comparison method is MNIR [5], a text regression method that is primarily designed to be distributed across many cores in order to be able to handle massive data. It parallelizes the regression by conducting individual inverse regressions of the counts of phrases onto the feature space (which is in our case the binary labeling). It is regularized, giving a sparse feature vector. The recommendation of MNIR, regarding tuning, is to use AIC or BIC penalization. We select BIC because we are more interested

in interpreting the covariates than in the quality of prediction, and BIC is known to be superior for model recovery as compared to AIC (see, e.g., [40] or, more generally, [16]). We used the **textir** [5] package. MNIR resulted in very long lists (of often more than a thousand words), so we truncated the list by taking the words with the top 1000 weights. For practical use, we would advocate greater regularization via, e.g. cross-validation, to restrict the list further.

We also compare to a classic linear Lasso approach (see, e.g., [6]). Earlier work [1] shows that the gains from logistic over linear are minimal, and the computational expense is large. We use the **glmnet** package [41], selecting the regularization parameter with cross-validation on the rootmean-squared error (RMSE) test error. The standard package automatically standardizes the columns, so this method uses L^2 -rescaling. Generally, the Lasso lists were short, but if they were above 1000 words we truncated as above.

We finally compare to the original *n*-gram method of Ifrim et al. Other than our own, there does not seem to be an R package implementing this approach, so we replicated their method by using our binary features option and not rescaling the columns. For a more direct comparison, we select the tuning parameter with our permutation approach. They initially advocated early stopping to regularize, but in later work (and in their C++ package which we extended and modified for our package) they introduced direct regularization.

In the first two studies, we used L^2 normalization for our method. We select *C* two ways: the permutation approach discussed above and to a fixed C = 8 to prune perfect predictors for three or fewer documents. We did not allow gaps in phrases, and did not upper-bound phrase length. For the permutation, we permuted ten times and took the maximum as the *C*, as the *C* from the permutations do not vary much.

We ran our trials primarily on the cleaned Fat/Cat dataset stemmed with Porter stemming [42] via the **tm** [43] package. See Section 7.1 for further details. For the latter two methods, the data was stored in a flat file of cleaned and stemmed text, with each line corresponding to a single document. For the first two methods we generated a document-term matrix from this text, dropping all terms that appeared fewer than five times to keep the matrix manageable. This resulted in 8698 unigrams, 81 863 bigrams, and 122 528 trigrams. There were 49 558 documents in total. We ran the Lasso twice, first with unigrams and then with all unigrams through trigrams. For MNIR we only used unigrams; the number of features generated when we expanded to trigrams was computationally prohibitive on a single computer.

To obtain labeling, we selected a random sample of 100 of the 1400 keywords associated with the Fat/Cat reports,

weighted by the frequency of the keywords's appearance, to form the labeling schemes to evaluate. For each keyword, we dropped any phrases from consideration as a feature (e.g., phrases with 'carbon' or 'monoxide' for the carbon monoxide keyword were dropped) by either passing these words as banned words to our algorithm or dropping the relevant columns from the document-term matrix.

To further understand computational timing, we also replicated our first comparison study on the ECAB legal decisions discussed below. Here we had seven judges that were part of overlapping subsets of decisions, and we compared each judge to the baseline (this is not our labeling of primary interest in the case study below).

For the first computational comparison, we ran the four methods on the full data and compared the runtimes and characteristics of the final word lists. For the second computation comparison, we set aside a random 15% of the data, and then predicted the labeling of the set-aside data using each method. The Lasso, the Ifrim et al., and our method all produce nominal predictive scores, often overly low due to the massive imbalance of the positive and negative classes. To find a cutoff point for classification, we fit a logistic regression model on whatever predictive value came out of the method on the training set, and then classified by determining whether the predicted log-odds were above or below 0. For the MNIR method, we first projected the results of the inverse regression, as described in [5], and used the resulting score as the covariate in the logistic regression modeling step. We also calculated AUC (area under the curve) scores using the raw predictive scores for all methods.

Our third investigation was on the impact of different choices of q for L^q -rescaling when using our method. We again generated word lists for each of our 100 keywords, and calculated average length and average frequency of words for the resulting lists for q = 1.2, 2, and 4. We put no upper bound on phrase length.

Simulations were run on a MacBook pro with a 2.9 GHz Intel Core i5 processor, 16 GB memory, and a solid-state drive. Reproducible R code for all simulations is available on request.

6.2. Comparison Results

For each labeling and each method, we calculated different statistics on the word lists. We also timed the total time to generate a list. Table 2 shows the average of these measures across the 100 labelings. A few general trends are evident.

6.2.1. List characteristics

Overall, the resulting lists are very different in character. We calculated the average list length, and then, for lists

Table 2. Results of comparison study on OSHA reports (top) and ECAB legal decisions (bottom). Runtime does not include time to generate the document-term matrix for the Lasso and MNIR.

	Time	List length	Wo	rd freq
Method	(avg. seconds)	(avg.)	(avg.)	(median)
Ifrim	21.5	25	15 285	10178
Lasso	87.6	194	5441	217
Lasso (trigrams)	86.6	440	657	13
Lasso (pentagrams)	101.1	502	503	12
MNIR	42.1	8696	3818	392
textreg	71.2	25	1441	371
textreg (fixed C)	46.4	24	1711	472
Ifrim	63.8	8	17861	16872
Lasso	38.5	261	3643	109
Lasso (trigrams)	138.8	800	898	21
MNIR	163.9	25 990	31	16
textreg	540.3	14	12 595	10639
textreg (fixed C)	555.6	64	2481	653

truncated at the top 1000 terms, calculated the mean and median frequency of the words to gauge how common or rare the selected phrases were. We finally averaged these means and medians across the different labeling runs. We see that the methods are quite different in terms of these scores.

MNIR gives very long lists of very specific words; this is probably due to the inverse regression, which selects all phrases that are relevant without much regard to correlation structure between them. The Lasso also tends to have longer lists, but the average appearance of the selected words is on par with the *n*-gram feature-rescaled regressions. However, when the Lasso had access to trigrams, which can be highly targeted and specific, the median frequency plummeted to 13. For pentagrams, it went to 12. As expected, the Ifrim et al. method, due to no column rescaling, selects *very* general terms, with a median frequency of around 10 000.

For the OSHA dataset, the permutation C values tended to be close to the fixed C, giving overall word lists that were similar as well. For the ECAB dataset, with longer documents, the permutation-selected C was much higher, and thus the word lists were much shorter due to the greater regularization. The resulting words were also typically more general.

Overall, it is clear that the practitioner can use different methods to get different types of lists. We believe, generally, that one wants short lists of phrases, and that those phrases should not be overly general.

6.2.2. Runtime

Runtimes were generally comparable for our OSHA data but widely different for the ECAB data. We discuss the OSHA data first. The Lasso method, even including its cross-validation step, was quite comparable to the textreg method with respect to time due to its very fast implementation. It is also robust to very wide feature matrices; note the average time for using all trigrams is the same as for just unigrams. The Ifrim method is faster than the textreg methods, likely due to improved pruning. MNIR is also quite fast, even though it is designed for parallel systems and we ran on a single core. Generally, the computational times to generate summaries are comparable. However, MNIR on the trigrams was not workable. As MNIR is linear in the number of features, the blowup of features was too much of a time increase. Of course, multiple machines and a parallel structure could avoid this.

We were quite surprised by the time statistics being insensitive to number of terms for the **glmnet** package. To investigate further, we calculated a document-term matrix for all phrases up to five words, giving 363 132 unique terms, giving the extra row in Table 2. Here, the average runtimes for the Lasso increased modestly by about 16% to a mean of 101 seconds.

The computational times were more spread out for the ECAB data; see the bottom half of Table 2. We now see a substantially increased time from moving to unigrams to trigrams for the Lasso, and the timing of the textreg methods exploded. On investigating this time differential further, we found that many selected phrases had five or more words. Furthermore, most runs reached the maximum number of iterations before convergence, indicating flat surfaces. This is a weakness of greedy coordinate descent, which we discuss further below. Overall, we potentially have, due to the use of boilerplate language in legal writing, long yet informative phrases that we wish to see in our summaries. This could make pre-generation of candidate phrases prohibitively expensive. The MNIR method selected all unigrams; each unigram apparently has enough difference in use across judges to be selected under the BIC penalization.

The times for the Lasso and MNIR do not include the time to generate the document-term matrix, which was 7.4 minutes for 5 g and 3.7 minutes for trigrams.

6.3. Predictive Accuracy Results

To assess predictive accuracy, we used the macroaveraged F1-statistic [44] (F1 calculated for each of the 100 trials and then averaged). See Table 3. All methods would sometimes score none of the test set documents as positive, giving undefined F1 scores. These are also indicated in the table. Figure 2 shows these scores along with the component precision and recall, as well as AUC scores, for those keywords where all methods had defined F1. The predictive accuracy is comparable across the methods, although our text regularization does suffer some due to overregularization from the permutation method.

Table 3. F1 scores for different methods averaged across keywords examined (with standard deviations). Number of keywords with no defined F1 (out of 100) are also indicated.

Method	F1	SD	No. missing
MNIR	0.22	0.22	19
Lasso	0.28	0.24	6
Lasso (trigrams)	0.36	0.25	6
Ifrim	0.25	0.24	16
textreg	0.23	0.23	7

Overall, the predictive accuracy is low; recall, in particular, tends to run around 20% for all methods, although the Lasso with trigrams was noticeably superior.

We also examined AUC scores. Here, we again see the cost of textreg's overregularization: conservative classification reduces sensitivity greatly, lowering the ROC (receiver operating characteristic) and AUC scores. The AUC scores are deceptively high due to the imbalance between positive and negative examples.

Generally, the superior performance of the Lasso on trigrams indicates that rescaling features coupled with the richer feature set of multiple words is useful for prediction tasks. Our methods results indicate that overregularization is detrimental to prediction due to, primarily, diminished ability to predict positive documents as positive (shown by the recall scores).

6.4. Selection of *q* Results

As anticipated, different values of q produce lists of different quality. Results are given in Table 4. Higher q

corresponds to lists with relatively more common words and phrases. Low q produces lists that tend to have phrases with more words. For the q = 1.2 list, more than half the phrases were three words or longer. We also found that the length of the list increased with q. For q = 4, we had longer lists with more common words. For low q, the algorithm selected very targeted phrases, and not too many of them. The regularization parameter C is not comparable across q. Therefore we recalculated it via the permutation approach for each run and value of q. Mean values for C are shown in the table for reference.

Generally, we find a rule of thumb worth remembering: longer lists tend to have more general terms. This can happen by adjusting either C or q (see the first simulation and compare different C, for example).

6.5. Discussion

Overall, our method succeeds in accommodating multiple word phrases while also allowing for an intercept and the rescaling of features. As shown, these extensions are critical for generating manageable lists with phrases that are not overly common. Further, picking C and q does control both list length and commonality of words on the lists as predicted by our initial discussion.

In terms of computational time, our method did not perform particularly well when compared to the Lasso of the **glmnet** package, especially considering that the Lasso's overall time includes that of cross-validation. As our method is similar to Lasso regression on a

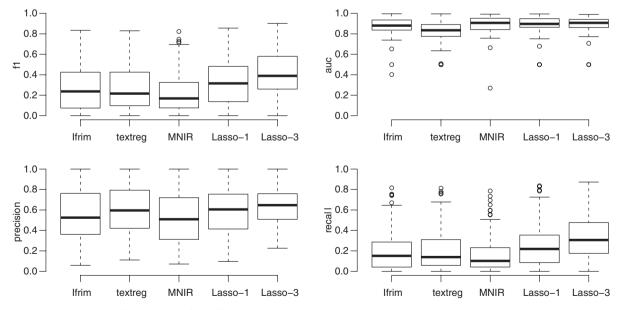


Fig. 2 Out-of-sample classification rates for different methods. Lasso on trigrams (Lasso-3) is generally the best, although there is substantial variability. Keywords in which any of the methods failed to return F1 scores were dropped. Charts are substantively the same with these keywords included.

Method	Time (avg. seconds)	List length (avg.)	Phrase length (avg.)	Word freq (avg.)	Word freq (median)	C (avg.)
textreg (q=1.2)	78.7	8	2.9	310	110	4.6
textreg $(q=2)$	72.8	24	2.2	1492	435	7.1
textreg $(q = 4)$	135.3	35	1.8	8035	1761	19.0

Table 4. Word list characteristics of different rescaling norms. Length is length of list, mean size is the average number of words in the selected phrases, and frequency refers to phrase occurrence in the corpus. Mean C is the average regularization value.

preconstructed phrase matrix (differing only in that it uses a hinge loss instead of a squared loss), one might naturally ask whether an alternative approach would be to simply generate the full document-term matrix and use **glmnet**. We are not entirely convinced, as we discuss next.

First, generating full document-term matrices including phrases with wildcards could be computationally tedious, especially if we imagine extensions such as optionally included words. At the very least, it is expensive in both time and memory, and grows increasingly so with the number of possible phrases; the 7 minutes to generate all 5-grams is well over the model-fitting times. The stored matrix was twice the size of the initial raw text. Second, the hinge loss is different from the classic L^2 loss, and there could be core differences in overall behavior here; this is an area for future investigation. Third, the speed of the glmnet package is partially due to the LARS (least angle regression) approach where the entire regularization path is computed at once. It is arguably also due to the package being particularly well developed and optimized for efficiency. There may be similar ways to substantially speed up our package and method to make it more competitive.

For example, one potential slowdown, determined from examining the convergence paths of many runs, is that we often see an initial selection of a small list of phrases and then a slow hill climb where with each step a different already-selected phrase is selected for adjustment. This comes from the coordinate descent; the optimal gradient path is at an angle, and so following it requires small steps in the associated coordinate directions to trace that path. Unfortunately, with each such step the algorithm conducts a full search for the highest gradient across all phrases, which is expensive. Instead we might after each phrase is selected, find the maximum point given the set of all phrases selected at that point. Then, only if a new phrase were introduced would the algorithm have further steps. This is effectively the LARS [6] approach, and is an important area for enhancement.

Furthermore, the greedy ascent iteration often involves the intercept ratcheting down as features ratchet up. This suggests another possible direction of forcing the intercept to be -1 (for corpora with rare positive labeling rates) rather than estimating it; a -1 intercept corresponds to predicting all documents as negative by default, distinct from no intercept which predicts 50–50 uncertainty as default or allowing the intercept to move which tends to predict the overall base rate as default. Specifying the intercept value would save an intercept update with each step; the optimization problem is then finding a collection of phrases to give positive documents positive weight without affecting the negative documents too heavily.

Regardless, for truly massive data, especially when the number of documents grows large, neither our approach nor the Lasso will work. Instead, methods that allow for easy parallelization, such as MNIR, will be the key. Another area for future exploration, therefore, is to determine how to over-regularize MNIR to get shorter lists, which also might induce lists with more common phrases.

7. CASE STUDIES

We illustrate the CCS tool with two case studies. For the first case study, we also compare the resulting summaries from CCS with the three other methods discussed above. An overview of the code to generate these results using our textreg R package (available on CRAN) is in the Appendix. Full scripts and data are available on our website.¹ For other studies using similar tools, see, for example, [1] or [33]. Before presenting our results, we discuss data representation.

7.1. Data Representation and Cleaning

There are many choices one might make in preparing a corpus for statistical analysis. It is common to, for example, convert text to lowercase and to drop all punctuation. We take that approach here, although we convert all digits to 'X' (to preserve the presence of numbers, in case that is informative) and convert hyphens to blank spaces (so the sequence of hyphenated words would coincide with a nonhyphenated similar phrase—something not possible with single word analyses).

Most text analysis packages would then convert the raw text into an $m \times p$ matrix of counts, dropping

¹Website address is http://scholar.harvard.edu/lmiratrix/software/textreg-r-package

any stop words, but because of the greedy coordinate descent algorithm, unknown phrase length, and the related generation of features on the fly, we store the text as raw strings, with one string per document.

There is some controversy as to whether to 'stem' documents, which is where the tails of many words are cropped so as to collapse the number of possible words. For example 'clean,' 'cleaning', and 'cleaner' might all be cropped to 'clean+' This has the disadvantage of making the resulting text output somewhat difficult to read, especially when considering phrases. Stemming can also lose textual meaning if the different suffixes are in fact important in the context. It has the advantage, particularly for phrases, of collapsing several different versions of phrases into one. We provide stemming as an option but, to enhance readability of output, append a '+' to the end of all stemmed words (and their roots) to indicate they have been potentially cropped. We also provide tools to wildcard search for stemmed phrases in the original text so as to recover examples of the complete phrases.

Sometimes a given context would involve words that are known a priori, or nearly a priori, to have no meaning. We therefore provide an option for custom-made, short stop-word lists (i.e., a list of banned words) that are prohibited from being in any summary phrase. Generally, these lists are built in an iterative process. The first summary generated will often contain words that are immediately recognizable as inconsequential to a researcher with pre-existing contextual knowledge, even though they are correlated with the labeling. The researcher would then drop these words, rerun the algorithm, and repeat as necessary. We do not see any way to avoid this; the case studies below illustrate why.

7.2. Fat/Cats

Our main investigation relies on OSHA's publicly available summaries of occupational Fat/Cats (Fat/Cats),² in the United States, from 2000 to 2010. These summaries describe workplace incidents that have resulted in death or the inpatient hospitalization of three or more workers. When such an event occurs, an employer must report it to OSHA.³⁴ In the course of conducting the resulting investigation, OSHA generates a narrative report, part of which becomes publicly available and is annotated with

² Website address is http://ogesdw.dol.gov/views/data_catalogs. php.

any of a set of about 1400 keywords to categorize the narrative reports in terms of specific chemicals involved, machinery involved, body parts affected, and other salient features. The publicly available records primarily consist of a title and a short paragraph summary of the incident, along with the date, whether the incident involved a fatality, and several other covariates.

We concatenated the title and paragraph description to form the documents. These documents tend to be 56-136 words long (these are the first and third quantiles), with a minimum length of four words and a longest report of 791 words. After stemming, there were 49 840 unique unigrams (word stems), of which 12 700 appeared 10 or more times and 4704 appeared 100 or more times.

To investigate this corpus, we can, for any keyword, generate a labeling of the narrative reports by setting those reports tagged with the keyword as +1 and the remainder as -1. Using CCS on this would then summarize the collection of reports marked with a keyword by comparing them to all other reports. Ideally, this would take out words common to these reports (e.g., 'employee' or other general work-place terms), leaving us with phrases that make the identified set stand out. We would interpret this summary as a distillation of what is distinct about this category of Fat/Cats as compared to Fat/Cats in general. By periodically summarizing reports for each keyword of interest, researchers may gain information about emerging hazards and trends in circumstances. Hopefully, the resulting summaries would be faster to read than the individual narratives, but still contain hints as to general themes within these narratives.

As chemical exposure is an area of particular interest for enhanced surveillance and understanding, we generated a background comparison set of documents by identifying keywords that we deemed to be at least loosely associated with chemical exposure. We then defined the 'chemical family' of narratives as all narratives that were labeled with at least one of these keywords. This allowed us to compare various categories of narratives within the limited context of this chemical family, as well as within the larger context of all other types of narratives. Changing the background set highlights different aspects of what sets apart a marked collection of reports.

As an overview of the overall number of narratives of different topics of interest, Table 5 shows the appearance pattern of the categories examined. We discuss Methylene chloride and Carbon monoxide here, and defer Chemical reaction to a supplementary document. The table also shows how many narratives involved a fatality.

7.2.1. Methylene chloride

As it is our motivating example, we first examined Methylene chloride. We initially selected a value of C = 4

³ 29 C.F.R. 1904.39(a)

⁴ OSHA recently expanded the list of reportable events to include the loss of an eye, amputation, or inpatient hospitalization. Occupational Injury and Illness Reporting Requirements-NAICS Update and Reporting Revisions, 79 Fed. Reg. 56129-56188 (September 18, 2014)(amending 29 C.F.R. 1904.39).

 Table 5.
 Number of narratives with different keyword labelings.

 Second pair of columns restrict database to only reports related to chemical exposure.

	Ge	eneral	Ch	emical
	#	%	#	%
Methylene chloride	17	0.03%	17	0.2%
Carbon monoxide	243	0.5	243	3.5
Chemical reaction	115	0.2	115	1.6
Fatality	20 691	41.8	2575	36.7
Total	49 558	100.0	7014	100.0

Table 6. Methylene chloride (L^2 -rescaling, against all reports). # Phrase is total occurrence of phrase; # Reports is number of reports containing phrase; # Tag is number of methylene chloride reports containing phrase; % Tag is the percent of phrase appearances in methylene chloride reports; and % Phrase is percent of methylene chloride reports containing phrase.

Phrase	# Phrase	# Reports	# Tag	% Tag	% Phrase
A bathtub	9	9	5	56	29
Paint stripper	5	5	3	60	18
Stripper contained	3	3	3	100	18
And reglazing	2	2	2	100	12
From a bathtub	2	2	2	100	12
Remover contained	2	2	2	100	12
Stripping agent	2	2	2	100	12
Tub head	2	2	2	100	12

to ensure that we prune all singleton perfect predictors (see Section 5.2). There are 17 reports marked with the 'Methylene Chloride' keyword. Running CCS on these reports returns two words, 'methylene' and 'chloride'. As these words are not of interest, we immediately added these words to the ban list and reran. Table 6 displays the resulting summary comparing these 17 narratives to all the other narratives.

The summarizer picks up on the coherent theme across these reports of bathroom refinishing. This example is quite encouraging, given our prior knowledge of the dangers of methylene chloride, but the utility of CCS in detecting yet unknown patterns remains to be seen. If we select *C* based on the 95th percentile of 100 permutations, we obtain C = 5.65. The needed *C* to result in a null summary is, by comparison, $C^{obs} = 6.92$. We conclude that there is a statistically significant relationship between the text and the keyword (beyond the presence of the banned words), and that the summary is thus informative. The corresponding summary for C = 5.65 is quite succinct, containing only 'a bathtub' and 'stripper contained'. Picking *C* to give statistical significance appears, here, to drop informative phrases.

The summaries do not necessarily capture information in all tagged documents. In this case, for example, six of the methylene chloride reports do not have any of the phrases in Table 6, and so are not represented. A manual review of these reports revealed that four involved 'stripper's for tile, floors, and furniture. One involved an explosion and one, quite terse, only referred to methylene chloride gas.

7.2.2. Carbon monoxide

We also examined reports relating to carbon monoxide, an asphyxiant odorless gas. We ran CCS with different values of q for L^q -rescaling to examine the impact of different levels of rescaling. We compared the CO cases to all other cases involving any of a set of keywords predetermined to be related to some sort of chemical exposure (i.e., those narratives marked as members of the 'chemical family'). To reduce computational time, we limited attention to phrases that appear at least five times in the corpus. Results are in Table 7.

To obtain these results, we summarized in an iterative process; words such as 'carbon', 'monoxide', 'gas', 'poisoning', 'exposed', 'exposure', 'overexposed', 'carboxyhemoglobin', 'ppm', 'levels', 'partspermillion', 'overcome', and 'co' were eventually dropped. We also removed the more specialized 'hyperbaric', having to do with a medical intervention for CO poisoning and 'cohb', an abbreviation for carboxyhemoglobin, a molecular complex that hemoglobin and carbon monoxide form in the body. All of these words appeared in initial summaries and are due to the technical and/or obvious aspects of CO poisoning; they do not reveal trends or characteristics of interest and thus obscure the desired results. None of these words would have appeared on any conventional stop-word list. As they are in fact correlated with the category, we see no way of automatically removing them.

The final results reflect several known patterns in CO poisoning. For example, the exhaust from gasoline- and propane-powered engines are major culprits of these exposures, particularly in combination with poorly ventilated enclosed spaces. The appearance of the phrase 'cold room' appears to reflect incidents in which propane-fueled fork-lifts and floor cleaning devices were the source of carbon monoxide exposure within cold storage areas, where ventilation can be poor.

In investigating hospitalization, we found 17% of the CO poisoning cases contained 'were hospitalized' versus only 5% of the other chemical-related narratives and 1% of the non-chemical narratives. The fire department was mentioned in nearly 9% of these narratives versus a baseline of 0.5%. This all may be due to lower rates of fatality, with only 36% of the CO poisoning cases involving fatalities as compared to 37% for other chemical family reports and 43% for non-chemical-family reports. Interestingly, 'dead'

Phrase	q = 1.2	q = 1.5	q = 2	q = 3	q = 4	# Reports	# Tagged	% Tagged	% Phrase
A propane powered	0.25	0.02		0.05	0.07	17	15	88	6
Their blood	0.16	0.13				5	5	100	2
Gasoline powered	0.15	0.29	0.28	0.25	0.21	41	30	73	12
Concrete saw	0.11	0.24	0.18	0.17	0.08	10	9	90	4
An X hour	0.10					6	6	100	2
The fire department measured	0.09					4	4	100	2
In a cold	0.09					4	4	100	2 2 2 2 2 2
Operating a propane	0.09	0.13				6	6	100	2
Propane operated	0.07					4	4	100	2
Were using a gasoline	0.07					5	5	100	2
Propane powered		0.25	0.23	0.22	0.13	34	28	82	12
Powered		0.20	0.35	0.41	0.44	169	85	50	35
Forklifts		0.13	0.21	0.26	0.22	23	16	70	7
For fresh		0.13				4	4	100	2
The generator was		0.06				9	7	78	3
X hour		0.02	0.03			16	11	69	5
Overexposure			0.15	0.26	0.28	37	18	49	7
Exhaust			0.04	0.10	0.10	112	35	31	14
Generator was			0.04	0.08		14	10	71	4
Blood			0.03	0.06	0.04	98	30	31	12
A gasoline				0.08	0.10	48	26	54	11
Average				0.07	0.05	16	10	62	4
Found				0.06	0.06	554	68	12	28
Were treated				0.05	0.05	111	22	20	9
The fire department				0.05	0.01	90	21	23	9
Hour				0.03	0.03	85	23	27	9
Source of the				0.01	0.02	32	12	38	5
Ventilation				0.01	0.02	130	32	25	13
Enclosed				0.01	0.01	66	18	27	7
Were taken				0.01	0.02	97	23	24	9
Employees				0.01	0.01	1400	126	9	52
Employees were				0.01	0.01	590	66	11	27
Were				0.00	0.01	2777	168	6	69
Fire department					0.02	211	36	17	15
A propane					0.02	96	26	27	11
Were hospitalized					0.01	384	42	11	17
Dead					0.01	372	38	10	16
Warehouse					0.01	92	19	21	8

Table 7. Different summaries of carbon monoxide (compared to chemical family narratives).

appears in 16% of the CO narratives as compared to 5% in the remainder of the chemical family.

Different rescalings give different styles of summaries. The smaller q = 1.2 and q = 1.5 have very specific phrases (e.g., 'were using a gasoline' and 'their blood') that appear only in the positively marked documents. Larger q give more phrases overall, and give phrases that appear at higher rates in both the positive and negative class. For example, 'employees', with more than 10 000 appearances, appears for q = 4.

Overall, infrequent and specific phrases are relatively easy to interpret, and the more common phrases less so. But their patterns of appearance are striking. 'Employees', for example, appears in 52% of the CO narratives versus a baseline of a mere 11%! Less surprisingly, 'enclosed' appears 7% of the time versus less than 0.5% at baseline. Table 7 contrasts 'Carbon Monoxide' to all incidents labeled with other chemical-related keywords. We also compared CO cases to the full set of cases in the database. That is, we summarized the same collection of reports but used a different baseline point of comparison. Results are in Table 11, in the Appendix. They are broadly similar.

We also analyzed the data using stemming. See Table 12 in the Appendix. Results are again broadly similar (but possibly harder to read). Stemming collapses phrases, which can be helpful, but hampers human readability.

Finally, we compare CCS to the other methods of MNIR, the Lasso, and the Ifrim et al. approach. The methods returned summaries of very different lengths: MNIR was 293, Lasso 65, Ifrim 20, and textreg 12. We had to truncate MNIR and Lasso to display the lists, but we did this by taking the union of the top 20 words of each list to

see maximum overlap. See Table 8, with words sorted by frequency of appearance in the corpus.

We see that the lists are quite different with mild overlap. MNIR generally targets rare phrases, most of which are not displayed, and Ifrim very general ones. MNIR, restricted to unigrams in this instance due to computational concerns, has less overlap than it might otherwise. Textreg has mid-range phrases with a few very rare phrases that are perfect predictors. These phrases were not included in the document-term matrix due to their rarity, so could not be on the Lasso or MNIR lists. Many of the phrases have similar meanings. For example, Ifrim has the general versions (e.g., 'gasolin+') for the more specific (gasolin+ power+) of the Lasso and textreg.

7.3. Legal Decisions

In the context of legal decisions, our motivating question is whether we can efficiently learn about the characteristics of certain types of cases by extracting associated phrases and topics from a corpus of those cases. As an exploratory case study, we chose to examine publicly available decisions from the ECAB, which considers appeals to determinations by the OWCP in the U.S. DOL. OWCP handles compensation claims from federal workers injured during the course of employment. EBAC handles as many as 2000 appeals per year.

Within this case law, one particular area of interest is how ECAB handles compensation claims for so-called 'emotional conditions'. These cases can be challenging for a number of interesting reasons. For example, establishing whether an employee's condition was legitimately caused by workplace conditions requires an analysis of causation that is unique in many ways from that which is appropriate in the context of a physical condition.

To further probe the potential utility of CSS in extracting useful information from large bodies of technical text, we performed an exploratory analysis on the collection of ECAB decisions that relate to both mental health conditions and causality. We sought to determine whether automated summaries would reveal meaningful patterns. These decisions are publicly available through DOL⁵. We examined the years 2005–2010 by scraping them from the website.

We ended up with 11 214 legal decisions, documents generally ranging in length from 1602 to 2691 words (these being the first and third quartiles) and a median length of 2074 words. The shortest was 281 words and the longest 12 664. There are 107 302 unique words, of which 37 474 appear 10 or more times and 11 264 appear 100 or

⁵ Website address is http://www.dol.gov/ecab/decisions/main. html

Table 8.	Different	summaries	of	carbon	monoxide	(comparing
different r	nethods).					

PhraseMNIRLassoIfrimtextregFreq.Power+ generat+ to+0.43Level+ of approxim+0.44Food+ poison+0.15For fresh+0.35Propan+ power+ floor+0.25	
Level+ of approxim+ 0.4 4 Food+ poison+ 0.1 5 For fresh+ 0.3 5 Propan+ power+ floor+ 0.2 5	
Food+ poison+0.15For fresh+0.35Propan+ power+ floor+0.25	
For fresh+0.35Propan+ power+ floor+0.25	
Propan+ power+ floor+ 0.2 5	
Propan+ power+ floor+ 0.2 5	
Tailpip+ 6.6 5	
Their blood 0.2 5	
Unventil+ 5.6 5	
Vanguard 5.6 5	
Over + expos + to + 0.2 6	
Decatur 6.2 7	
Receiv+ oxygen+ 0.1 7	
Mek 5.2 8	
Twa 7.2 8	
Carbonyl 7.3 9	
Exhaust+ fume+ 0.1 9	
Intern+ combust+ engin+ 0.2 9	
Newton 5.5 9	
Power+ concret+ 0.1 11	
Qa 5.4 11	
Transient 5.4 11	
Stratton 5.6 12	
Three employe+ are expos+ 0.3 12	
Weight+ averag+ 0.1 12	
$\begin{array}{ccc} \text{weight} + \text{averag} + & 0.1 & 12 \\ \text{Brigg} + & 5.4 & 13 \end{array}$	
Emiss+ 5.4 13	
Fd 6.2 14	
I Stitle 1 Stitle 1	
1	
1	
· · · · ·	
Overcom+ by 0.1 0.1 207	
Poison+ 5.8 0.2 0.9 0.5 210	
Overcom+ 4.3 0.1 249	
Ventil+ 3.5 0.0 321	
Gasolin+ 3.2 0.1 412	
Expos+ to+ $0.0 0.0 490$	
Exposur+ 4.1 0.1 0.2 0.1 519	
Exhaust+ 3.1 0.0 565	
Propan+ 2.9 0.1 620	
Expos+ 2.8 0.0 1053	
Level+ 2.1 0.0 3282	
Found+ 1.7 0.0 3502	
Power+ 1.4 0.0 5779	
Forklift+ 0.6 0.0 9118	
Fell+ 0.0 19483	
Were+ 1.2 0.0 27668	
When 0.0 36629	
His 0.0 53632	
Was 0.0 176175	

more times. These counts include case identifiers and other character strings as words. We do not attempt to remove them directly.

We automatically labeled all of the decisions with two sets of dummy variables, one for emotional condition and one for discussion of causality or work-relatedness of the injury. For each, we labeled documents if they

		% Correct	Estimated				
Labeling	Total	(sample)	Positive	Negative	Sense.	Spec.	
'Emotional condition'	1479	94/100	1390	89	0.83	0.99	
No 'emotional condition'	9735	97/100	292	9443			
Total cases	11214						
'Causality/work-relatedness'	4236	99/100	4194	42	0.75	0.99	
No 'causality/work-relatedness'	6978	80/100	1396	5582			
Total Cases	11214						

Table 9. Manual review of labeling quality for legal decisions.

Table 10. Summary of cases involving an 'emotional condition' and a discussion of causality. Different columns correspond to the number of dropped phrases. First column is only 'emotion' and 'condition'. The rest are adding more and more phrases. Column B adds 'ecab', Column C 'depression', and Column D 'xx' (for illustration purposes). Column D includes many other terms.

	Main	+'ecab'	+'depression'	+'xx'	+Many
Illness that is	0.46	0.48	0.48	0.60	0.60
Cutler xx ecab	0.32				
Requirement imposed by the	0.31	0.31	0.31	0.32	0.32
Psychiatrist	0.05	0.05	0.05	0.06	0.06
Incidents alleged to have caused	0.05	0.05	0.05		
Lillian Cutler	0.04	0.03	0.03		
Compensable	0.03	0.03	0.03	0.03	0.03
Disorder	0.01	0.01	0.01	0.01	0.01
Factor of employment and	0.01			0.01	0.01
A factor of employment	0.01	0.01	0.01	0.02	0.02
Not covered	0.01	0.01	0.01	0.03	0.03
Reaction to	0.01	0.00	0.00		
Anxiety	0.00	0.00	0.01	0.01	0.01
Cutler xx ecab xxx xxxx	0.00				
depression	0.00	0.00			
Lillian Cutler xx	0.00	0.30	0.30		
Compensable factor of employment and		0.01	0.01		
Results from an				0.06	0.06
Environment or				0.01	0.01
An administrative or personnel				0.01	0.01
Requirement imposed				0.01	0.01
Allegations				0.00	0.00

contained any of a set of handpicked key phrases. Once we tuned our collection of key phrases, we took a random sample of the positively and negatively marked documents and conducted a manual review. The labeling is clearly not perfect, as is illustrated in Table 9. Ideally, the CCS method will still be able to produce relevant summaries despite the noise of the missed labels. Although it is possible that specific types of positive decisions are systematically missed due to the ad hoc labeling, discovery of meaningful summary phrases would nevertheless be suggestive.

Column 1 of Table 10 shows a first-pass summary of those cases that both involve an emotional condition and revolve around issues of causality. We see fairly general terms and some boilerplate language. Here, it is necessary to explore the raw text to discover the contexts for these phrases. This is easily done using our package. For example, one positively marked decision has:

Not every injury or **illness that is** related to a claimant's employment is compensable. In the case of **Lillian Cutler**, the Board explained some distinctions as to the type of employment situations giving rise to a compensable emotional condition under FECA.

C.E., Docket No. 10-461 (issued November 23, 2010) (emphasis added).

Another is:

To establish that an emotional condition arose in the performance of duty, a claimant must submit the following: (i) medical evidence establishing that she has an emotional or psychiatric disorder; (ii) **factual evidence identifying** employment factors or

incidents alleged to have caused or contributed to the condition; and (iii) rationalized medical opinion evidence establishing that the emotional condition is **causally related** to the identified compensable employment factors.

T.G., 58 ECAB 189 (2006) (emphasis added).

As an illustration of stability of CCS, consider the other columns of Table 10. Each column corresponds to dropping more and more terms from consideration. Note that the transition from the third to fourth column drops one of the case references, 'Lillian Cutler', even though we did not explicitly drop those words and phrases. CCSselected phrases are picked in the context of other phrases. Because we removed 'xx' (indicating a case number), 'Lillian Cutler' is no longer selected, along with other phrases including parts of this phrase and 'xx'. Instead, we obtain a cluster of phrases showcasing different aspects of these cases. Dropping phrases can affect a summary only if those phrases are in the summary. The final two columns are the same because none of the additional phrases was in the summary from column 4. Care must be taken to understand the complex dependencies between phrases.

Thus, in the context of ECAB decisions, the CCS tool provides phrases that flag boilerplate language and case citations. To some degree, these phrases appear to reflect precedent and common statements of law that characterize a given category of cases. While our results are exploratory, inexact, and not particularly revealing in and of themselves, they do suggest that a refined CCS tool might one day facilitate the development of automated case content analysis or aid the development of refined legal taxonomies.

7.4. Discussion

As the above studies illustrate, using these tools to understand text is a very different, and far less precise, activity than working to correctly classify text. The common problems with machine-learning approaches (selecting methods, selecting tuning parameters, etc.) are only exacerbated by this uncertain area. The researcher is left with many decisions to make and only vague guidance on how to make them.

With our method, two such decisions are prominent: what method to use for selection of the regularization parameter C, and what method to use for selection of the feature-normalization parameter q. We also need to determine how to remove domain-specific stop words.

7.4.1. Picking C

For selecting C, we have several options, especially if we include picking a regularization parameter by optimizing predictive performance. Which option to select is a difficult, especially since there is no easy metric of final quality if one's focus is not prediction. The computational investigations shed some light on this problem. Ideally, one would use the maximum of the permutation approach and the rare-phrase pruning approach. This will guarantee finding a summary only if one should be found, and also will discard rare phrases that do not speak of general trends across the positive documents. The free test from the permutation-selected C of whether the phrases as a whole are in fact significantly associated with the text is a real boon, in our view. It moves toward presentation of results that are known to not be entirely noise.

That said, future work on stability (where documents are perturbed to see how the selected phrases change, for example) is a must. Furthermore, we acknowledge that the above examples suggest that the permutation-selected C is severe, more severe than from cross-validation or similar methods. This means we can lose human meaning as illustrated by the methylene chloride example: the richness of the summary was much greater with a slightly reduced C. Relaxing regularization toward what would be achieved with prediction-oriented approaches to achieve longer lists may be informative, but (other than improved prediction) this could undermine the guarantees provided by the above. Perhaps, work on testing individual phrases via false discovery rates could find a better balance.

Regardless, one should always compare the finally chosen C to Table 1 to see to what degree it is discarding perfect predictors, and to what degree it would leave the remainder to be potentially picked up. This provides a human interpretation of the impact of the regularization.

7.4.2. Picking q

Selecting q is also admittedly difficult. By design, it gives different views of the data, from the quite general to the very specific. We advocate for exploring a range of values as the best practice. In the above case studies, for example, the full range of terms on the tables provided the most complex and rich story; perhaps pooling the lists and exploring these pooled lists would be one way forward. We underscore that we view these tools as exploratory; the researcher can extract small snippets of text to see whether they offer some clue toward a more thorough investigation. This is similar in spirit to, for example, XGobi [45].

7.4.3. Stop words

Even though we avoid using stop-word lists as a first-pass approach, we still need to generate specialized stop-word lists. We see no way to have removed these words without human intervention, as, from a prediction standpoint, the removed words are key indicators of a given labeling. Unfortunately, selecting them occludes terms that could enhance human understanding. We can see this with the calculated C thresholds: as we add words to the ban list, the C plummets because we are removing the words that are most correlated with the labeling. Eventually, we could reach a point where we have 'conditioned out' all the connection of the labeling to text. This is another potential avenue for exploring such data. Overall, we advocate generating modest-length stopword (ban) lists using substance-matter knowledge, coupled with rescaling, over using generic stop lists that allow milquetoast words through due to their being not obviously wrong.

8. CONCLUSIONS

We presented a method for comparing sets of documents that is simple, sparse, and fast. We argued that these qualities are important for text analysis, especially if it is to be used for surveillance or other exploratory tasks. Here, 'simple' means the summaries cannot be too technical in nature. For example, the presence/absence of features is easier to interpret than regression weights. We need sparsity as humans are lazy; the number of phrases in a summary must be few. The faster the summaries can be computed, the better. Otherwise, exploratory analysis and discovery are bogged down.

We do not, however, argue that the results of these tools, or in fact any other text analysis tools that we know of, can be taken as ultimate proof of any particular substantive theme or meaning. Summaries can be quite suggestive, but researchers would need to investigate further to substantiate those suggestions. Alternatively, secondary analyses such as blind scoring of the key phrases for sentiment could lead to traditional statistical conclusions. In these cases, CCS should be viewed as a dimension reduction tool, providing a targeted, small number of informative features for a very complicated form of data.

On the technical side, we have effectively provided an implementation of Lasso-style regression where the full set of features is dynamically created and the loss is a squared hinge loss rather than a normal quadratic loss. This work shows that implementing sparse regression with greedy coordinate descent offers a viable direction for summarization using phrases rather than words. Furthermore, the approach of dynamically building features shows promise for other customizations such as skipping or dropping words to automatically detect related phrases, collapsing them into single features.

Admittedly, more work needs to be done to optimize this particular implementation to see exactly how fast it could be; currently the algorithm is suboptimal because it does not fully fit currently selected features at each iteration. That said, compared to methods with a precomputed design matrix, it is comparably fast, is more flexible in the *n*-grams considered, and allows for some trickiness such as having gaps in the key phrases (i.e., wildcard words) and the enforcement of nonnegative weights. Additionally, the **textreg** package is quite natural to use, allowing users to avoid calculating the phrase matrix and, instead, work with raw text. It also easily allows for rescaling schemes other than L^2 .

Going beyond text analysis, these methods also hint at ways of incorporating many interaction terms among features in high-dimensional regression. Phrase features are simply interactions of nearby word features, and thus similar bounding methods may exist. This is another area for future exploration.

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APPENDIX A

Appendix A consists of supplemental tables showing alternate summaries of the case studies discussed above. Table A1 compares carbon monoxide narratives to all the other narratives. Table A2 demonstrates a series of summaries on a stemmed corpus.

Were using a gasoline 0.28 0.31 5 Propane-powered floor 0.28 5 Their blood 0.28 0.52 0.45 0.64 0.65 5	Tagged % 5 5 5 5 5 5 5	Tagged % 100 100 100 100 100 100	Phrase 2 2
Propane-powered floor 0.28 5 Their blood 0.28 0.52 0.45 0.64 0.65 5	5 5 5	100 100	2
Their blood 0.28 0.52 0.45 0.64 0.65 5	5 5	100	2
	5		•
H 11: 015 5		100	2
Hemoglobin 0.15 5	5	100	
In the cold room 0.15 0.16 5		100	2 2
They were found 0.34 0.15 4	4	100	2
Powered forklifts 0.31 0.18 9	8	89	3
Gasoline powered 0.30 0.47 0.56 0.53 66	30	45	12
Propane powered 0.28 0.40 0.53 0.53 52	28	54	12
Overexposure 0.18 0.34 0.53 0.58 37	18	49	7
Exhaust fumes 0.13 6	5	83	2
Exhaust 0.03 0.10 0.11 196	35	18	14
Calculated 0.18 12	7	58	3
Generator was 0.13 0.17 24	10	42	4
Employees were treated 0.10 0.13 40	12	30	5
Evacuated 0.07 0.10 104	18	17	7
The fire department0.06241	21	9	9
Ventilation 0.05 0.07 203	32	16	13
Were treated 0.05 0.04 217	22	10	9
A propane 0.05 0.05 128	26	20	11
Powered 0.03 0.03 740	85	11	35
Blood 0.03 0.04 355	30	8	12
Were hospitalized 0.02 0.03 778	42	5	17
Found 0.01 0.02 2885	68	2	28
Employees 0.00 0.01 5575	126	2	52
Stratton 0.27 12	7	58	3
Headaches 0.16 36	10	28	4
Source of the 0.09 60	12	20	5
A gasoline 0.07 70	26	37	11
Passed out 0.06 108	11	10	5
Enclosed 0.05 180	18	10	7
Fire department 0.03 778	36	5	15
Cold 0.03 219	16	7	7
Hours 0.01 1277	33	3	14
Room 0.01 1580	42	3	17

Table A1. Different summaries of carbon monoxide (when comparing against all other cases).

Table A2. Different summaries of carbon monoxide (using stemming, compared to chemical family).

Phrase	q = 1.2	q = 1.5	q = 2	q = 3	q = 4	# Reports	# Tagged	% Tagged	% Phrase
Poison+ at+	0.33	0.12	0.10	0.08	0.03	17	17	100	7
Propan+ power+	0.12	0.38	0.55	0.56	0.49	34	28	82	12
Gasolin+ power+	0.11	0.38	0.54	0.56	0.54	43	32	74	13
For fresh+	0.10	0.19	0.10			4	4	100	2
Poison+	0.09	0.36	0.49	0.53	0.53	141	96	68	40
An+ X hour+	0.04					6	6	100	2
Were+ use+ a+ gasolin+	0.04					5	5	100	2
Was oper+ a+ propan+	0.03					6	6	100	2
The cold room+	0.03	0.02				7	7	100	3
Concret+ saw+		0.08	0.21	0.10		10	9	90	4
XXX cubic+		0.06				4	4	100	2
The generat+ was		0.06	0.09			10	8	80	3
Forklift+ were+		0.01	0.05			8	7	88	3
Averag+			0.13	0.16	0.10	17	11	65	5
Headach+ dizzi+			0.06			6	5	83	2
Intern+ combust+			0.05			8	6	75	2

Phrase	q = 1.2	q = 1.5	q = 2	q = 3	q = 4	# Reports	# Tagged	% Tagged	% Phrase
The cold			0.04			17	10	59	4
Exhaust+			0.03	0.07	0.08	142	38	27	16
Generat+			0.03	0.05	0.03	110	31	28	13
Sourc+ of the			0.02	0.10	0.08	33	13	39	5
X hour+			0.01	0.03		57	18	32	7
Hour+				0.03	0.04	362	47	13	19
Found+				0.02	0.03	561	69	12	28
A+ gasolin+				0.02	0.02	48	26	54	11
Cold				0.01		66	16	24	7
Fire+ depart+				0.01	0.01	217	37	17	15
Ventil+				0.01	0.01	173	42	24	17
Power+				0.01	0.02	459	95	21	39
Were+ treat+				0.01	0.01	112	22	20	9
Were+				0.01	0.01	2778	168	6	69
Employe+ were+				0.01	0.00	612	67	11	28
Room+				0.00	0.00	475	43	9	18
Employe+ are					0.01	305	37	12	15

 Table A2.
 Continued

APPENDIX B: DERIVATIONS

We here show three derivations used in the above work. We first show how to obtain the bound on the gradient. Second, we give an alternate formulation of the loss function, which gives a different approach for finding the feature with the maximal gradient. We then show how to obtain the minimal λ^* to ensure perfect predictors are pruned.

Bound on the Gradient

The gradient for phrase j is

$$\frac{d}{d\beta_j}\mathcal{L}(\beta) = \sum_{i=1}^n \xi'(m_i)y_i \frac{c_{ij}}{z_j} + C\frac{d}{d\beta_j}|\beta_j|$$
$$= \sum_{y_i=1,c_{ij}>0} \xi'(m_i)\frac{c_{ij}}{z_j} + \sum_{y_i=-1,c_{ij}>0}$$
$$-\xi'(m_i)\frac{c_{ij}}{z_j} + C\frac{d}{d\beta_j}|\beta_j|$$

Now consider all phrases k with phrase j as a prefix that are currently not in the model (i.e., which currently have $\beta_k = 0$), and maximize over the possible gradients (a similar argument gives a bound below for negative gradients). For vectors a, b, let $a \leq b$ denote a component-wise relationship of $a_i \leq b_i$ for i = 1, ..., m. Then the set $\{a : 0 \leq a \leq c_j\}$ contains all potential appearance patterns for a phrase with phrase j as a prefix. We do not wish to calculate what the actual phrases are, hence we optimize over this set of potential phrases. This results in the following optimization problem:

$$U = \max_{0 \le a \le c_j} \sum_{y_i = 1, c_{ij} > 0} \xi'(m_i) \frac{a_i}{z_a} + \sum_{y_i = -1, c_{ij} > 0} -\xi'(m_i) \frac{a_i}{z_a} + C \frac{d}{d\beta_j} |\beta_j|$$

$$\leq \max_{0 \le a \le c_j} \sum_{y_i = -1, c_{ij} > 0} -\xi'(m_i) \frac{a_i}{z_a} - C = \max_{0 \le a \le c_j} \frac{1}{z_a} \langle w, a \rangle - C$$

with $z_a = |a|_q$ being the L^q norm of a, w a *m*-vector of weights with $w_i = -\xi'(m_i)$, and $\langle \cdot, \cdot \rangle$ the inner product.

Because $\xi'(m_i)$ is everywhere nonpositive, the first term in the first line above is negative. The second line follows because setting $a_i = 0$ for any document with $y_i = 1$ only increases the gradient, as doing so will simultaneously drop negative terms and shrink z_j . The penalty term is negative because we are examining gradients at 0; if we step ϵ in the negative direction (as indicated by the first term), the gradient will immediately be shrunk toward 0. The m_i , which include the intercept μ , are fixed constants, determined by the current location of our optimization path. We are effectively maximizing over an inner product of a and a vector of weights w with $w_i = \xi'(m_i)$ for i = 1, ..., n. Overall, this bound is assessing the maximum possible utility of a hypothetical super phrase, which boils down to maximizing weights put on positive examples.

The normalization z_a renders this problem difficult. We can bound the optimization using the following relationship:

$$|x|_2|w|_2 \le |x|_q|w|_r$$

for q, r such that 1/q + 1/r = 1. This inequality gives

$$U = \max_{\substack{0 \le a \le c_j \ |a|_q}} \frac{1}{|a|_q} |a|_2 |w|_2 \cos \theta - C$$
$$\leq \max_{\substack{0 \le a \le c_j \ |a|_q}} \frac{1}{|a|_q} |a|_q |w|_r \cos \theta - C$$
$$= \max_{\substack{0 \le a \le c_j \ |w|_r \ \cos \theta - C}} |w|_r \cos \theta - C$$
$$\leq |w|_r - C$$

with θ being the angle between a and w.

Coupling this with a similar argument for minimization gives the overall bound.

A note on the Elastic Net: The elastic net [17] is where we penalize our loss function with (using Ifrim et al.'s notation)

$$CR_{a}(\beta) = Ca \sum_{j=1}^{p} |\beta_{j}| + C(1-a) \left(\sum_{j=1}^{p} |\beta_{j}|^{2}\right)^{1/2}.$$

This regularization tends to keep groups of correlated features rather than picking one; it can borrow from the stability of ridge regression. It can be potentially useful when many small features have weak signals. Setting a = 1 corresponds to L^1 regularization.

For our problem, the gradient search is just changed to a subtraction of Ca rather than C for the at-zero potential new features. The gradients calculated for features already in the model have an extra term of $C(1-a)2\beta_j$ due to the derivative of the second term above.

Alternate Gradient Formulation

By redefining β , we can change the optimization problem to have a L^q -rescaling term in the penalty. This gives different bounds on the gradients for super-phrases based on a sub-phrase. However, it also changes the gradients themselves, which would change the path of the optimization problem. That is, different features will initially have the largest gradient. Assuming true convergence, the final solution will be identical, however.

In particular, define $\tilde{\beta}_j = \beta_j / z_j$. The loss term can then be re-expressed as

$$\mathcal{L}(\beta) = \sum_{i=1}^{n} \xi(m_i) + C \sum_{j=1}^{p} z_j |\tilde{\beta}_j|.$$

The gradient for phrase j is then

$$\frac{d}{d\beta_j}\mathcal{L}(\beta) = \sum_{i=1}^n \xi'(m_i)y_ic_{ij} + Cz_j\frac{d}{d\beta_j}|\beta_j|$$
$$= \sum_{y_i=1,c_{ij}>0} \xi'(m_i)c_{ij} + \sum_{y_i=-1,c_{ij}>0}$$
$$-\xi'(m_i)c_{ij} + Cz_j\frac{d}{d\beta_i}|\beta_j|.$$

Again, consider all phrases with phrase j as a prefix and maximize over the gradient, yielding the following optimization problem:

$$U = \max_{0 \le a \le c_j} \sum_{y_i = 1, c_{ij} > 0} \xi'(m_i) a_i + \sum_{y_i = -1, c_{ij} > 0} -\xi'(m_i) a_i + C z_j \frac{d}{d\beta_j} |\beta_j| = \max_{0 \le a \le c_j} \sum_{y_i = -1, c_{ij} > 0} |\xi'(m_i)| a_i - C \left(\sum_{i=1}^n a_i^p\right)^{1/p}.$$

As before, the second line comes from noticing that setting $a_i = 0$ for any document with $y_i = 1$ will only increases the gradient due to dropping the negative terms and shrinking z_j . The penalty term is still negative because, if we step ϵ in the negative direction from 0 (which is indicated by the first term), the gradient will immediately shrink toward 0.

If we consider only phrases that have at least r occurrences in our corpus, then we can roughly bound with

$$\max_{k:x_k \leq x_j} \frac{d}{d\beta_k} L(\beta) \leq \sum_{y_i = -1, c_{ij} > 0} \left| \xi'(m_i) \right| c_j - Cm^{1/p}.$$

This is from maximizing both terms separately. For the first, we simply add the maximum weight, without regard to the normalizing constant. For the normalizing constant, given a total count of r occurrences, the

maximum z_j would be putting singletons on each of r documents, giving the $r^{1/p}$ total.

Similarly, bounding from below gives an overall bound of

$$\left| \max_{\substack{k:x_k \leq x_j \\ y_i = 1, c_{ij} > 0}} \frac{d}{d\beta_k} L(\beta) \right| \leq 0 \bigvee \max\left\{ \sum_{\substack{y_i = -1, c_{ij} > 0 \\ y_i = 1, c_{ij} > 0}} \left| \xi'(m_i) \right| c_{ij} \right\} - Cr^{1/p}$$

This bound appears to be less useful than the one presented in the main paper. Furthermore, not rescaling by z_j tend to make more common phrases be selected first, as we are not rescaling the first term, allowing it to grow quite large.

Perfect Predictors

Take the count vector for a perfect predictor c_j . It has r 1s and m - r 0s. For the regression, the count vector c_j is q-rescaled, giving

$$x_j = \frac{1}{z_j} c_j = \frac{1}{r^{1/q}}$$
 as $z_j = \left(\sum_{i=1}^m |c_{ij}|^q\right)^{1/q} = r^{1/q}$.

Assume that feature c_j has been set aside and we have optimized without it. We have $\hat{Y} = X\tilde{\beta}$ (with $\tilde{\beta}_j \equiv \beta_j/z_j$ except for the intercept) for our current set of predictions, and an overall predicted mean $\mu = \frac{1}{m} \sum \hat{Y}_i$. Now reintroduce feature c_j . Our loss function, when only considering feature c_j , is then

$$\ell(\beta_j) = \sum_{i=1}^m \left(Y_i - \hat{Y}_i\right)^2 + C|\beta_j|$$
$$= \sum_{i=1}^m \left(Y_i - X_i\tilde{\beta} - x_{ij}\beta_j\right)^2 + C|\beta_j|$$
$$= \sum_{i:x_{ij>0}} \left(1 - X_i\tilde{\beta} - \frac{\beta_j}{r^{1/q}}\right)^2 + C|\beta_j|,$$

where we have dropped those terms not dependent on β_j . This is convex. Take the derivative and set equal to 0 to find the minimum:

$$\begin{split} \ell'(\beta) &= -\frac{2}{r^{1/q}} \sum_{i:x_{ij>0}} \left(1 - X_i \tilde{\beta} - \frac{\beta_j}{r^{1/q}} \right) + Csgn(\beta) \\ &= -\frac{2}{r^{1/q}} \sum_{i:x_{ij>0}} \left(1 - X_i \tilde{\beta} \right) + \frac{2}{r^{1/q}} \sum_{i:x_{ij>0}} \frac{\beta_j}{r^{1/q}} + Csgn(\beta) \\ &= -\frac{2}{r^{1/q}} \sum_{i:x_{ij>0}} \left(1 - X_i \tilde{\beta} \right) + 2\beta_j r^{1-2/q} + Csgn(\beta). \end{split}$$

The β s will not be negative, and hence we examine the positive case, allowing us to drop the *sgn()* term. Set $\ell'(\beta)$ equal to 0 and solve, giving

$$\hat{\beta}_j = r^{1/q} \left(\frac{1}{r} \sum_{i: x_{ij>0}} \left(1 - X_i \tilde{\beta} \right) \right) - \frac{1}{2} r^{2/q-1} C.$$

The term in the outer parenthesis is the average prediction for the documents having the perfect predictor c_j .

If for a document *i* with $c_{ik} = 1$ there are not really any other predictive features, then $\hat{Y}_i = X_i \tilde{\beta} \approx \mu$. If this is true for all of the documents predicted by c_k , then the above is then approximately

$$\hat{\beta}_j \approx r^{1/q-1}r(1-\mu) - \frac{1}{2}r^{2/q-1}C$$

$$= r^{1/q}(1-\mu) - \frac{1}{2}r^{2/q-1}C.$$

If some documents are predicted by other features included in the model, then the sum will be less and the necessary C for pruning will be reduced. Rearrange to obtain an approximate cutoff for C to drop all perfect predictors that perfectly predict r documents.

The
$$q = 2$$
 case.: For $q = 2$, we have

$$\hat{\beta}_j = \sqrt{r} \frac{1}{r} \left(\sum_{i:x_{ij>0}} \left(1 - X_i \tilde{\beta} \right) \right) - \frac{1}{2}C$$
$$\approx \sqrt{r} (1 - \mu) - \frac{1}{2}C.$$

For the positive examples, $x_{ij} = 1/\sqrt{r}$, giving prediction of

$$\hat{Y}_i \approx \mu + \frac{1}{\sqrt{r}} \left(\sqrt{r}(1-\mu) - \frac{1}{2}C \right) = 1 - \frac{1}{2\sqrt{r}}.C$$

The first term of $\hat{\beta}_j$ takes our prediction perfectly to +1, and the second term shrinks the coefficient away from 1 by half of *C*. Predictions from predictors that predict for more documents will be shrunk less than those for fewer. The raw coefficients will also be larger.

APPENDIX C: USING THE TEXTREG PACKAGE

Our text regression package, **textreg** on CRAN, is a wrapper for an extensively modified version of the C++ code written by Georgiana Ifrim. It is also designed to integrate with the **tm** package, a commonly used R package for dealing with text. Our package is fully documented, but it is research code, meaning gaps and errors are possible; the authors would appreciate notification of anything that is out of order.

The primary method in this package is the regression call **textreg()**. This method takes a corpus and a labeling vector and returns a **textreg.result** object that contains the final regression result along with diagnostic information that can be of use. The (somewhat edited) function heading along with default values is

```
textreg <- function(corpus, labeling, banned=NULL,
    C = 1.0, Lq = 2,
    maxIter = 40,
    verbosity = 1,
    positive.only = FALSE, binary.features = FALSE,
    no.regularization = FALSE,
    min.support = 1, min.pattern = 1, max.pattern = 100, gap = 0,
    convergence.threshold=0.0001)
```

The main arguments to this method are listed below:

corpus A vector of strings or a corpus object built out of strings.

labeling A vector of +1/0/-1 values, where 0 means drop from consideration.

- **banned** A vector of unigrams (words) that should not be allowed in any summary phrase.
- **C** The *C* tuning parameter for regularization.
- Lq The q for the L^q -rescaling of terms. A value of 10 or above is treated as infinity.
- **maxIter** The maximum number of iterations allowed before terminating even under no convergence.
- verbosity 0 means silent. Larger numbers mean more diagnostic printout.
- **positive.only** Only allow positive features (other than the intercept). Useful if there are few positive documents and many negative, baseline, documents.
- **binary.features** The feature vectors are 0-1 vectors indicating whether a phrase is in or not in any given document. This is compared to vectors of counts of how many times a phrases in a document. These feature vectors are L^q -rescaled regardless.
- **no.regularization** If TRUE, then features will not be rescaled (which recovers the Ifrim et al. algorithm).
- **min.support** Phrases that do not appear this many times are not considered viable features. Increasing this number can greatly decrease the running time of the algorithm, but it will force the dropping of very rare phrases regardless of rescaling or regularization choice.
- min.pattern,max.pattern Minimum and maximum lengths for phrases that are considered.
- gap Number of words that can appear in a gap. A phrase can have multiple gaps of this length.

The resulting **textreg.result** object can be printed, plotted, and explored. Try, in an R Console, typing **rs** by itself or **plot(rs)**. The method **reformat.textreg.model(rs)** gives a nice table (see, e.g., Table 6) of summary statistics for the final phrases. The sideby-side summary table such as Table 7 is made by passing a list of **textreg.result** objects to **make.list.table()**. The method **calc.loss(rs)** gives the final loss of a result **rs**, and **predict** (**rs)** will return individual document-level predictions of the labeling. The method **rule.to.matrix(rs)** gives back the $n \times r$ design matrix for the final selected r phrases including intercept.

To pick a tuning parameter, one can use

Cs <- find.threshold.C(corpus, labeling, ban.words, R=100)

This method returns a R + 1 length list of numbers. The first number is the choice of C that will return a null model for the labeling given, and the subsequent R numbers constitute our found C values that return a null model under a random permutation of the labeling (holding the zeros fixed). It takes the same parameters as **textreg** except for **maxIter** and **C**. Be sure to use the same remaining values for both calls so that **find.threshold.C** culminates with a C corresponding to the correct model family.

For exploring text, **sample.fragments(phrase, labeling, corpus)** is useful. See also **grab.fragments()**. To profile specific phrases, possibly even phrases not in the results, use **make.phrase. count.table()**.

One can make cluster plots of how phrases relate with **cluster.phrases(rs)**, or make matrices of co-occurrence of phrases using **make.phrase.correlation.chart(rs)**.

All of the above, and a bit more, is demonstrated and more fully explained in the vignette 'Bathtub Demo', that comes with the package. Please read through it for further discussion and ideas.

REFERENCES

- J. Jia, L. Miratrix, B. Yu, B. Gawalt, L. E. Ghaoui, L. Barnesmoore, and S. Clavier, Concise comparative summaries (CCS) of large text corpora with a human experiment, Ann Appl Stat 8 (1) (2014), 499–529.
- [2] D. M. Blei, A. Y. Ng, and M. I. Jordan, Latent Dirichlet allocation, J Mach Learn Res, 3 2003, 993–1022, .
- [3] G. Ifrim, G. Bakir, and G. Weikum, Fast logistic regression for text categorization with variable-length n-grams, In 14th ACM SIGKDD International Conference on Knowledge Discovery and Data Mining, 2008, 354–362.
- [4] G. Ifrim and C. Wiuf, Bounded coordinate-descent for biological sequence classification in high dimensional predictor space, In 17th ACM SIGKDD International Conference on Knowledge Discovery and Data Mining, 2011, 708–716.
- [5] M. Taddy, Multinomial inverse regression for text analysis, J Am Stat Assoc, 108 (503) (2013), 755–770.
- [6] R. Tibshirani, I. Johnstone, B. Efron, and T. Hastie, Least angle regression, Ann Stat, 32 (2) (2004), 407–451.
- [7] J. Grimmer and A. Bayesian, Hierarchical topic model for political texts: measuring expressed agendas in senate press releases, Polit Anal, 18 (1) (2010), 1–35.
- [8] J. Chang, J. Boyd-Graber, S. Gerrish, C. Wang, and D. M. Blei, Reading tea leaves: how humans interpret topic models, In Neural Information Processing Systems (NIPS), 2009.
- [9] J. M. Bischof and E. M. Airoldi, Capturing semantic content with word frequency and exclusivity, In 29th International Conference on Machine Learning, Edinburgh, Scotland, February 2012.
- [10] G. Salton and C. Buckley, Term weighting approaches in automatic text retrieval, Inf Process Manag, 24 (5) (1988), 513–523.
- [11] G. Salton, Developments in automatic text retrieval, Science, 253 (5023) (1991), 974–980.
- [12] T. Kudo and Y. Matsumoto, A boosting algorithm for classification of semi-structured text, In Conference on Empirical Methods in Natural Language Processing, Barcelona, Spain, Association for Computational Lingusitics, 2004, 301–308.
- [13] R. Schapire and Y. Singer, BoosTexter: a boosting-based system for text categorization, Mach Learn, 39 (2/3) (2000), 135–168.
- [14] A. Genkin, D. D. Lewis, and D. Madigan, Large-scale Bayesian logistic regression for text categorization, Technometrics, 49 (3) (2007), 291–304.
- [15] T. Zhang and F. J. Oles, Text categorization based on regularized linear classification methods, Inf Retr, 4 (2001), 5–31.
- [16] T. Hastie, R. Tibshirani, and J. H. Friedman, The Elements of Statistical Learning (Second Edition). New York, NY, Springer, 2009.
- [17] H. Zou and T. Hastie, Regularization and variable selection via the elastic net, J R Stat Soc Series B Stat Methodol, 67 (2) (2005), 301–320.

- [18] F. Mosteller and D. L. Wallace, Applied Bayesian and Classical Inference: The Case of "The Federalist" Papers, New York, Springer-Verlag, 1984.
- [19] E. M. Airoldi, A. G. Anderson, and S. E. Fienberg, Who wrote Ronald Reagan's radio addresses? Bayesian Anal, 1 (2) (2006), 289–320.
- [20] G. King and W. Lowe, An automated information extraction tool for international conflict data with performance as good as human coders: a rare events evaluation design, Int Organ, 57 (3) (2009), 617–642.
- [21] D. J. Hopkins and G. King, A method of automated nonparametric content analysis for social science, Am J Polit Sci, 54 (1) (2009), 229–247.
- [22] C. Cortes and V. Vapnik, Support-vector networks, Mach Learn, 20 (1995), 273–297.
- [23] G. Forman, An extensive empirical study of feature selection metrics for text classification, J Mach Learn Res, 3 (2003), 1289–1305.
- [24] S. Dumais, J. Platt, and D. Heckerman, Inductive learning algorithms and representations for text categorization, In Proceedings of the Seventh International Conference on Information and Knowledge Management, ACM, 1998, 148–155.
- [25] T. Joachims, Learning to Classify Text Using Support Vector Machines. The Springer International Series in Engineering and Computer Science, New York, NY, Springer, 2002.
- [26] S. Rose, D. Engel, N. Cramer, and W. Cowley, Automatic keyword extraction from individual documents, In Text Mining: Applications and Theory, M. W. Berry and J. Kogan, eds. Chichester, UK, John Wiley & Sons, Ltd, 2010, 1–20.
- [27] E. Frank, G. W. Paynter, I. H. Witten, C. Gutwin, and C. G. Nevill-Manning, Domain-specific keyphrase extraction, In The Sixteenth International Joint Conference on Artificial Intelligence (IJCAI-99), California, Morgan Kaufmann, 1999, 668–673.
- [28] J. Chen, B. Zhang, D. Shen, Q. Yang, Z. Chen, and Q. Cheng, Diverse topic phrase extraction from text collection, In World Wide Web Conference, Edinburgh, Citeseer, 2006, 1–9.
- [29] S. Gerrish and D. M. Blei, Predicting legislative roll calls from text, In The 28th International Conference on Machine Learning (ICML-11), 2011, 489–496.
- [30] L.-W. Lee and S.-M. Chen, New methods for text categorization based on a new feature selection method and a new similarity measure between documents, In Advances in Applied Artificial Intelligence, 2006, 1280–1289.
- [31] Y. Yang and I. O. Pendersen, A comparative study on feature selection in text categorization, In ICML-97, 14th International Conference on Machine Learning, Nashville, TN, 1997, 412–420.
- [32] J. Eisenstein, A. Ahmed, and E. P. Xing, Sparse additive generative models of text, In 28th International Conference on Machine Learning, Bellevue, WA, 2013, 1–8.
- [33] L. E. Ghaoui, G.-C. Li, V.-A. Duong, V. Pham, A. N. Srivastava, and K. Bhaduri, Sparse machine learning methods for understanding large text corpora: application to flight reports, In Conference on Intelligent Data Understanding, 2011, 159–173.
- [34] B. L. Monroe, M. P. Colaresi, and K. M. Quinn, Fightin' words: lexical feature selection and evaluation for identifying the content of political conflict, Polit Anal, 16 (4) (2008), 372–403.

- [35] H. Zou, The adaptive lasso and its oracle properties, J Am Stat Assoc, 101 (476) (2006), 1418–1429.
- [36] T. T. Wu and K. Lange, Coordinate descent algorithms for lasso penalized regression, Ann Appl Stat, 2 (1) (2008), 224–244.
- [37] Z. Q. Luo and P. Tseng, On the convergence of the coordinate descent method for convex differentiable minimization, J Optim Theory Appl, 72 (1) (1992), 7–35.
- [38] L. E. Ghaoui, V. Viallon, and T. Rabbani. Safe feature elimination for the lasso and sparse supervised learning problems. arXiv preprint arXiv:1009.4219 (2010).
- [39] H. Wickham, D. Cook, H. Hofmann, and A. Buja, Graphical inference for Infovis, IEEE Trans Vis Comp Graph, 16 (6) (2010), 973–979.
- [40] T. P. Speed and B. Yu, Model selection and prediction: normal regression, Ann Inst Stat Math, 45 (1) (1993), 35–54.

- [41] J. Friedman, T. Hastie, and R. Tibshirani, Regularization paths for generalized linear models via coordinate descent, J Stat Softw, 33 (1) (2010), 1.
- [42] M. F. Porter, An algorithm for suffix stripping, Program, 14 (3) (1980), 130–137.
- [43] D. Meyer, K. Hornik, and I. Feinerer, Text mining infrastructure in R, J Stat Softw, 25 (5) (2008), 1–54.
- [44] Y. Yang and X. Liu, A re-examination of text categorization methods, In Proceedings of the 22nd Annual International ACM SIGIR Conference on Research and Development in Information Retrieval, ACM, 1999, 42–49.
- [45] A. Buja, D. Cook, and D. F. Swayne, Interactive highdimensional data visualization, J Comp Grap Stat, 5 (1) (1996), 78–99.