

## **THE BABY BOOM, THE BABY BUST, AND THE HOUSING MARKET**

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This paper examines the impact of major demographic changes on the housing market in the United States. The entry of the Baby Boom generation into its house-buying years is found to be the major cause of the increase in real housing prices in the 1970s. Since the Baby Bust generation is now entering its house-buying years, housing demand will grow more slowly in the 1990s than in any time in the past forty years. If the historical relation between housing demand and housing prices continues into the future, real housing prices will fall substantially over the next two decades.

### **1. Introduction**

The dramatic rise in the number of births in the 1950s and the subsequent decline in the 1970s – the Baby Boom and the Baby Bust – are widely recognized as among the most important changes in the United States in the past 50 years. At the peak of the Baby Boom in 1957, 4.30 million babies were born in the United States. The year before the boom began, in 1945, 2.86 million babies were born, and at the trough of the Baby Bust in 1973, the figure was only 3.14 million. In this paper we examine how such major demographic changes affect the market for housing.

Our goals are both retrospective and prospective. We want to assess what impact these major demographic changes have had on the demand for housing and, further, how these changes in demand have affected residential investment and the price of housing. In addition, we want to assess what more recent demographic patterns imply about the housing market over the next twenty years.

Our analysis of both cross-sectional and time-series data leads us to three conclusions. First, large demographic changes of the sort we have observed induce large (and mostly predictable) changes in the demand for housing.

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Second, these fluctuations in demand appear to have substantial impact on the price of housing. Third, recent demographic patterns imply that housing demand will grow more slowly over the next twenty years than in any time since World War II.

These findings have important implications for the public policy debate over housing. Between 1970 and 1980 housing prices rose dramatically: depending on the index, the real price of housing rose between 19 and 32 percent. This development generated many calls for government intervention to help provide more 'affordable' housing. Our results indicate that this increase in housing prices was largely attributable to the aging of the Baby Boom. Over the next twenty years, the Baby Bust generation will be in its house-buying years. As Kenneth Rosen (1984) has emphasized, this implies that housing demand will grow more slowly in the future. Our estimates suggest that real housing prices will fall substantially – indeed, real housing prices may well reach levels lower than those experienced at any time in the past forty years.

Our analysis proceeds as follows. We begin in section 2 by documenting the facts about the Baby Boom. We show the rise and subsequent decline in births and discuss the extent to which these demographic changes were predicted by contemporary observers.

In section 3 we examine the link between age and housing demand. Using cross-sectional data from the Census for years 1970 and 1980, we find that an individual generates little housing demand until age 20 – that is, children do not substantially increase a family's quantity of housing. Housing demand rises sharply between ages 20 and 30, and remains approximately flat after age 30. This finding implies that an increase in the number of births has little immediate effect on the housing market, but generates an increase in housing demand twenty years hence.

We examine in section 4 how demographic changes in the United States have affected the demand for housing. We combine our cross-sectional results on age and housing demand with time series on the age composition of the population. We find that the Baby Boom of the 1950s led to rapid growth in housing demand in the 1970s, and that the Baby Bust of the 1970s will lead to slow growth in housing demand in the 1990s.

In sections 5 and 6 we examine how housing demand affects the price of housing and the amount of residential investment. Section 5 is in the nature of an exploratory data analysis of the impact of changes in housing demand. We are unable to detect a statistically significant relation between demographic housing demand and the quantity of residential capital. Residential investment is such a 'noisy' time series that the standard errors we obtain are very large. There is, however, a significant relation between housing demand and the price of housing: a one percent increase in housing demand leads to a five percent increase in the real price of housing. We use this estimated

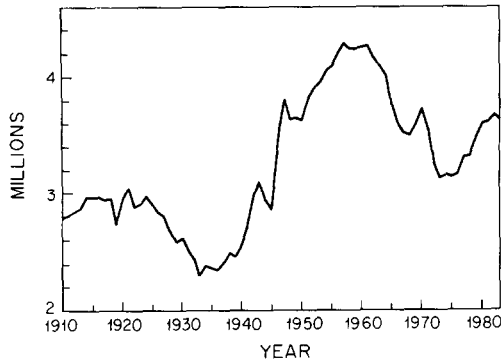


Fig. 1. Live births in the United States 1910-1983.

relationship to examine how the slow growth in demand over the next twenty years will likely affect housing prices.

In section 6 we use an intertemporal model of the housing market, in the spirit of the one proposed by Poterba (1984), to examine the impact of changing housing demand. One implication of our findings is that the Baby Boom caused an increase in housing demand in the 1970s that was predictable far in advance. The model makes precise predictions about how such a forecastable increase in demand should affect the housing market. We examine what the model predicts and compare these predictions to experience. We conclude that the housing market probably should not be characterized as an efficient asset market in which prices reflect available information on future demand.

## 2. The facts about the Baby Boom

Fig. 1, which graphs the number of births over time, shows the Baby Boom very clearly. The low level of fertility during the Great Depression and the boom in births that lasted from 1946 to 1964 combine to produce a sharp step in the population structure. As this step aged, it had effects on the educational system, the labor market, the housing market, and the social security system.

One way to look at the magnitude of the Baby Boom is to look at the number of people at a given age. In 1960, 24.0 million people, or 13.3 percent of the population, were between ages 20 and 30; in 1980 the corresponding number was 44.6 million, or 19.7 percent of the population. Since this is the age in which people are forming new households and increasing their demand for housing, it is clear that the boom should have had a large effect on the housing market.

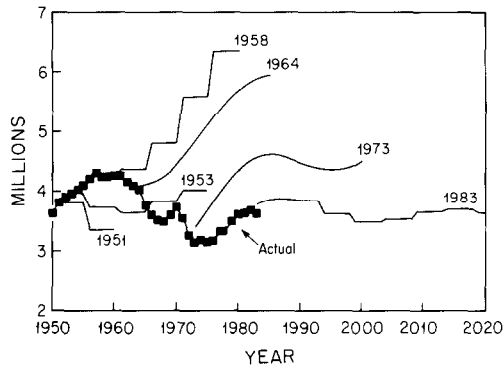


Fig. 2. Number of births: Actual and Census Bureau projections.

For our purposes, the exact cause of the Baby Boom is not important.<sup>1</sup> That the bulge in the population was significant and could be expected to move up through the age structure is clear enough that it does not need to be defended. Whether the boom was seen as being temporary or permanent is a tougher question to answer.

Fig. 2 presents the actual number of births per year from 1950 through 1983, and several contemporary forecasts made by the Census Bureau.<sup>2</sup> The lesson to be learned from looking at these forecasts is that forecasting births is a risky business. Any forecast of housing demand that depended sensitively on births would be highly suspect. Luckily, as we shall show below, forecasts of housing demand depend (as a first approximation) only on the population above the age of 20. Thus, housing demand can be forecast 20 years into the future before the unreliability of birth forecasts becomes a problem.<sup>3</sup>

### 3. Cross-sectional evidence on the demand for housing by age

We are interested in how housing demand is affected by changes in the

<sup>1</sup>Russell (1982) notes that the boom was caused by increases in the number of women who married and the number of children per married woman, and the fact that married women tended to have children earlier. In terms of fertility, the boom can be seen in the number of births per 1,000 women aged 15 through 44, which jumped from a depression low of below 80 to a peak of above 120 around 1960, and fell to below 70 by 1980.

<sup>2</sup>The Census Bureau generally provides several forecasts for births, based on different assumptions about fertility. In cases where three forecasts were made, we took the middle one; in cases where four were made, we took the average of the middle two.

The series of actual births for the years before 1959 was subsequently adjusted upward to reflect underregistration. The result is that census forecasts are well below the actual number of births (even over short horizons for which birth forecasts should be highly accurate). We therefore adjust forecasts made before 1959 by a constant multiple computed by assuming that the first year of any forecast was correct.

<sup>3</sup>The 1983 forecast has been more accurate than most of its predecessors, at least so far. Actual births in 1987 were 3,829,000, compared to a forecast of 3,879,000.

size of different age cohorts. We begin our examination of this issue by using cross-sectional data to determine the link between age and the quantity of housing demanded.

Looking across individuals, the quantity of housing demanded is a function of age, income, and a variety of other household characteristics. Yet here we use data on only the first of these attributes: age. Our ultimate goal is to construct a variable on the aggregate demand for housing given information only on the age composition of the population. We are therefore not interested in the value of the true coefficient on age in a multiple regression. Instead, we are interested in the best predictor of a household's quantity of housing given information only on the age of its members. Any correlation of age with income and other household characteristics does not pose a problem – indeed, multicollinearity may be a strength, for it acts to eliminate any worry over the role of omitted variables.

We model demand for housing by a household as an additive function of the demand for housing of its members:

$$D = \sum_{j=1}^N D_j, \quad (1)$$

where  $D_j$  is the demand of the  $j$ th member and  $N$  is the total number of people in the household. To the extent that there are economies of scale in the provision of housing services, this would not be the best way to estimate the housing demand of a given household. Yet if we are interested in predicting the housing demand of an entire population, and if the extent of household formation is fairly constant, then our approach should be accurate.<sup>4</sup>

The demand for housing of each individual is taken to be a function of age. We allow each age to have its own housing demand parameter, so that an individual's demand is given by

$$D_j = \alpha_0 DUMMY0_j + \alpha_1 DUMMY1_j + \cdots + \alpha_{99} DUMMY99_j, \quad (2)$$

where  $DUMMY0 = 1$  if age = 0,  $DUMMY1 = 1$  if age = 1, etc. The parameter  $\alpha_i$  tells us the quantity of housing demanded by a person of age  $i$ .

Combining (1) and (2) gives the equation for household demand:

$$D = \alpha_0 \sum DUMMY0_j + \alpha_1 \sum DUMMY1_j + \cdots + \alpha_{99} \sum DUMMY99_j, \quad (3)$$

<sup>4</sup>Hendershott (1987) studies the effects of changes in the propensity to form households on the demand for housing. We do not deny that such effects are important, but our primary interest is in changes in demand that are forecastable; we do not think that such changes are nearly as forecastable as changes in the age structure of the population.

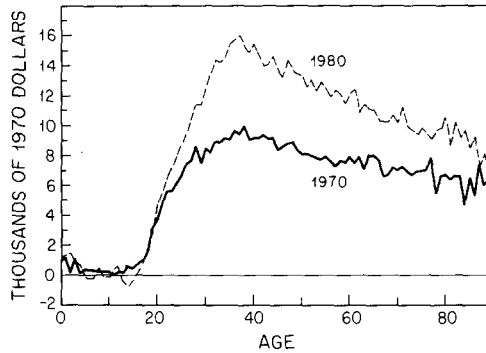


Fig. 3. Estimated housing demand by age.

We estimated (3) on a 1-in-1,000 sample of the 1970 Census. The sample consists of 203,190 individuals grouped into 74,565 households. The left-hand side variable is the value of the property for the unit in which the household resides. For owner-occupied units this is reported directly. In the case of rental units, we used the approximation that the value is equal to 100 times the gross monthly rent.<sup>5</sup> Leaving out units for which neither of these figures was available, our sample consisted of 53,518 households.

The solid line in Fig. 3 plots the estimated  $\alpha$ 's, while the estimates themselves are in table A.1 in the appendix. Since the sample is so large, the standard errors are extremely small (less than \$300 for all of the estimates below age 64). The dotted line in fig. 3 plots the  $\alpha$ 's for the same regression run on a sample from the 1980 Census, deflated into 1970 dollars by the GNP deflator.

The primary feature of the estimates is a sharp jump in the demand for housing between the ages of 20 and 30. As mentioned earlier, people below the age of 20 have little impact on the demand for housing. The result is qualitatively the same for 1970 and 1980 data. In the work below we use only the estimates from 1970.

A second feature of the results for both 1970 and 1980 is that the quantity of housing demanded appears to decline after age 40 by about one percent per year. This decline is probably attributable to the fact that, because of productivity growth, older cohorts have lower lifetime income than younger cohorts and therefore demand less housing.

A third feature of fig. 3 is the large shift upward between 1970 and 1980. The real value of housing for an adult of any given age increased almost 50 percent over this decade. Part of this increase is attributable to productivity growth: real disposable personal income per capita rose 22 percent from 1970

<sup>5</sup>To test the robustness of this approximation, we ran eq. (3) leaving out rental units and also with the value/rent ratio set to 80, 90, 110, and 120. The results were quite similar to our baseline case.

to 1980. But much of the rise in house value must be attributable to the 20 to 30 percent increase in the real price of housing. As long as the price elasticity of housing demand is less than one, an increase in the price of housing will increase the value of housing. The large increase in age-specific house value between 1970 and 1980 thus suggests that housing demand is fairly inelastic.

#### 4. Shifts in housing demand due to the Baby Boom

Here we examine how changes in the age composition of the population affect the demand for housing over time. Our approach is to assume that the age structure of housing demand (that is, the set of  $\alpha$ 's estimated in the last section) is constant over time. We can then see how the age structure of housing demand interacts with shifts in the age structure of the population.<sup>6,7</sup>

More precisely, to obtain a measure that we interpret as the shift in housing demand due to demographics, we multiply the age structure of the population by the coefficients estimated in eq. (3) and sum for all cohorts. That is, if  $N(i, t)$  is the number of people of age  $i$  in year  $t$ , then housing demand in year  $t$  is

$$D_t = \sum_i \alpha_i N(i, t). \quad (4)$$

We use the  $\alpha$ 's from the 1970 cross-sectional demand for housing. This time series on housing demand, which is measured in millions of 1970 dollars, is presented in table A.2 in the appendix. The growth rate of housing demand is plotted in fig. 4. For comparison, we also present in fig. 4 the growth in housing demand computed with the 1980  $\alpha$ 's.

The arrival of the Baby Boom in the housing market appears clearly as a swelling in the rate of growth of demand that peaks around 1980. The rate of increase in housing demand from 1940 to 1950 was 1.84 percent per year; from 1950 to 1960, 1.16 percent; from 1960 to 1970, 1.31 percent; and from 1970 to 1980, 1.66 percent. Our forecast is that the rate of growth from 1980 to 1990 will be 1.33 percent per year; from 1990 to 2000, 0.68 percent, and from 2000 to 2010, 0.57 percent.

It is instructive to compare our estimate of the growth in housing demand with simpler demographic variables. Since our cross-sectional estimates indicate a large increase in housing demand from age 20 to 30 and

<sup>6</sup>To obtain the age structure of the population on an annual basis we combined data on births with estimates of mortality. Actual births are used through 1983, and the Census Bureau median forecast thereafter.

<sup>7</sup>Our technique is similar to that employed by Hickman (1974); in place of our estimated  $\alpha$ 's he uses age-specific rates of household headship.

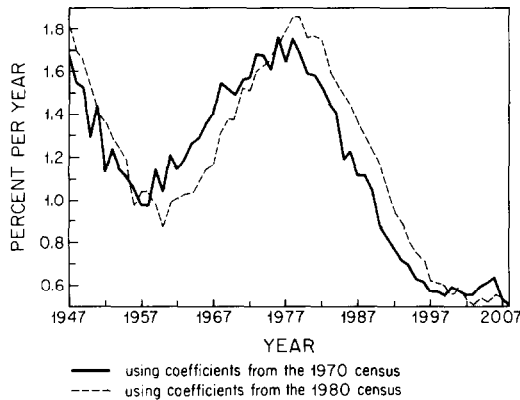


Fig. 4. Growth rate of housing demand.

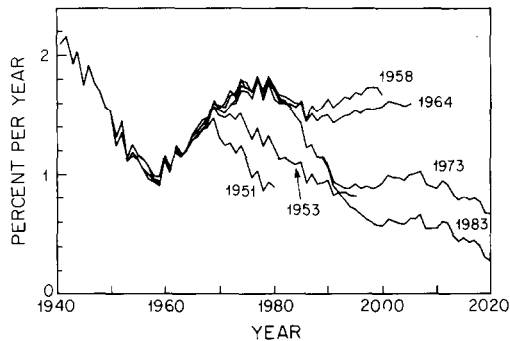


Fig. 5. Housing demand growth rate: Actual and projected.

*Note:* Each series is marked with a date indicating the year of the birth forecast on which it is based. For example, the series marked 1973 uses actual births through 1973 and the forecast of future births made in that year.

approximate constancy thereafter, our time series on housing demand is not very different from a time series on the adult population. The correlation between the growth in the population over 21 and the growth in our housing demand variable is 0.86. Our estimate of housing demand is, however, very different from the population including children. The correlation between the growth in the total population and the growth in our housing demand variable is  $-0.57$ . Hence, although our housing demand variable is quite similar to the adult population, it is not at all approximated by the total population.



Fig. 5 shows the forecasts of change in demand that would have been made using Census birth projections starting at various points in the postwar period. Despite the fact that birth projections were not very accurate, in every case forecasted demand growth tracks actual demand growth quite well for twenty years after the forecast is made. Because of the low demand for housing of children, forecasts of total housing demand made in the 1960s would have correctly predicted the increase in the rate of growth of housing demand in the 1970s.

## 5. From housing demand to prices and quantities

In the last section we combined the cross-sectional results on age and the quantity of housing demanded with time-series data on the age-composition of the population to generate a new time series on housing demand. This time series shows that the Baby Boom profoundly affected the demand for housing in the 1970s and that the Baby Bust will have the opposite impact on the housing market over the next twenty years. Our goal now is to examine the link between housing demand as measured by this time series and developments in the housing market.

We take two approaches to examining how these fluctuations in housing demand affect the housing market. Our first approach, which is pursued in this section, is statistical and relatively atheoretical: we examine how our time series on housing demand correlates with data on the housing market. Our second approach, which is pursued in the next section, is more theoretically correct but is relatively data-free: we calibrate a variant of Poterba's model of the housing market and examine how, according to that model, large and predictable shifts in housing demand should affect the housing market. We also examine the extent to which available evidence is consistent with the model.

The reason we call the statistical analysis of this section relatively atheoretical is that any good theory of the housing market must take into account many subtle intertemporal issues. At any point in time, the stock of housing depends on past flows of investment; the flow of investment depends on the price of housing; the price of housing depends on current and expected future rents; the rent depends on the stock and the state of housing demand. While the model of the next section incorporates all these feedbacks, here we ignore them. The goal of the exploratory data analysis of this section is to see what stylized facts emerge.

### 5.1. Quantities

We begin by looking at whether there is any correlation between our housing demand variables and the quantity of housing. We measure the

Table 1  
 Housing demand and the housing stock.<sup>a</sup>  
 Dependent variable:  $\log(\text{stock})$ ; sample period: 1947–1985.

<i>constant</i>	8.01 (7.81)	5.14 (6.35)	4.99 (7.28)
<i>time</i>	0.0095 (0.0366)	−0.0006 (0.0419)	−0.0006 (0.0436)
$\log(\text{demand})$	0.010 (0.652)	0.173 (0.547)	0.182 (0.574)
$\log(\text{gnp})$		0.149 (0.036)	0.149 (0.037)
<i>cost of funds</i>			−0.00003 (0.00073)
$\rho$	0.971 (0.035)	0.976 (0.031)	0.976 (0.034)
$\bar{R}^2$	0.9996	0.9997	0.9997
<i>DW</i>	1.28	1.13	1.13
<i>see</i>	0.00704	0.00581	0.00590

<sup>a</sup>Standard errors are in parentheses.

quantity of housing as the net stock of residential capital.<sup>8</sup> In table 1 we regress the log of the stock of housing on a time trend and the log of our demographic housing demand variable. We correct for serial correlation by allowing the residual to follow a first-order autoregressive process, which is estimated using an iterated Cochrane–Orcutt procedure.

The results are disappointing. The point estimate of the coefficient for housing demand is near zero, but the standard error is extremely large. We cannot reject that housing demand has no impact on the stock of housing. We also cannot reject that housing demand has a proportionate impact on the stock of housing – that is, that the coefficient is one.

Residential investment is of course a highly cyclical component of GNP. So perhaps it is not surprising that the standard error we obtain is large. In an attempt to reduce the residual variance, and thus obtain more precise estimates, we include the log of real GNP in column 2 of table 1. As expected, GNP enters significantly and positively. (Since the equation is almost differenced, it is essentially relating residential investment – the change in the stock – to the change in GNP. Thus, we are picking up the

<sup>8</sup>End of year total net stock of residential capital, constant cost valuation, from *Fixed Reproducible Tangible Wealth*. In this measure, housing is valued at a base year price regardless of the price in the year of acquisition – thus it corresponds to a ‘physical volume’ measure (and not to price times quantity). In 1985 residential capital was made up of 67.9% owner occupied non-farm; 27.9% private tenant occupied non-farm; 1.7% owner occupied farm; and 2.0% federal, state, and local.

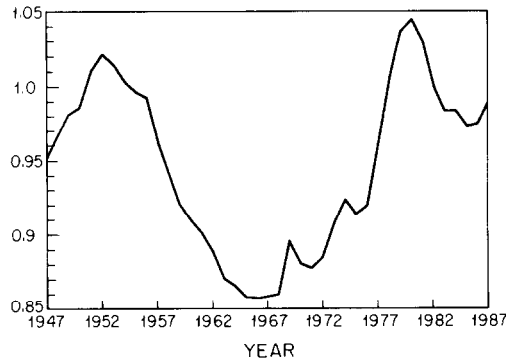


Fig. 6. Real price of housing.

investment accelerator.) The coefficient on demand and its standard error, however, are not substantially affected by correcting for the business cycle.

The next correction we attempt is for the after-tax real interest rate, which affects the user cost of housing and thus housing demand. The after-tax real interest rate is  $(1 - \tau)i - \pi$ , where  $\tau$  is the marginal tax rate,  $i$  is the nominal interest rate, and  $\pi$  is expected inflation. We take  $\tau$  to be 0.3,  $i$  to be the yield on long-term Treasury bonds,<sup>9</sup> and  $\pi$  to be the average rate of change in the GNP deflator over the past two years. When this variable is entered into the regression, in column 3 of table 1, it is not significant and does not alter any of the other estimates.

In summary, we cannot find a relation between our demographically driven housing demand variable and the stock of housing. Yet we cannot conclude there is no relation. Residential investment is just too 'noisy' to allow any firm inference.

## 5.2. Prices

Next we examine whether there is any relation between our housing demand variable and the price of housing. We run regressions analogous to those above: we regress the log of the real price of housing on a time trend and the log of our demographic housing demand variable. We also include the log of real GNP and the after-tax interest rate to correct for other macroeconomic effects.

The time series on housing prices we use is the residential investment deflator relative to the GNP deflator. It is displayed in fig. 6. We choose this series because it is available for a long sample. For the period during which it overlaps with other existing series, such as the median sales price of

<sup>9</sup>From *International Financial Statistics*; from 1953 onward, these are 20 year constant maturities.

Table 2  
 Housing demand and the real price of housing.<sup>a</sup>  
 Dependent variable:  $\log(\text{price})$ ; sample period: 1947–1987.

<i>constant</i>	–63.1 (9.2)	–70.3 (8.7)	–73.4 (7.9)
<i>time</i>	–0.065 (0.010)	–0.078 (0.010)	–0.081 (0.009)
$\log(\text{demand})$	4.65 (0.68)	5.04 (0.62)	5.29 (0.56)
$\log(\text{gnp})$		0.260 (0.098)	0.234 (0.097)
<i>cost of funds</i>			–0.0035 (0.0021)
$\rho$	0.770 (0.102)	0.757 (0.114)	0.690 (0.109)
$\bar{R}^2$	0.940	0.950	0.952
<i>DW</i>	1.29	1.44	1.49
<i>see</i>	0.0152	0.0139	0.0136

<sup>a</sup>Standard errors are in parentheses.

existing single family houses collected by the National Association of Realtors, the different series move closely together. In particular, all series show real housing prices rising sharply in the 1970s and relatively flat in the 1980s.

In contrast to the results for the quantity of housing, the results for housing prices are encouraging. The results in table 2 show a strong and highly significant relation between housing demand and the real price of housing. In the full specification in column 3, the coefficient on housing demand is 5.3, which implies that a one percent increase in the demand for housing leads to a 5.3 percent increase in the real price of housing. The *t*-statistic on this coefficient exceeds 9 – the correlation between our housing demand variable and the price of housing is unlikely to be spurious.<sup>10,11</sup>

<sup>10</sup>The results in table 2 also show that the relative price of housing is procyclical – a one percent increase in real GNP is associated with a 0.23 percent increase in the relative price of housing. While the point estimates imply that high interest rates exert a depressing effect on housing prices (a one percentage point increase in the after-tax real interest rate depresses housing prices by 0.35 percent), this effect is not statistically significant.

<sup>11</sup>Replacing our demand variable with the adult population in the regressions produces almost identical results. When both our demand variable and the adult population are included, our demand variable works slightly better as measured by the *t*-statistic or the size of the coefficient. By contrast, when our demand variable is replaced with the total population, the total population enters with the wrong sign. When both our demand variable and the total population are included, the coefficient on our demand variable remains positive and significant while the coefficient on the total population is negative and insignificant. Taken together, these time-series results strongly support the validity of the housing demand variable generated from the cross-section estimates and the conclusion that children generate little demand for housing.

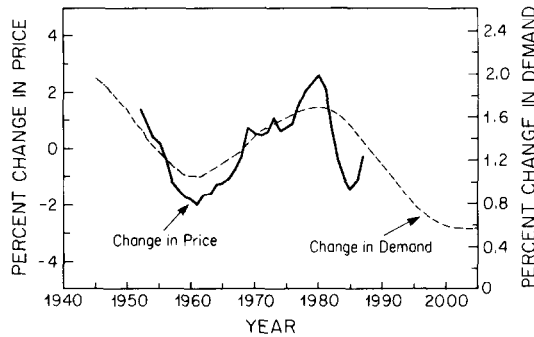


Fig. 7. Rates of growth: Demand and prices.

Note: Series are average rate of change over the past five years.

The strong association between housing demand and housing prices also appears when one simply plots the data. Fig. 7 shows the percentage change over the previous five years in the real price of housing and in our housing demand variable. The demographic shifts appear to explain the increase in housing prices around 1950, the fall around 1960, and the housing boom throughout the 1970s.

Our finding that shifts in housing demand have great impact on housing prices suggests that both the supply and demand for housing are highly inelastic. To see why, consider the following static model of housing supply and demand:

$$H^s = a + bP, \quad b > 0,$$

$$H^d = c - dP + D, \quad d > 0,$$

where  $H$  is log quantity of housing,  $P$  is log price, and  $D$  is an exogenous shift in demand. These two equations imply that

$$P = (c - a)/(b + d) + (1/(b + d))D.$$

For the coefficient on housing demand to be large, as we found in table 2, the sum of the supply elasticity and the demand elasticity ( $b + d$ ) must be small. Of course, the housing market is not adequately described in such a static model. Yet even within dynamic models, such as that examined in the next section, the substantial impact of demand on prices can only be explained if supply and demand elasticities are small.

Since we have found a highly significant relation between housing demand and housing prices, it is natural (at least for the heroic) to extrapolate this

relation forward to see what it implies for future housing prices. As we have emphasized earlier, the changes in housing demand caused by changes in birth rates are forecastable far in advance. Therefore, we can be confident about our predictions regarding future housing demand.

The implication for future housing prices is perhaps apparent from fig. 7, which graphs the percentage increase in housing prices and the percentage increase in housing demand. It shows that housing demand will grow more slowly over the next twenty years than at any time in our sample. If the historical relation between demand and prices continues to hold, it appears that the real price of housing will fall about 3 percent a year. More formal forecasting using the regressions yields the same answer. The regression in the first column of table 2 implies that real housing prices will fall by a total of 47 percent by the year 2007. Thus, according to this forecasting equation, the housing boom of the past twenty years will more than reverse itself in the next twenty.

At this point we should interject a note of caution about this forecast. Every good student of econometrics can recite the perils of forecasting beyond the experience of the data. The predicted growth of our housing demand variable is lower than has been experienced over the past forty years, and the period of low growth is protracted. Hence, we cannot be confident about precisely what effects this slow growth will have. Yet experience does tell us that slow growth in demand is associated with falling prices. Even if the fall in housing prices is only one-half what our equation predicts, it will likely be one of the major economic events of the next two decades.

## 6. Housing demand in an intertemporal model

We now turn to examining the impact of changing demand in an intertemporal model of the housing market. The model that we use is a slight variation on that of Poterba (1984). In contrast to Poterba, we ignore issues of taxation and of the effect of inflation on the cost of owning a home, and concentrate on the effects of changes in demand attributable to demographic changes.<sup>12</sup>

### 6.1. *The elements of the model*

Let  $H$  be the stock of housing. We assume that the flow of housing services is proportional to the stock of housing. The demand for housing is given by the equation

<sup>12</sup>The model we examine is partial equilibrium in nature. For a general equilibrium treatment of some of these issues, see Manchester (1988).

$$H^d = f(R)N, \quad f' > 0,$$

where  $R$  is the real rental price and  $N$  is the adult population.  $N$  is a shift variable which is meant to capture the effect of demographic changes of the type discussed in section 4. The market-clearing rent is thus given by

$$R = R(h), \quad R' < 0,$$

where  $h = H/N$  is housing per adult, and  $R(\cdot)$  is the inverse of  $f(\cdot)$ .

We let  $P$  represent the real price of a standardized unit of housing. (We ignore any distinction between land and structures.) For simplicity we assume that the operating cost of owning a home is some constant,  $r$ , times the value of the house. This constant is meant to incorporate the opportunity cost of capital, property taxes, maintenance, and depreciation. The arbitrage condition for the path of housing prices is

$$R(h) = rP - \dot{P}. \tag{5}$$

This equation says that the rent must equal the user cost, which equals the operating costs minus the capital gain. This implies

$$\dot{P} = rP - R(h). \tag{6}$$

This equation tells us how the price of housing evolves over time.

Gross investment in housing is taken to be an increasing function of the price of housing and proportional to the scale of the economy as measured by the adult population:

$$\dot{H} + \delta H = \psi(P)N, \quad \psi' > 0,$$

where  $\delta$  is the rate of depreciation. Let  $n$  be the rate of growth of the population – that is,  $n = \dot{N}/N$ . We can rewrite this equation terms of  $h$  rather than  $H$ . Differentiating  $H/N$  with respect to time and substituting gives

$$\begin{aligned} \dot{h} &= \dot{H}/N - n(H/N) \\ &= \psi(P) - (n + \delta)h. \end{aligned} \tag{7}$$

Population growth,  $n$ , thus enters as a shift variable in the model.

Fig. 8 combines eqs. (6) and (7) to give the familiar phase plane representation of the housing market. In steady state, the state variable  $h$  and the costate variable  $P$  are constant. The arrows show the implied dynamics

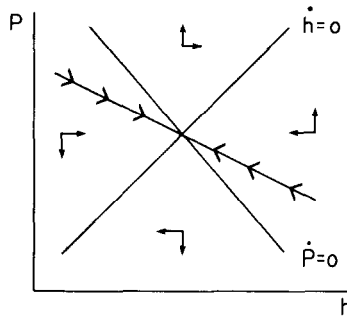


Fig. 8. The model's dynamics.

of the economy when it is out of steady state. For any given value of  $h$ ,  $P$  jumps to the stable arm and the economy converges to the steady state.

### 6.2. *Simulating a Baby Boom*

Now consider the effects of a hypothetical Baby Boom. The economy is in a steady state with growth at a rate of one percent per year. In hypothetical year 1960 it is announced that from years 1970 through 1979, the growth rate of the adult population will be two percent per year, after which it will return permanently to one percent. As can be seen in the phase diagram in fig. 9, the price of housing jumps up upon the announcement, rises until sometime in the middle of the high growth period, then gradually falls back to its steady state level.

To get a feel for the potential magnitude of the price changes, we simulate the model under the assumption that the demand elasticity is  $1/2$ , the supply elasticity is 1, the operating cost  $r$  is 5 percent, and the depreciation rate  $\delta$  is

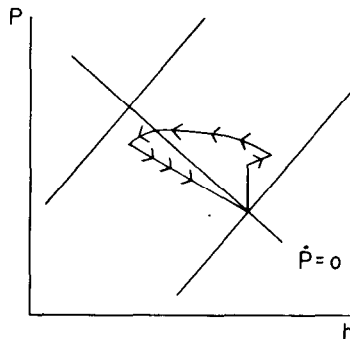


Fig. 9. A baby boom: An anticipated temporary increase in population growth.



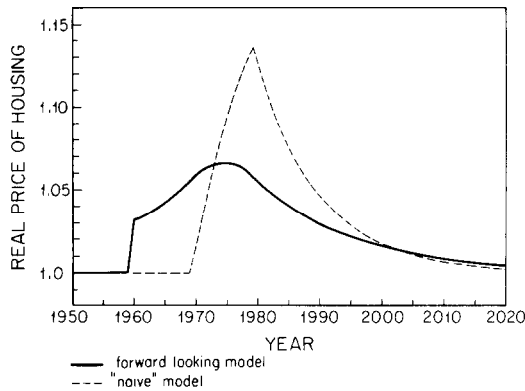


Fig. 10. Response of housing price to a theoretical baby boom.

2 percent. Although these supply and demand elasticities are somewhat smaller than is generally accepted, we choose them to generate large price responses. Rosen (1979) estimates that the elasticity of demand is about 1.<sup>13</sup> Poterba (1984) estimates that the supply elasticity is between 0.5 and 2, while Topel and Rosen (1988) estimate that the supply elasticity is about 3. After studying our base case, we will consider the price responses generated by these alternative parameters.

The solid line in fig. 10 plots the simulated path of prices. We see that upon the announcement of the Baby Boom, the price of housing jumps about three percent. From 1960 to 1970, the price rises about 3 percent more in anticipation of the increased demand. The price rises an additional one percent from 1970 to 1975, and then gradually falls back to the original level. Thus, the price changes generated by the model under perfect foresight are not very large, and almost all of the price rise takes place before the increased demand arrives.

In assessing this forward-looking model of the housing market, it is instructive to consider a simple alternative: suppose that market participants are 'naive' in the sense that, whatever the price of housing at a given time, they expect it to remain constant at that level. If this is the case, there are no expected capital gains, and the price of housing is determined simply by the rental market. That is,

$$P = R(h)/r.$$

<sup>13</sup>Rosen's estimate, which is based on cross-sectional data, should probably be considered a long-run elasticity. Over short horizons, perhaps even as long as a decade, the demand elasticity is likely smaller. Moving costs play a key role here, since they make the demand elasticity zero for many people. One weakness of the intertemporal model used here is that it does not distinguish between long-run and short-run demand elasticities.

Table 3  
Forward-looking and naive forecasts under alternative parameters.<sup>a</sup>

Parameters		Forward-looking model			Naive model
Demand elasticity	Supply elasticity	Initial jump	Value in 1970	Peak value (year)	Peak value
1/2	1/2	4.3	7.5	8.4 (1975)	15.8
1/2	1	3.2	5.9	6.6 (1975)	13.6
1/2	2	2.0	4.2	4.8 (1974)	10.2
1/4	1	4.1	8.6	9.9 (1974)	21.5
1/2	1	3.2	5.9	6.6 (1975)	13.6
1	1	2.1	3.7	4.1 (1975)	7.6

<sup>a</sup>All values are expressed in percent differences from the steady state. The hypothetical Baby Boom modelled is an increase in the growth rate of demand from one percent to two percent, lasting from 1970 to 1979, this increase having been announced in 1960.

In the phase diagram, the economy is always on the  $\dot{P}=0$  schedule. One simple consequence of this assumption is that the anticipation of a change in demand growth can have no effect at all on the market.<sup>14</sup>

The dotted line in fig. 10 plots the simulated path of  $P$  under naive expectations. In contrast to the forward-looking model, housing prices do not begin to rise until the beginning of high growth in 1970 and they reach their peak at the end of the period of high growth in 1980. The price changes are also much larger. From 1970 to 1980, the price rises by 14 percent – about twice as large as the total change under perfect foresight.

In table 3 we explore the effects of alternative assumptions about the demand and supply elasticities on the path of prices in both the forward-looking and naive expectations models. The hypothetical Baby Boom that we consider is the one described above. For each set of elasticity parameters, we give the total amount of the price increase in each model, and for the forward-looking model we also present the amount of the initial jump in price and the level that the price has reached in 1970, just as the actual boom in demand is beginning.

The result that almost all of the increase in price in the forward-looking model takes place before the actual boom in demand arrives is quite robust to the choice of elasticities. In no case that we look at does more than one-fifth of the total increase take place after the boom has arrived. The result that the total increase in price under the naive model is about twice as large as that under the forward-looking model is also robust. Raising either the supply or demand elasticity lowers the total amount of the price increase in

<sup>14</sup>The essence of this naive model is that the quantity of housing demanded depends on the current price of housing and not on the anticipated capital gain or loss. Other reasons beyond naive expectations can potentially explain this behavior – a binding borrowing constraint is one example.

either model. In general, alternative assumptions about the supply and demand elasticities do not change the qualitative properties of either model. Setting these elasticities as high as suggested by some of the literature discussed above, however, does reduce the size of the price increase in both models, but especially in the forward-looking model, to near insignificance.

### 6.3. *Does the model fit experience?*

While we do not formally test the forward-looking model of the housing market, there are several reasons to think that it cannot come close to fitting the data. First, consider the timing of the run-up in prices in the 1970s. Both housing prices and our housing demand series rose swiftly in this decade. But the increase in demand growth could have been perfectly predicted ten years in advance. In a forward-looking model, most of the increase in prices should have taken place *before* the increase in demand actually arrived.

Similarly, the forward-looking model does not properly capture the timing of the turn-around in prices. An examination of fig. 7 shows that housing prices peaked at almost exactly the time that the demand growth began declining. The forward-looking model implies that prices should turn down *before* demand growth. The model with naive expectations described at the end of the last section, by contrast, does have the property that prices turn down at the same time as demand.

Second, consider the magnitude of the price increase. The arbitrage condition in the forward-looking model makes it difficult for prices to rise very quickly in the absence of news. The forward-looking model reacts to our simulated 'Baby Boom' with a price increase of seven percent – far from the 20 to 30 percent rise observed from 1970 to 1980. Again the model with naive expectations comes closer to matching the facts: the total rise in prices in response to a simulated Baby Boom in this model is much greater than in the forward-looking model, and it takes place over a shorter period of time.

As a means of salvaging the forward-looking model, one might argue that the rise in prices in the 1970s was due not to the anticipated demand increases in that decade but to the gradual arrival of 'news' about future demand growth. Fig. 5 shows, however, that considering the arrival of news only makes the forward-looking model look worse. The positive news about demand in our sample period arrived during the 1950s, when it became clear that the forecasts of births made in the early 1950s were too low. During the 1970s by contrast, the news that arrived was negative: the low birth rates of the decade showed that earlier forecasts were too high. News about births in the 1970s should have made housing prices fall.

### 6.4. *Is the housing market an efficient asset market?*

Our simulation of the intertemporal model suggests that naive expect-

tations better characterizes the housing market than does perfect foresight. In other words, the fluctuations in prices caused by fluctuations in demand do not appear to be foreseen by the market, even though these fluctuations in demand were foreseeable (at least in principle). Thus, the arbitrage condition (6) appears not to characterize housing prices.

More direct tests also suggest that the housing market cannot be viewed as an efficient asset market in which prices fully reflect available information and returns are unforecastable. We can test the proposition that real housing prices follow a random walk by regressing the change in the log of housing prices on the change in the log of our demographic demand variable. We obtain, with standard errors in parentheses,

$$\Delta \log P = -0.06 + 4.7 \Delta \log D,$$

(0.02) (1.1)

$$N = 40, \quad D.W. = 1.36, \quad \bar{R}^2 = 0.30, \quad s.e.e. = 0.016.$$

Remember that this housing demand variable, which forecasts 30 percent of real capital gains in housing, is known about 20 years in advance. Thus, housing prices are not all a random walk.

The failure of the random walk hypothesis for housing prices, however, need not imply the failure of (6) and the existence of profit opportunities. If the rent-price ratio moves in the opposite direction from the expected capital gain, then the total return (rent and capital gain together) could be unforecastable. In fact, using the CPI's component for rent, we find that the rent-price ratio is negatively related to next year's capital gain:<sup>15</sup>

$$\Delta \log P = 0.03 - 0.024R/P,$$

(0.02) (0.018)

$$N = 40, \quad D.W. = 0.99, \quad \bar{R}^2 = 0.02, \quad s.e.e. = 0.019.$$

Yet this statistical relation is very weak: the  $R^2$  is far smaller for the rent-price ratio than for the change in demand. When both regressors are included, we obtain

$$\Delta \log P = -0.11 + 5.8 \Delta \log D + 0.024R/P,$$

(0.04) (1.4)                      (0.019)

<sup>15</sup>Unfortunately, since we have only an index of rents, we cannot compute the total return on housing and examine directly whether it is related to our demand variable. This also implies that the coefficient on  $R/P$  cannot be easily interpreted. We should note that Apgar (1987) has argued that the CPI for rent is a bad measure because of changes in the quality of the rental units over time.

$$N = 40, \quad D.W. = 1.41, \quad \bar{R}^2 = 0.32, \quad s.e.e. = 0.016.$$

The change in demand remains significant, while the rent–price ratio has the wrong sign. In contrast to what an efficient market would require, the rent–price ratio is not the best predictor of the capital gain. It is possible, of course, that if we had better data on rents, we would find the evidence more favorable to the efficient markets hypothesis. Based on the available evidence, however, it seems that the housing market should not be viewed as an efficient asset market.<sup>16</sup>

## 7. Conclusion

We have documented that changes in the number of births over time lead to large and predictable changes in the demand for housing. These changes in housing demand appear to have substantial impact on the price of housing. If the historical pattern continues over the next twenty years, housing prices will fall to levels lower than observed at any time in recent history.

Does our finding imply that readers of this paper should sell their homes and become renters? There are at least three reasons that such an action may not be called for. First, there continue to be substantial advantages to home ownership. Some of these advantages are attributable to the tax code and some are attributable to solving the principal–agent problem that exists between landlord and tenant. Second, there is substantial uncertainty about future housing prices. Not only are there unforeseeable macroeconomic developments, but individual regions of the country will experience housing booms or busts. The best way to hedge the uncertainty about future housing costs is to pay them in advance – that is, to be a homeowner. Third, most homeowners have unrealized capital gains (at least in nominal terms). Becoming a renter requires realizing these capital gains and paying tax at a current rate of 28 percent. For these three reasons, there is no easy way for the typical person to take advantage of advance knowledge of a fall in housing prices.

What effect will the fall in housing demand have on the economy as a whole? It is of course difficult to judge. Since it appears that current housing prices do not fully reflect low future demand, the United States may be currently overinvesting in residential capital. When such drop in demand does become apparent, it is conceivable that we will see a large and sudden drop in housing prices and residential investment, which may be a potential source of macroeconomic instability. Falling housing prices may also induce increases in saving, as individuals perceive their housing equity as insufficient to fund their retirement. The macroeconomic effects of falling housing demand appear to be a fruitful topic for future research.

<sup>16</sup>For further evidence on this question, see Case and Shiller (1988).

## Appendix

Table A.1  
The cross-sectional estimates of housing demand by age.<sup>a</sup>

Age	$\alpha$	Age	$\alpha$	Age	$\alpha$
0	857	34	9,091	68	6,694
1	1,175	35	9,006	69	7,146
2	180	36	9,608	70	6,976
3	1,066	37	9,360	71	7,233
4	110	38	9,856	72	6,918
5	385	39	8,994	73	6,660
6	371	40	9,122	74	6,896
7	340	41	9,096	75	6,968
8	221	42	9,246	76	7,012
9	244	43	9,017	77	7,816
10	211	44	9,052	78	5,416
11	10	45	8,326	79	6,635
12	188	46	8,532	80	6,716
13	143	47	8,731	81	6,343
14	536	48	8,805	82	6,652
15	392	49	8,314	83	6,627
16	639	50	7,999	84	4,666
17	911	51	8,085	85	6,506
18	1,498	52	7,901	86	5,241
19	3,065	53	7,780	87	7,614
20	3,673	54	7,699	88	6,028
21	4,623	55	7,884	89	6,347
22	5,629	56	7,528	90	8,309
23	5,578	57	7,207	91	6,407
24	6,138	58	7,645	92	6,756
25	6,678	59	7,487	93	6,091
26	7,463	60	7,893	94	6,664
27	7,647	61	7,423	95	7,222
28	8,491	62	7,871	96	3,850
29	7,453	63	7,010	97	2,716
30	8,404	64	7,964	98	4,777
31	8,130	65	7,961	99	2,318
32	8,879	66	7,654		
33	8,864	67	6,591		

<sup>a</sup>Estimates are for 1970, as described in text.  $\alpha$  is expressed in 1970 dollars. Standard errors are in the range \$180 to \$225 for ages 0 through 27; \$225 to \$360 for ages 28 through 70; and remain under \$1,000 through age 87. The *R*-squared for this regression is 0.70.

Table A.2  
The time series on housing demand (D).\*

1940	701,244	1971	1,091,308	2002	1,566,799
1941	715,990	1972	1,108,434	2003	1,575,496
1942	731,466	1973	1,126,025	2004	1,584,220
1943	745,519	1974	1,145,156	2005	1,593,669
1944	760,620	1975	1,164,476	2006	1,603,432
1945	773,894	1976	1,183,382	2007	1,613,635
1946	788,645	1977	1,204,403	2008	1,622,251
1947	802,574	1978	1,224,315	2009	1,630,538
1948	816,193	1979	1,245,966	2010	1,639,140
1949	828,924	1980	1,267,075	2011	1,648,877
1950	841,658	1981	1,287,339	2012	1,658,450
1951	852,629	1982	1,307,880	2013	1,666,729
1952	865,024	1983	1,328,015	2014	1,673,842
1953	874,894	1984	1,347,417	2015	1,681,571
1954	885,822	1985	1,366,484	2016	1,688,858
1955	896,022	1986	1,382,844	2017	1,696,287
1956	906,033	1987	1,399,964	2018	1,703,310
1957	915,681	1988	1,415,691	2019	1,708,276
1958	924,648	1989	1,431,576	2020	1,712,960
1959	933,652	1990	1,446,637		
1960	944,395	1991	1,459,477		
1961	954,227	1992	1,471,519		
1962	965,829	1993	1,482,796		
1963	976,944	1994	1,493,492		
1964	988,660	1995	1,503,859		
1965	1,001,255	1996	1,513,377		
1966	1,014,264	1997	1,522,723		
1967	1,028,125	1998	1,531,408		
1968	1,042,675	1999	1,540,102		
1969	1,058,925	2000	1,548,612		
1970	1,075,162	2001	1,557,809		

\*Expressed in millions of 1970 dollars. Construction described in text.

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