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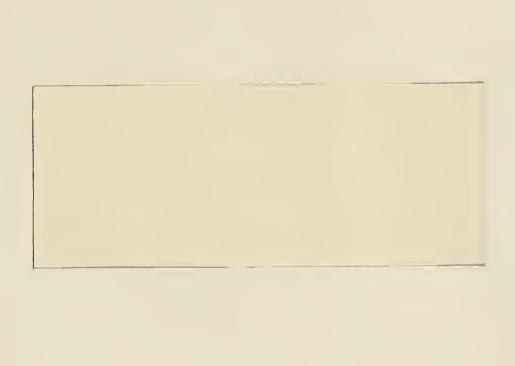
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M.I.T. Working Paper #358

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# OPTIMAL BAYESIAN MECHANISMS

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#### Abstract

Arrow [1079] and D'Aspremont and Gerard-Varet [1070] showed that when agents' preference parameters are independently distributed there exist public decision-making mechanisms that take optimal decisions in Bayesian equilibrium and balance the budget. We show that this result extends to the case where the parameters are correlated.



## Optimal Bayesian Mechanisms

#### Introduction

Suppose that a population of agents i = I,...,n have von Neumann-Morgenstern utility functions

$$v_i(d,\Theta_i) + t_i$$

where d is a public decision and  $t_i$  is a transfer of a private good to agent i. The parameter  $\theta_i$  is private information for agent i and the joint distribution of  $(\theta_1,\ldots,\theta_n)$  is given by the c. d. f.  $F(\theta_1,\ldots,\theta_n)$ , which is common knowledge. Suppose that  $d^*(\theta_1,\ldots,\theta_n)$  solves the problem

(0) 
$$\max_{\substack{\text{deD } i=1}}^{n} v_i(d, \theta_i),$$

where D is the space of possible public decisions (assume a solution to (0) exists). We shall refer to  $d^*(\theta_1,\ldots,\theta_n)$  as the optimal public decision given  $\theta_1,\ldots,\theta_n$ .

A <u>mechanism</u> is a public decision function  $d(\theta_1,\dots\theta_n)$  and a set of transfers  $\{t_i(\theta,\dots,\theta_n)\}$ . The interpretation of such a mechanism is that agents announce the values of their parameters (possibly untruthfully) and, on the basis of the announcements  $(\theta_1,\dots,\theta_n)$ , the public decision  $d(\theta_1,\dots,\theta_n)$  is taken and agent i receives transfer  $t_i(\theta_1,\dots,\theta_n)$ . In an <u>optimal Bavesian mechanism</u>, for all  $(\theta_1,\dots,\theta_n)$ ,

(1) 
$$d(\theta_1, \dots, \theta_n) = d^*(\theta_1, \dots, \theta_n)$$

(2) 
$$\theta_{i} \text{ maximizes } E_{\theta_{-i} \mid \theta_{i}} [v_{i}(d*(\theta_{i}, \theta_{-i}), \theta_{i}) + t_{i}(\theta_{i}, \theta_{-i})]^{1}$$

and

(3) 
$$\sum_{i} t_{i}(\theta_{1}, \dots, \theta_{n}) = 0,$$

where  $\mathbf{E}_{\theta_{-i} \mid \theta_{i}}$  denotes the expectation operator with respect to  $\theta_{-i}$  conditional on

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truth about his parameter, assuming that other agents are truthful, (b) chooses the optimal public decision and (c) balances the budget.

Arrow [1979] and D'Aspremont and Gerard-Varet [1979] (ADG) showed that optimal Bayesian mechanisms exist when the  $\theta_i$ 's are independently distributed. The transfers in these mechanisms take the form

(4) 
$$t_{i}(\theta_{1},...,\theta_{n}) = t_{ii}(\theta_{i}) - \frac{1}{n-1} \sum_{j\neq i} t_{jj}(\theta_{j}),$$

where

(5) 
$$t_{ii}(\theta_i) = E_{\theta_{-i}, j\neq i} v_j(d*(\theta_i, \theta_{-i}), \theta_j).^2$$

Laffont and Maskin [1979] and Riordan [1984] showed that the independence hypothesis in the ADG proposition can be weakened. Making stronger assumptions about the  $v_i$ 's and  $\theta_i$ 's, they demonstrated that when, roughly speaking, the  $\theta_i$ 's are nonpositively correlated, optimal Bayesian mechanisms can be found where transfers take the additively separable form (4), although  $t_{ii}$  is no longer given by (5).<sup>3</sup>

In this paper we present two results. We first provide a proposition in the spirit of the Laffont-Maskin-Riordan results on additively separable transfers (Theorem 1). We then argue (Theorem 2) that if one does not impose the additively separable form, one can find optimal Bayesian mechanisms regardless of the nature of the correlation across  $\theta_i$ 's.

### 1. Additively Separable Transfers

For the purposes of this section only, assume that (a) the space of public decisions D is (0,1); (b) agent i's parameter  $\theta_i$  is a number in (0,1) and that agent i has a differentiable conditional density function  $f_i(\theta_{-i}|\theta_i)$  corresponding to  $F(\theta_1,\ldots,\theta_n)$ ; (c) the function  $v_i(\cdot,\cdot)$  is strictly concave in

its first argument, twice continuously differentiable, and satisfies

(6) 
$$\frac{\partial^2 v_i}{\partial d\theta\theta_i} > 0.$$

Suppose first that we attempt to mimic the ADG solution (5). That is, suppose we take

(7) 
$$t_{ii}(\overset{\circ}{\theta}_{i}) = E_{\theta_{-i}}|\overset{\circ}{\theta_{i}}\overset{\sum}{j\neq i}v_{j}(\overset{\circ}{d*}(\overset{\circ}{\theta}_{i},\theta_{-i}),\theta_{j})$$
$$= \int_{\overset{\circ}{j\neq i}}v_{j}(\overset{\circ}{d*}(\overset{\circ}{\theta}_{i},\theta_{-i}),\theta_{j})f(\theta_{-i}|\overset{\circ}{\theta}_{i})d\theta_{-i}$$

and then define t by (4). Then, assuming that other agents are truthful, agent i maximizes

(8) 
$$\int v_{i}(d^{*}(\theta_{i},\theta_{i}),\theta_{i})f(\theta_{-i}|\theta_{i})d\theta_{-i} + t_{ii}(\theta_{i}),$$

since  $t_{jj}(\theta_j)$ ,  $j \neq i$ , does not depend on  $\theta_i$ . The first derivative of (7) is

$$\int \left[\frac{\partial v_{i}}{\partial d}(d^{*}(\Theta_{i},\Theta_{-i})\Theta_{i})\frac{\partial d^{*}}{\partial \Theta_{i}}(\Theta_{i},\Theta_{-i})f(\Theta_{-i}|\Theta_{i})\right]$$

$$(9) + \sum_{j \neq i} \frac{\partial v_{j}}{\partial a} (d^{*}(\Theta_{i}, \Theta_{-i}), \Theta_{i}) \frac{\partial d^{*}}{\partial \Theta_{i}} (\Theta_{i}, \Theta_{-i}) f(\Theta_{-i} | \Theta_{i})$$

$$+ \sum_{j \neq i} v_{j} (d^{*}(\Theta_{i}, \Theta_{-i}), \Theta_{i}) \frac{\partial f_{i}}{\partial \Theta_{i}} (\Theta_{-i} | \Theta_{i}) ] d\Theta_{-i}.$$

Now, because  $d^*(\theta_1, \dots, \theta_n)$  solves (1)

(10) 
$$\sum_{j=1}^{n} \frac{\partial v_{j}}{\partial a} (d^{*}(\Theta_{i}, \dots, \Theta_{n}), \Theta_{j}) = 0.$$

Therefore, when  $\theta_{i} = \theta_{i}$ , (9) becomes

But unless  $\frac{\partial f_i}{\partial \theta_i} = 0$ , there is no reason why (11) should vanish. Hence, although truthtelling satisfies the first-order conditions for a maximum when the  $\theta_i$ 's are independent, truthtelling is not optimal when the  $\theta_i$ 's are correlated. That is why the ADG procedure requires independence.

However, suppose instead that we define  $t_{ii} (\theta_i)$  so that it satisfies the first order conditions for a maximum at  $\theta_i = \theta_i$ . Then, for all  $\theta_i$ ,

$$(12) \qquad -\int_{\overline{\partial d}}^{\partial v_{\underline{i}}} (d^{*}(\Theta_{\underline{i}}, \Theta_{-\underline{i}}), \Theta_{\underline{i}}) \frac{\partial d^{*}(\Theta_{\underline{i}}, \Theta_{-\underline{i}}) f_{\underline{i}}(\Theta_{-\underline{i}} | \Theta_{\underline{i}}) d\Theta_{-\underline{i}} = \frac{\partial t_{\underline{i}\underline{i}}}{\partial \Theta_{\underline{i}}} (\Theta_{\underline{i}}).$$

With  $t_{ij}(\cdot)$  defined by (12), the first derivative of agent i's maximand becomes

$$\int \left[\frac{\partial v_{i}}{\partial d}(d^{*}(\theta_{i},\theta_{-i}),\theta_{i})f_{i}(\theta_{-i}|\theta_{i})\right] \\
-\frac{\partial v_{i}}{\partial d}(d^{*}(\theta_{i},\theta_{-i}),\theta_{i})f_{i}(\theta_{-i}|\theta_{i})\frac{\partial d}{\partial \theta_{i}}(\theta_{i},\theta_{-i})d\theta_{-i}.$$
(13)

Clearly, (13) vanishes at  $\theta_{\bf i} = \theta_{\bf i}$ . Also, in view of (6),  $\frac{\partial d^*}{\partial \theta_{\bf i}}(\theta_{\bf i}^0, \theta_{-\bf i}) > 0$ . Therefore, to establish that truth-telling is optimal, it suffices to show that (13) is nonpositive for  $\theta_{\bf i} > \theta_{\bf i}$  and nonnegative for  $\theta_{\bf i} < \theta_{\bf i}$ . Now these second-order conditions will not, in general, be satisfied. To ensure that they hold, we must impose stronger conditions on  $f_{\bf i}$ . For any vector  $\theta_{-\bf i}$ , let  $\theta_{\bf i}^*(\theta_{-\bf i})$  be a value of agent i's parameter such that

(14) 
$$\frac{\partial v_{i}}{\partial d}(d*(\Theta_{i}^{*}(\Theta_{-i}),\Theta_{-i}),\Theta_{i}^{*}(\Theta_{-i})) = 0.$$

If  $\theta_{i}^{*}(\theta_{-i})$  exists, it is unique. To see this, suppose that  $\theta_{i} = \theta_{i}^{*}$  satisfies

(15) 
$$\frac{\partial v_{i}}{\partial d}(d^{*}(\theta_{i}^{*}, \theta_{i}^{*}), \theta_{i}^{*}) = 0.$$

Differentiating  $\frac{\partial v_i}{\partial d}(d^*(\theta_i,\theta_{-i}),\theta_i)$  with respect to  $\theta_i$  we obtain

(16) 
$$\frac{\partial^{2} v_{i}}{\partial d^{2}} (d^{*}(\theta_{i}, \theta_{-i}), \theta_{i}) \frac{\partial d^{*}}{\partial \theta_{i}} (\theta_{i}, \theta_{-i}) + \frac{\partial^{2} v_{i}}{\partial d \partial \theta_{i}} (d^{*}(\theta_{i}, \theta_{-i}), \theta_{i}).$$

From (0),

(17) 
$$\frac{\partial d^*}{\partial \theta_{i}} = \frac{-\partial^2 v_{i}}{\partial d \partial \theta_{i}} / \sum_{j=1}^{n} \frac{\partial^2 v_{j}}{\partial d \partial \theta_{i}}$$

Substituting (17) into (16) we obtain

(18) 
$$\frac{\sum_{\substack{j \neq i \text{ } \partial d^2}}^{\delta^2 v_j}}{\sum_{\substack{j=1 \text{ } \partial d^2}}^{n} \frac{\delta^2 v_j}{\delta d \delta \Theta_i}},$$

which is positive. Hence,

(19) 
$$\frac{\partial v_{i}}{\partial d}(d^{*}(\theta_{i}, \theta_{-i}), \theta_{i}) \quad \begin{cases} > 0, \text{ if } \theta_{i} > \theta_{i}^{*} \\ < 0, \text{ of } \theta_{i} < \theta_{i}^{*} \end{cases}$$

establishing the uniqueness of  $\theta_{\hat{i}}^*(\theta_{-\hat{i}})$ . If  $\theta_{\hat{i}}^*(\theta_{-\hat{i}})$  fails to exist, set it equal to 1 if  $\frac{\partial v_{\hat{i}}}{\partial \hat{d}}(\hat{d}^*(\theta_{\hat{i}},\theta_{-\hat{i}}),\theta_{\hat{i}})$  is positive for all  $\theta_{\hat{i}}$ , and equal to zero, otherwise. From the definition of  $\hat{d}^*$ ,

$$\sum_{j=1}^{n} \frac{\partial v_{j}}{\partial \hat{d}} (d^{*}(\theta_{i}^{*}, \theta_{-i}), \theta_{-i}) = 0.$$

Thus,  $\theta_{\mathbf{i}}^*(\theta_{-\mathbf{i}})$  is that value of  $\theta_{\mathbf{i}}$  which makes the social and individual marginal products of d both zero. In that case,  $\theta_{\mathbf{i}}^*(\theta_{-\mathbf{i}})$  is the <u>representative</u> or <u>average</u> value of  $\theta_{-\mathbf{i}}$ .

One way of formalizing the idea that  $\theta_i$  and  $\theta_{-i}$  are nonpositively correlated is to suppose that as  $\theta_i$  moves away from the average value of  $\theta_{-i}$  the conditional density  $f_i(\theta_{-i}|\theta_i)$  does not fall. That is,

(20) 
$$\frac{\partial f_{i}}{\partial \theta_{i}}(\theta_{-i}|\theta_{i}) \quad \left\{ \begin{array}{l} \geq 0, \text{ if } \theta_{i} \geq \theta_{i}^{*}(\theta_{-i}) \\ \leq 0, \text{ if } \theta_{i} \leq \theta_{i}^{*}(\theta_{-i}) \end{array} \right.$$

This is the condition we need to establish that truthtelling is optimal.

Theorem 1: Under the assumptions we have made about  $v_i$  and  $\theta_i$  in this section, an optimal Bayesian mechanism exists if, for all i,  $f_i$  satisfies (20).

<u>Proof:</u> We need only show that the bracketed expression of (13) is nonpositive o connegative) for  $\Theta_{\bf i}$  greater (less) than  $\Theta_{\bf i}$ . Consider  $\Theta_{\bf i} > \Theta_{\bf i}$ .

Suppose first that

(21) 
$$\frac{\partial v_{i}}{\partial d}(d*(\theta_{i}, \theta_{-i}), \theta_{i}) > 0.$$

Then, because  $d^*$  is increasing in  $\theta_i$  and  $v_i$  is concave in d,

(22) 
$$\frac{\partial v_{i}}{\partial d}(d^{*}(\theta_{i}, \theta_{-i}), \theta_{i}) > 0.$$

Furthermore, from (6)

(23) 
$$\frac{\partial v_{\underline{i}}}{\partial d}(d^{*}(\theta_{\underline{i}}, \theta_{-\underline{i}}), \theta_{\underline{i}}) > \frac{\partial v_{\underline{i}}}{\partial d}(d^{*}(\theta_{\underline{i}}, \theta_{-\underline{i}}), \theta_{\underline{i}}) > 0$$

From (19), (22) and (23) we have

$$\Theta_{i}^{*}(\Theta_{-i}) < \Theta_{i} < \Theta_{i}.$$

Therefore, from (20),

(24) 
$$f(\theta_{-i} | \theta_{i}) \leq f(\theta_{-i} | \theta_{i}).$$

But (23) and (24) together imply that the bracketed expression in (13) is nonpositive.

Suppose next that

(25) 
$$\frac{\partial v_{i}}{\partial d}(d^{*}(\theta_{i}, \theta_{-i}), \theta_{i}) \leq 0.$$

Then, the bracketed expression is nonpositive unless

(26) 
$$\frac{\partial v_{\underline{i}}}{\partial d}(d^*(\theta_{\underline{i}}, \theta_{-\underline{i}}), \theta_{\underline{i}}) < 0.$$

But (26) implies that

$$\Theta_{i} < \Theta_{i} < \Theta_{i}^{*}(\Theta_{-i}),$$
 which in turn means that

(27) 
$$f(\theta_{-i} | \theta_{i}) \geq f(\theta_{-i} | \theta_{i}).$$

Combining (25) - (27), we conclude once again that the bracketed expression is one nonpositive. The argument for  $\theta_i$  <  $\theta_i$  is similar.

Q.E.D.

If the inequalities in (20) are reversed, then one cannot find a transfer of function  $t_{ii}(\theta_i)$  that induces agent i to tell the truth. Hence, with positive correlation, there does not exist an optimal Bayesian mechanism with transfer functions of the form (4).

The conclusion that negative rather than positive correlation makes incentive requirements easier to fulfill accords well with intuition, as Laffont and Maskin [1979] point out. Positive correlation aggravates the free rider problem. If an agent believes his tastes are similar to those of others, he can relatively safely leave provision of a public good in their hands.

## 2. Nonseparable Transfers

One limitation of using an additively separable transfer function is that it does not exploit the differences in beliefs corresponding to different values of  $\theta_i$ . If  $f_{-i}(\theta_{-i}|\theta_i)$  depends on the value of  $\theta_i$ , then an agent of type  $\theta_i$  will view a transfer depending on  $\theta_{-i}$  as a different gamble than will an agent of type  $\theta_i$ . But those different views are irrelevant to i's maximization problem if the terms in  $\theta_i$  are separable from the terms in  $\theta_{-i}$ . Thus additively separable transfer functions reduce our ability to discriminate among types. This assertion is confirmed by our next result, which illustrates the power of more general transfers. Henceforth, we drop the special assumptions of section 1 and revert to the less structured model of the introduction.

Theorem 2: In the model of section 1, there exists an optimal Bayesian mechanism.

<u>Proof</u>: Becasue the proof is virtually entirely algebraic manipulation, it will be helpful to consider the simplest possible case to illustrate the ideas as clearly as possible. Accordingly, suppose that there are just two agents and that  $\theta_i$  can assume just two values:  $\theta^1$  and  $\theta^2$ . After we go through the argument for this case, it should be apparent how the proof generalizes.

For any i, j, k,  $\epsilon$  {1,2} let

$$v_h^{ijk} = v_h(d*(\Theta^i,\Theta^j),\Theta^k)$$

and let  $p^{ij}$  be the joint probability that  $\theta_1 = \theta^i$  and  $\theta_2 = \theta^j$ . We shall take  $t^{ij}$  to be the transfer to agent 1 if  $\theta_1 = \theta^i$  and  $\theta_2 = \theta^j$ . Hence, for balance, the transfer to agent 2 in that event is  $-t^{ij}$ . Our problem is to find numbers  $t^{11}$ ,

t<sup>12</sup>, t<sup>21</sup>, and t<sup>22</sup> such that

(30) 
$$\sum_{i} p^{i1} (v_2^{i11} - t^{i1}) \ge \sum_{i} p^{i1} (v_2^{i21} - t^{i2})$$

(31) 
$$\sum_{p}^{i2}(v_2^{i22} - t^{i2}) \ge \sum_{i} p^{i2}(v_2^{i12} - t^{i1}).$$

It clearly suffices to find tij's such that

(32) 
$$v_1^{iji} + t^{ij} \ge v_1^{kji} + t^{kj}$$

and

(33) 
$$v_2^{ijj} - t^{ij} \ge v_2^{ikj} - t^{ik},$$

for all i, j, and k. To see that we can find such tij's, first set

$$(34)$$
  $t^{11} = 0.$ 

From (32), we must have

$$v_1^{111} \ge v_1^{211} + t^{21}$$
.

Hence,

(35) 
$$t^{21} = v_1^{111} - v_1^{211} - \alpha,$$

where

(36) 
$$\alpha \geq 0$$
.

From (32),

(37) 
$$v_1^{212} + t^{21} \ge v_1^{112}$$
.

From (35) and (37) we have

(38) 
$$v_1^{212} - v_1^{112} + v_1^{111} - v_1^{211} \ge \alpha$$

We must show that the left hand side of (38) is nonnegative. But for any i, j, and k

$$v_1^{kjk} + v_2^{kjj} \ge v_1^{ijk} + v_2^{ijj},$$

that is,

(39) 
$$v_2^{kjj} - v_2^{ijj} \ge v_1^{ijk} - v_1^{kjk}$$

Similarly, we have

(40) 
$$v_2^{ijj} - v_2^{kjj} \ge v_1^{kji} - v_1^{iji}$$

Adding (39) and (40), we obtain

(41) 
$$0 \ge v_1^{ijk} - v_1^{kjk} + v_1^{kji} - v_1^{iji}.$$

Therefore, the left hand side of (38 is nonnegative after all. Similarly, we have

(42) 
$$0 \ge v_2^{ikj} - v_2^{ijj} + v_2^{ijk} - v_2^{ikk}$$

for all j and k. Now from (33) and (35)

$$v_2^{211} - v_1^{111} + v_1^{211} + \alpha \ge v_2^{221} - t^{22}$$
.

Therefore,

(43) 
$$t^{22} = v_2^{221} - v_2^{211} + v_1^{111} - v_1^{211} - \alpha + \beta,$$

where

$$(44) \beta \geq 0.$$

From (33) we require

$$(45) v_2^{222} - t^{22} \ge v_2^{212} - t^{21}.$$

Substituting from (35) and (43), we can rewrite (45) as

(46) 
$$v_2^{222} - v_2^{212} + v_2^{211} - v_2^{221} \ge \beta$$

From (42) the left hand side of (46) is nonnegative.

Now, from (32),

(47) 
$$v_1^{222} + t^{22} \ge v_1^{122} + t^{12}$$
.

Using (43) we can rewrite (47) as

(48) 
$$t^{12} = v_1^{222} - v_1^{122} + v_2^{221} - v_2^{211} + v_1^{111} - v_1^{211} - \alpha + \beta - \gamma,$$

where

$$(49) \qquad \qquad \gamma \geq 0.$$

From

$$v_1^{121} + t^{12} \ge v_1^{221} + t^{22}$$

we deduce that

(50) 
$$v_1^{121} - v_1^{221} + v_1^{222} - v_1^{122} \ge \gamma$$
,

which is possible in view of (41).

We have shown that when we define the  $t^{ij}$ 's by (34), (35), (43), and (48), all but two of the eight incentive constraints given by (32) and (33) are satisfied if  $\alpha$ ,  $\beta$  and  $\gamma$  are nonnegative and satisfy (38), (46), and (50). It remains to show that we can choose  $\alpha$ ,  $\beta$ , and  $\gamma$  within this range so that the last two incentive constraints,

(51) 
$$v_2^{111} - t^{11} \ge v_2^{121} - t^{12}$$

and

(52) 
$$v_2^{122} - t^{12} \ge v_2^{112} - t^{11}$$

also hold.

From (48), (51) and (52) can be reexpressed as

$$(v_{2}^{122}-v_{2}^{112}) - (v_{1}^{222}-v_{1}^{122}) - (v_{2}^{221}-v_{2}^{211}) - (v_{1}^{111}-v_{1}^{211})$$

$$\geq -\alpha+\beta-\gamma \geq (v_{2}^{121}-v_{2}^{111}) - (v_{1}^{222}-v_{1}^{122}) - (v_{2}^{221}-v_{2}^{211}) - (v_{1}^{111}-v_{1}^{211}).$$

Now, it is possible to satisfy (53) as long as (i) the maximum value of  $-\alpha+\beta-\gamma$  is not less than the rightmost expression in (53), and (ii) the minimum value of  $-\alpha+\beta-\gamma$  is not greater than the leftmost expression in (53).

The maximum value of  $-\alpha+\beta-\gamma$  occurs when  $\alpha=\gamma=0$  and  $\beta=v_2^{222}-v_2^{212}+v_2^{211}-v_2^{221}$ . Therefore, to establish (i) amounts to showing that  $v_2^{222}-v_2^{212}+v_2^{211}-v_2^{221} \geq v_2^{121}-v_2^{111}-(v_1^{222}-v_1^{122})-(v_2^{221}-v_2^{211})-(v_1^{111}-v_1^{211}),$ 

which reduces to

$$(v_2^{222} - v_2^{212}) + (v_2^{111} - v_2^{121}) + (v_1^{222} - v_1^{122}) + (v_1^{111} - v_1^{211}) \ge 0.$$

Now,

$$\mathbf{v}_{1}^{121} + \mathbf{v}_{2}^{122} \geq \mathbf{v}_{1}^{211} + \mathbf{v}_{2}^{212} = \mathbf{v}_{1}^{111} + (\mathbf{v}_{1}^{211} - \mathbf{v}_{1}^{111}) + \mathbf{v}_{2}^{222} + (\mathbf{v}_{2}^{212} - \mathbf{v}_{2}^{222}).$$

That is,

(55) 
$$0 \ge (v_1^{111} - v_1^{121}) + (v_2^{222} - v_2^{122}) + (v_1^{211} - v_1^{111}) + (v_2^{212} - v_2^{222}).$$

Furthermore, from (40) we have

(56) 
$$v_1^{111} - v_1^{121} \ge v_2^{121} - v_2^{111},$$

and

(57) 
$$v_2^{222} - v_2^{122} \ge v_1^{122} - v_1^{222}$$

Substituting (56) and (57) in (55) we obtain

$$(58) 0 \ge (v_1^{122} - v_1^{222}) + (v_2^{121} - v_2^{111}) + (v_1^{211} - v_1^{111}) + (v_2^{212} - v_2^{222}).$$

Similarly, by permuting indices, we get

(59) 
$$0 \ge (v_1^{112} - v_1^{212}) + (v_2^{112} - v_2^{122}) + (v_1^{221} - v_1^{121}) + (v_2^{221} - v_2^{211}).$$

Now, (58) is just (54), establishing (i).

The minimum value of  $-\alpha+\beta-\gamma$  occurs when  $\beta=0$ ,  $\alpha=v_1^{212}-v_1^{112}+v_1^{111}-v_1^{211}$ , and  $\gamma=v_1^{121}-v_1^{221}+v_1^{222}-v_1^{122}$ . Therefore, establishing (ii) amounts to showing that

$$(v_{2}^{122} - v_{2}^{112}) - (v_{1}^{222} - v_{1}^{122}) - (v_{2}^{221} - v_{2}^{211}) - (v_{1}^{111} - v_{1}^{211})$$

$$\geq -[(v_{1}^{212} - v_{1}^{112}) + (v_{1}^{111} - v_{1}^{211}) + (v_{1}^{121} - v_{1}^{221}) + (v_{1}^{222} - v_{1}^{122})]$$

After cancellation and rearrangement, (55) becomes

$$(v_2^{122}-v_2^{112}) + (v_2^{211}-v_2^{221}) + (v_1^{212}-v_1^{112}) + (v_1^{121}-v_1^{221}) \ge 0,$$

which is just (59). Hence (ii) is established.

Q.E.D.

Notice that because the transfers  $t^{ij}$  do not depend on the probabilities, the theorem and method of proof continue to hold when agents have different beliefs about the joint distribution of the  $\theta_i$ 's (as long as those beliefs are common knowledge).

#### Footnotes

If f is a function of  $\theta_1, \dots, \theta_n$ , the notation  $f(\theta_i, \theta_{-i})$  is shorthand for  $f(\theta_1, \dots, \theta_{i-1}, \theta_i, \theta_{i+1}, \dots, \theta_n)$ .

<sup>2</sup>Because the  $\theta_i$ 's are presumed to be independent, the expectation in (5) is unconditional.

<sup>3</sup> See also D'Aspremont and Gerard-Varet [1982] for results on the correlated case.

"d\* is differentiable because it solves (1) and because v is strictly concave and twice continuously differentiable.

<sup>5</sup>Laffont and Maskin [1979] show that, under the hypotheses of Theorem 1,  $\theta_i = \theta_i$  is a <u>local</u> maximum when  $t_{ii}$  is defined by (12). Riordan [1984] establishes a result related to the Theorem when n=2. Although his result is couched in terms of <u>positive</u> correlation, he correlates costs and benefits. Therefore, his positive correlation of costs and benefits amounts to negative correlation of benefits.

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